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The structure of psychopathology in adolescence and its common personality and cognitive correlates.

Natalie Castellanos-Ryan¹,* Frederic N. Brière¹, Maeve O'Leary-Barrett², Tobias

Banaschewski³, Arun Bokde⁴, Uli Bromberg⁵, Christian Büchel⁵, Herta Flor⁶, Vincent

Frouin⁷, Juergen Gallinat⁵, Hugh Garavan⁸, Jean-Luc Martinot⁹, Frauke Nees⁶, Tomas

Paus¹⁰, Zdenka Pausova¹¹, Marcella Rietschel⁶, Trevor W. Robbins¹², Michael N.

Smolka¹³, Robert Whelan¹⁴, Gunter Schumann¹⁵, Patricia Conrod^{16,17}* and The IMAGEN

Consortium

1) Department of Psychoeducation, University of Montreal, Montreal, Canada; 2) the Department of Psychology, McGill University, Montreal, Canada; 3) Department of Child and Adolescent Psychiatry and Psychotherapy, Central Institute of Mental Health, Medical Faculty Mannheim, Heidelberg University, Germany; 4) the Institute of Neuroscience and Discipline of Psychiatry, School of Medicine, Trinity College Dublin, Ireland; 5) University Medical Center Hamburg-Eppendorf, Hamburg, Germany; 6) Department of Cognitive and Clinical Neuroscience, Central Institute of Mental Health, Medical Faculty Mannheim, Heidelberg University, Germany; 7) Neurospin, Atomic Energy and Alternative Energies Commission, Paris, France; 8) the Departments of

Psychiatry and Psychology, University of Vermont, Burlington, USA; 9) Institut National de la Santé et de la Recherche Médicale, INSERM Unit 1000 "Neuroimaging & Psychiatry", University Paris Sud, University Paris Descartes - Sorbonne Paris Cité, and Maison de Solenn, Paris, France; 10) Rotman Research Institute, Departments of Psychology and Psychiatry, University of Toronto, Toronto, Canada; 11) Hospital for Sick Children, University of Toronto, Toronto, Canada; 12) the Department of Experimental Psychology, University of Cambridge, Cambridge, United Kingdom; 13) the Department of Psychiatry and the Neuroimaging Center, Technische Universität Dresden, Germany; 14) the Department of Psychology, University College Dublin, Ireland; 15) and MRC Social, Genetic, and Developmental Psychiatry (SGDP) Centre, London and Psychology & Neuroscience, King's College London, United Kingdom; 16) Department of Psychiatry, University of Montreal and CHU Sainte-Justine Hospital, Montreal, Canada; 17) Department of Psychological Medicine and Psychiatry, Institute of Psychiatry, Psychology & Neuroscience, King's College London, United Kingdom.

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Corresponding Authors: Natalie Castellanos-Ryan, <u>Natalie.castellanos-ryan@umontreal.ca</u>, Department of Psychoeducation, University of Montreal; Patricia Conrod, <u>patricia.conrod@umontreal.ca</u>, Department of Psychiatry, University of Montreal, CHU Ste-Justine.

Abstract

The traditional view that mental disorders are distinct, categorical disorders has been challenged by evidence that disorders are highly comorbid and exist on a continuum (e.g., Caspi et al, 2014; Tackett et al, 2013). The first objective of this study was to use structural equation modelling to model the structure of psychopathology in an adolescent community-based sample (N=2144) including conduct disorder, attention-deficit/ hyperactivity disorder (ADHD), oppositional-defiant disorder (ODD), obsessivecompulsive disorder, eating disorders, substance use, anxiety, depression, phobias and other emotional symptoms, assessed at 16 years. The second objective was to identify common personality and cognitive correlates of psychopathology, assessed at 14 years. Results showed that psychopathology at 16 years fit two bifactor models equally well: a) a bi-factor model, reflecting a general psychopathology factor, as well as specific externalizing (representing mainly substance misuse and low ADHD) and internalizing factors; and b) a bi-factor model with a general psychopathology factor and three specific externalizing (representing mainly ADHD and ODD), substance use and internalizing factors. The general psychopathology factor was related to high disinhibition/impulsivity, low agreeableness, high neuroticism and hopelessness, high delay-discounting, poor response inhibition and low performance IQ. Substance use was specifically related to high novelty-seeking, sensation-seeking, extraversion, high verbal IQ and risk-taking. Internalizing psychopathology was specifically related to high neuroticism, hopelessness and anxiety-sensitivity, low novelty-seeking and extraversion and an attentional bias toward negatively-valenced verbal stimuli. Findings reveal several non-specific or trans-

diagnostic personality and cognitive factors that may be targeted in new interventions to potentially prevent the development of multiple psychopathologies.

Keywords: psychopathology, comorbidity, personality, cognition, adolescence

General Scientific Summary: This study provides support for a spectrum and common liability model of psychopathology, with findings showing that in a community sample of European adolescents: a) internalizing and externalizing psychopathology share substantial variance that can be modeled as a general psychopathology factor; and b) that the common variance across psychopathologies is associated with personality traits related to high disinhibition/impulsivity, low agreeableness, high neuroticism and hopelessness, as well as different aspect of cognitive function, namely high delay-discounting, poor response inhibition and low performance IQ.

Introduction

The traditional view that mental disorders are distinct, categorical syndromes has been challenged by evidence that many disorders are highly comorbid and exist on a continuum. Indeed, comorbidity rates are high among psychiatric disorders (Kessler, Chiu, Demler, Merikangas, & Walters, 2005) and comorbid cases represent the rule rather than the exception. This suggests that mental disorders have more in common than implied by current nosology, or at least that there may be a more parsimonious structure to psychopathology. Since the mid-90s, there has been substantial research suggesting that certain psychiatric disorders in adulthood may share substantial common variance and exist on at least two liability continuums or spectra (e.g., Kendler, Prescott, Myers, & Neale, 2003; Krueger, 1999): an externalizing dimension conveying risk for disinhibited, antisocial and/or substance use disorders, and an internalizing dimension conveying liability to experience mood and anxiety disorders. Importantly, these findings in adulthood confirmed dimensional models of psychopathology frequently used and replicated in child samples since the 80's (Achenbach & Edelbrock, 1981), suggesting stability of this structure across developmental periods. However, although these internalizing and externalizing dimensions have been replicated across populations and developmental periods, studies have also shown that there are significant positive correlations among these latent dimensions/factors. This has motivated a new wave of studies investigating the structure and implications of these inter-factor correlations. In line with this, a number of recent studies (Carragher et al., 2016; Caspi et al., 2014; Lahey et al., 2012; Laceulle et al., 2015; Lahey et al., 2015; Murray, Eisner & Ribeaud, 2016; Noordhof et al., 2015; Patalay et al., 2015; Petterson et al., 2015; Tackett et al.,

2013; Wright & Simms, 2015), have shown that psychopathology diagnoses and symptoms may be best modelled with a bi-factor (or general-specific) model, which includes 1) a general factor capturing the common variance shared across all symptoms and 2) specific factors reflecting the additional shared variance among symptoms, be they internalizing and externalizing factors (Lahey et al., 2015; Patalay et al., 2015; Tackett et al., 2013), externalizing, internalizing and thought disorder factors (Carragher et al., 2016; Caspi et al., 2014; Laceulle et al., 2015), externalizing, distress and fear factors (Lahey et al., 2012) or other factors (Murray et al., 2016; Noordhof et al., 2015; Wright & Simms, 2015). The replication of findings across studies is noteworthy given the range of samples and psychopathology disorders/symptoms included. Remarkably, these studies demonstrated that a general psychopathology factor (or P factor, as coined by Caspi et al., 2014) can be reliably replicated across samples and psychopathology disorders/symptoms, as well as developmental periods. Indeed, Murray et al., (2016) recently demonstrated that the P factor remained stable from age 7 to 16 years. Finally, these recent studies have also provided initial insights into the nature of the shared variance across psychopathology. For example, higher scores on the P factor were associated with greater life impairment (Caspi et al., 2014), lower school attainment, and higher psychopathology later in life (Patalay et al., 2015; Petterson et al., 2015), greater economic deprivation in childhood (Caspi et al., 2014; Patalay et al., 2015) and lower IQ and school functioning (Caspi et al., 2014; Lahey et al., 2015), as well as higher neuroticism and lower agreeableness and conscientiousness (Caspi et al., 2014). Using a behavioural genetics design, Tackett et al. (2013) found that high scores on the P factor were associated with higher negative emotionality, and that there was substantial variance

shared between the P factor and negative emotionality at both phenotypic and genetic levels.

Taken together, these findings suggest that a general psychopathology factor can be modelled across different measures of psychopathology, is robust across developmental periods, reflects shared personality and IQ correlates, as well as environmental and genetic influences, and may improve prediction of future individual functioning. However, the nature of a general psychopathology factor in adolescence remains understudied. Importantly, although there are a number of studies that include adolescent samples (Carragher et al., 2016; Laceulle et al., 2015; Murray et al., 2016; Noordhof et al., 2015; Patalay et al., 2015; Tackett et al., 2013), only one study so far has modelled substance use (i.e., alcohol use problems, Carragher et al., 2016), and none has included eating disorders. This is an important limitation since the peak of onset for many psychopathology disorders, particularly the onset of SU problems (Kessler et al., 2005) and eating disorders (Stice et al, 2009, Nagl et al, 2016) occurs during adolescence. Second, only two studies have examined the personality correlates of a general psychopathology factor in adolescence. As previously mentioned, Tackett et al. (2013) found that negative emotionality may be associated with a liability to developing a range of psychopathologies, which is in line with the wider personality-psychopathology literature showing that neuroticism is the Big Five trait most strongly associated with many forms of psychopathology, particularly mood and anxiety disorders (Widiger & Costa, 1994; Kotov, Gamez, Schmidt, & Watson, 2010), as well as their comorbidity (Khan, Jaconson, Gradner, Prescott & Kendler, 2005). Low conscientiousness and low extraversion have also consistently been associated with several mental health disorders

(Malouff, Thorsteinsson & Schutte, 2005; Kotov et al, 2010). Carragher and colleagues (2016) also modelled the common variance across several symptoms of emotional, hyperactive and conduct problems in Australian youth and found that several lower-order personality traits were associated with the common variance across these symptoms, with hopelessness and impulsivity being the strongest predictors. This finding is consistent with other studies showing associations between these lower-order factors and a range of psychopathologies (Castellanos-Ryan, O'Leary-Barrett, Sully, & Conrod, 2013; Kuo, Gallo, & Eaton, 2004) and their comorbidity (Castellanos-Ryan & Conrod, 2012; Castellanos-Ryan et al., 2014). A limitation of these two studies, however, is that assessments have been concurrent, which precludes drawing conclusions regarding the temporality of the associations between personality traits and the P factor. Other limitations are that these studies have not included both higher- and lower-order personality dimensions in relation to general-specific psychopathology factors and have not included a wide range of potential substance use problems, although these problems appear to be related to a distinct personality profile compared to many other forms of psychopathology – in that they are less related to neuroticism but more elevated on disinhibition, sensation-seeking and disagreeableness (Kotov et al, 2010; Castellanos-Ryan & Conrod, 2012). Thus, a longitudinal examination of associations between general and specific dimensions of psychopathology, which include a wider range of disorders and personality measures, would be useful.

Third, very few studies have been designed and sufficiently powered to examine the cognitive correlates of psychopathology when modelled hierarchically. Due to its large scale and comprehensive cognitive assessment battery, the IMAGEN study provides a

unique opportunity to examine the hierarchical structure of psychopathology and the cognitive correlates of such latent dimensions. We previously reported on the cognitive and functional imaging correlates of the latent structure of externalizing problems using this large sample of European adolescents (Castellanos-Ryan et al., 2014). A generalspecific factor was supported by structural equation modelling, and neuropsychological and functional imaging tasks were shown to dissociate the three latent factors concurrently and longitudinally. A general externalizing factor was specifically associated with delay discounting and hyperactivation of the pre-Supplementary Motor Cortex (coupled with hypoactivity of the Substantial Nigra and Subthalamic Nucleus) during successful behavioural inhibition. However, because this study only focused on externalizing symptoms, it is unclear to what extent this externalizing factor captured shared variance between externalizing symptoms that is in fact shared by a wider range of psychopathology symptoms and could be captured by a more general psychopathology factor. It also did not allow the identification of the cognitive correlates that are associated to the shared variance across externalizing and internalizing symptoms. Caspi et al. (2014) examined the associations between the P factor and cognitive functions and other brain measures (i.e., retinal imaging), and showed that a derived P factor correlated negatively with adult and child measures of cognitive function and brain integrity generally. Similarly, Lahey et al. (2015) found that the P factor modelled in childhood was also associated with poor verbal and spatial IQ. These results are consistent with studies on individual disorders showing that lower intelligence and cognitive impairments, such as deficits in response inhibition and working memory, are shared by a wide range of externalizing problems, such as aggression and ADHD/CD (Lijffijt et al.,

2005; Oosterlaan & Sergeant, 1998; Peeters, Monshouwer, Janssen, Wiers, & Vollebergh, 2014; Young et al., 2009) and internalizing disorders, such as depression (Wagner et al., 2015; Snyder, 2013) and various anxiety disorders including social anxiety disorder, generalised anxiety disorder, obsessive-compulsive disorder and posttraumatic stress disorder (Ferreri, Lapp, & Peretti, 2011; Polak, Witteveen, Reitsma, & Olff, 2012; Olley, Malhi, & Sachdev, 2007). However, other cognitive risk factors which have been investigated in relation to individual disorders may be more specific to internalizing or externalizing dimensional factors. For instance, attentional biases towards stimuli with a negative emotional valence have been associated with several internalizing problems, but may be less characteristic of externalizing problems (Peckham, McHugh, & Otto, 2010). On the other hand, deficits related to delay discounting or risky decision making may be more specific to externalizing problems (Bickel, Koffarnus, Moody & Wilson, 2014; Castellanos-Ryan et al., 2014). However, these cognitive variables have never been examined in previous bifactor studies. Interestingly, Caspi et al. (2014) found no correlation between a specific externalizing factor and any brain measures, which is unexpected considering the large literature on cognitive and brain abnormalities in ADHD and CD (e.g., Castellanos-Ryan, Rubia, & Conrod, 2011) and the findings reported by Castellanos-Ryan et al., (2014) suggesting that significant cognitive impairment, particularly poor response inhibition, is mediated by hypofunction of prefrontal cortical circuits. However, this could be due to the fact that specific externalizing behavior in the Caspi et al. (2014) study was assessed mostly with substance use behaviours and included only conduct problems, but not ADHD, as a non-

substance use-related externalizing problem. Additional structural studies which include a wider range of potential externalizing symptoms are required to clarify this question. Within this context, our first aim was to model the structure of dimensional psychopathology in adolescence and determine whether it is best described by a generalspecific bifactor model, as found in previous studies. As adolescence is a developmental period characterised by important biological, cognitive, emotional and social changes, and the peak of onset for many psychopathology disorders, particularly the onset of SU problems (Kessler et al., 2005) and eating disorders (Stice et al, 2009, Nagl et al, 2016), we also aimed to examine the stability of the model across the transitional period of early- to mid-adolescence. To do this, we model the structure of psychopathology at two different ages in adolescence (14 and 16 years) in a community-based European sample including a diverse range of mental disorder symptoms. Importantly, we examine how three different substance misuse indicators (alcohol problems, drug use and tobacco use) and eating disorders integrate into a bifactor psychopathology model in adolescence, which has not been done before. Our third aim is to examine the prospective association between cognitive and personality correlates at age 14 and general-specific factors at age 16. Due to the extensive neuropsychological battery used in the IMAGEN study, the current analysis is unique in its ability to assess the extent to which delay discounting, risky decision making and attentional biases to positive and negative emotional stimuli (which have previously been associated with neuroticism) are associated with the general P factor or more specific factors. In this way, we will investigate the relationships between specific adolescent liability factors and cognition, beyond IQ and executive function, which has never been done before. We will also examine the prospective link

between higher- and lower-order personality measures and general-specific psychopathology factors, which has not been done in previous bi-factor studies. A final contribution of this study is to put our previous findings (Castellanos-Ryan et al., 2014) into context with respect to the personality and cognitive correlates of externalizing psychopathology, when models of psychopathology consider the internalizing spectrum in addition to the externalizing spectrum and their common variance.

These objectives will hopefully lead to a better understanding of common and specific correlates of psychopathology that can inform the development of new interventions (as we have done successfully for externalizing neurophenotypes; *e.g.*, Conrod et al., 2013; O'Leary-Barrett et al., 2013) that could potentially impact a multitude of outcomes by targeting personality and cognitive risk dimensions.

Method

Participants

A total of 2232 participants aged 14 years across 8 European sites were recruited via high-schools. Parents gave informed written consent and adolescents gave written assent to the study procedure prior to inclusion. All procedures were approved by each local institutional ethics committee. Further descriptions of the study design, sample and recruitment procedure, including data storage and safety can be found in the supplementary material and elsewhere (Schumann et al., 2010).

After data quality control, complete and reliable data sets for 2144 volunteers with an average age of 14.39 years (SD = 0.77) and an even sex ratio (n = 1093 girls, *i.e.*, 51 %) were included in analyses. Reliable follow-up data was gathered for 1603 (75%) participants at 16 years. Participants who were followed up did not differ significantly

from those not followed on demographic, psychopathology symptoms or cognitive variables, except for language (Odds Ratio=2.79, 95%CI=2.20-3.55, with English speakers being more likely to be followed) and parent-reported ADHD symptoms (Odds Ratio=0.93, 95%CI=0.87-0.99, with those scoring higher being less likely to be followed). All participants for whom we had reliable data at 14 years (N=2144) were included in analyses.

Measures

All measures were selected on the basis of brevity, age-appropriateness, and validity in their variant forms (English, German and French). Psychopathology and substance misuse symptoms over the last 12 months were assessed at 14 and 16 years of age.

Personality and cognitive function were only assessed at 14 years.

Psychopathology symptoms. Self- and parent-report behavioral and clinical measures were assessed via online computer platforms provided by Psytools ® (Delosis Ltd, London, UK), administered at participants' homes, and The Development and Well-Being Assessment interview (DAWBA, Goodman, A., Heiervang, E., Collishaw, S., & Goodman, R., 2011; see also http://www.dawba.info), was administered at the research site. The DAWBA interview was administered to adolescents and parents and assessed psychiatric symptoms of CD, ODD, ADHD, generalized anxiety, depression, specific phobia, social phobia, agoraphobia, panic disorder, OCD and eating disorders. A prognosis for the likelihood of having a disorder was calculated by computer algorithms that use the symptoms and impact recorded in the structured sections of both the parent-rated and self-rated DAWBA. These computer-generated band scores range from level 0 up to level 5, corresponding to the approximate prevalence rates in an epidemiological

sample for the disorder in question (ranging from less than 0.1% up to 70%). Diagnostic criteria were based on the Diagnostic Statistical Manual, Version 4. Because of low prevalence, likelihood of specific phobia, agoraphobia and panic disorder were averaged to create a combined "panic and other phobia" score. This was deemed justified as these three band scores were associated similarly to all other psychopathology indicators and correlates in this sample.

Substance misuse was assessed using the Alcohol Use Disorders Identification Test (AUDIT; Saunders, Aasland, Babor, de la Fuente, & Grant, 1993) and the European School Survey Project on Alcohol and Drugs (ESPAD; Hibell et al, 1997). The AUDIT was developed and validated by the World Health Organization to provide a brief assessment of alcohol use disorders and was specifically designed for international use. It exists in all three languages, and has been validated on primary health care patients and community samples. For this study, the scale total for problematic or harmful alcohol use in the last year included feelings of guilt or remorse after drinking, being unable to remember what happened the night before because of drinking, being injured or having injured someone as a result of drinking and relevant others being concerned about their drinking and suggestions to cut down. The ESPAD items used in this study comprised tobacco use frequency and the number of drugs used over the last 12 months. See Table 1 for prevalence and correlations between all psychopathology measures at 16 years. **Personality** was assessed with the self-reported Substance Use Risk Profile Scale (SURPS; Woicik, Stewart, Pihl, & Conrod, 2009), the NEO Five Factor Inventory (NEO-FFI; Costa Jr & McCrae, 1992) and the Temperament and Character Inventory (TCI; Cloninger, Przybeck, Syrakic, & Wetzel, 1994). The SURPS assessed the personality

traits of hopelessness, anxiety sensitivity, impulsivity and sensation-seeking. The reliability and concurrent and predictive validity of this measure has been well established in several adolescent and adult samples in different countries (Castellanos-Ryan et al., 2013; Krank et al., 2011; Woicik et al., 2009). The NEO-FFI assessed five higher order personality characteristics: neuroticism, conscientiousness, extraversion, agreeableness and openness to experience. The TCI was used to assess novelty-seeking, which is considered a good general measure of impulsive tendencies that also includes sensation-seeking (see Table S2 in supplementary material for correlations between personality and cognitive measures).

IQ and cognitive measures. Estimates of *intelligence* were derived from the vocabulary and similarities subtests (verbal IQ) and block design and matrix reasoning subtests (performance IQ) of the Wechsler Intelligence Scale for Children – Fourth Edition (WISC-IV; Wechsler, 2003). Digit Span forward and backward subtests from the WISC-IV were also administered and used to assess short-term auditory memory and auditory working memory (Groth-Marnat & Baker, 2003; Reynolds, 1997). *Poor Response Inhibition* was measured using the number of commission errors in a go-nogo Passive Avoidance Learning Paradigm (Newman & Wallace, 1993). *Delay discounting* was assessed with the Kirby Delay Discounting Questionnaire (Kirby, Petry, & Bickel, 1999). This measure was scored as described previously by Kirby et al. (1999), with k values (an index of delay discounting) assigned according to choice patterns across the 27 items. *Spatial working memory, risky decision-making and information processing biases for positive and negative stimuli* were assessed with three tasks from the Cambridge Cognition Neuropsychological Test Automated Battery (CANTAB; Cambridge

Cognition), the Spatial Working Memory (SWM), Cambridge Gambling task and affective go/ no-go task, respectively. SWM is a self-ordered test that requires retention and manipulation of visuospatial information. A modified version of the Cambridge Gambling Task (which reduced the time between stakes from 5 to 2 seconds to make the task shorter to avoid boredom effects in adolescents) was used to assess risky decision-making. Finally, the affective go/no-go is a task of emotional processing, in which affectively valenced words (happy and sad) are presented one at a time on screen.

Performance variables-of-interest are the target (omission) errors to positive and negative words, with an attentional bias towards negative versus positive words being assessed with a difference score between omission errors to each set of stimuli. For further details on the cognitive tasks see supplementary material and http://www.cambridgecognition.com/academic/cantabsuite.

Data Analysis

A series of structural equation models on computer-generated scaled likelihood of diagnosis scores and self-reported substance use were analysed using MPlus version 7.11 (Muthén & Muthén, 2010). Based on previously reported theoretical models and analyses (*e.g.*, Carragher et al., 2015; Caspi et al., 2014; Castellanos-Ryan et al., 2014; Lahey et al., 2012; Lahey et al., 2015; Tackett et al., 2013), several models were assessed for goodness of fit: 1) a single "Psychopathology" factor loading on all indicators; 2) two correlated factor models, where variables assessing internalizing (INT) psychopathology (generalised anxiety, depression, social phobia, panic and other phobias, OCD and eating disorders)¹ and CD, ODD, ADHD and substance use (SU) loaded on two specific INT

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¹ Alternative models were also tested in which OCD and eating disorder were entered as independent variables that did not load onto the internalizing factor, but rather correlated with each other and the internalizing, externalizing and

and EXT factors (with SU variables loading on the EXT factor; Model 2a) or loaded on three specific INT, externalizing (EXT) and SU factors (Model 2b). Model 2a and 2b allowed subfactors to covary. Lastly (3), two bi-factor (or general-specific) models were assessed, in which a general psychopathology factor (P) was added at the same level as the specific factors from the previous (2a and 2b) models (model 3a and 3b). In these last models, factors were not allowed to covary (i.e., they were constrained to zero), consistent with a classic bi-factor model, but because many recent studies present modified versions of bi-factor models (e.g., Carragher et al., 2015; Caspi et al., 2014), in which the specific factors are allowed to covary, two final revised bi-factors models (3a' and 3b') that allowed the specific factors to correlate were also examined (See Figures S1-S7 in supplementary material for a graphic representation of all models tested). In all models, gender and language (English versus other) were entered as covariates (at the level of observed variables). In addition, all models were fit using a complex random effects design to control for testing site as a cluster variable, and used robust maximum likelihood estimation (MLR). MLR has been shown to perform well when modeling low prevalent behaviors or nonnormal data (Asparouhov & Muthén, 2005). Full Information Maximum Likelihood was used to handle missing data.

Once the best fitting model was established, two sets of correlates (personality and cognitive indices) were each entered into the model separately. That is, unadjusted associations were examined by entering the personality and cognitive variables and the psychopathology factors into the same model and allowing them to correlate. Adjusted associations were examined in four separate models in which the psychopathology

substance misuse factors (as a form of correlated model), or loading directly on the P factor and no other factors (as an alternative form of bi-factor model). However, in both cases these models fit the data significantly worse than the models in which OCD and eating disorder loaded on to the internalizing factor.

factors were regressed onto 1) all SURPS subscales, 2) all NEO subscales, 3) novelty seeking (on its own), and 4) all cognitive variables entered together in the same model. The Benjamini-Hochberg procedure (Benjamini & Hochberg, 1995) was used to correct for multiple testing. Once p values are sorted in ascending values, the Benjamini-Hochberg procedure allows one to calculate the False Discovery Rate (FDR) for each of the p-values (*i.e.*, at each "position" in the sorted list of p-values, it will indicate what proportion of those are likely to be false rejections of the null hypothesis). This procedure to control for multiple testing has been shown to be less stringent and have more power than Bonferroni correction or other types of familywise error rate corrections. See online supplementary material for further description of the sample, measures and analytic approach.

Results

The structure of psychopathology in adolescence

Of the models tested, the three correlated factor model (model 2b) and all bi-factor models (both classic or modified) fit the data well, with the bi-factor model with the EXT and INT specific factors (Models 3a and 3a') or the modified bi-factor model with the three specific ADHD/CD/ODD, SU and INT factors fitting the data best (see Table 2). The fit was equally good for these three models, but model 3a (see figure 1) was chosen over model 3a' because of parsimony and the fact that the additional correlation did not improve model fit and was non-significant). Factor loadings for model 3a appear in Table 3, and factor loadings for model 3b' appear in Table S3 in supplementary material. In both models, all psychopathology indicators loaded significantly on the P factor, with CD, ODD, ADHD, smoking frequency, numbers of drugs used and depression loading

the strongest on this factor $(\geq .42)$, and eating disorders and social phobia loading the weakest on this factor (\leq .22). Indeed, at 16 years, the P factor represented mainly the substantial common variance shared among externalizing disorders (including smoking and drug use), depression and, to a slightly lesser extent, anxiety, OCD, phobias and drinking problems. The specific INT factors in both models 3a and 3b' reflected variance unique to all internalizing problem (albeit a low loading for eating disorders) measures. In model 3a the specific EXT factor reflected variance unique to all substance use measures and a low likelihood of ADHD diagnosis, whereas in model 3b', the specific SU and EXT factors respectively reflected the variance unique to all substance use measures and or variance unique to ADHD, ODD and CD. In model 3b', which allowed specific factors to covary, the specific SU factor was negatively associated with the specific INT factor (r=-.27, p=.014) and, the specific EXT factor at a trend level (r=-.15, p=.085); and the specific EXT and INT factors were not significantly associated (r=.26, p=.554). The negative loading of ADHD and ODD on the specific EXT factor in model 3a and the negative correlation between the specific EXT and SU factors in model 3b' suggest that, once the common variance shared among disorders is removed by the P factor, SU is associated with a lower likelihood of having ADHD and ODD. This is not the case for CD though, with modification indices for both models suggesting that there is some small residual variance shared (positive correlations) between CD and all SU items. Both models fit the data well and have advantages and disadvantages: model 3a is more parsimonious but includes positive and negative loadings on the EXT factor, while model 3b' includes factors with only positive loadings and is easier to interpret, but very little variance is captured by the specific EXT factor, which does not differ significantly

from zero (p=.694). Because of parsimony, and the fact that all of its factors had good variability, results for model 3a are provided in main text, while results for Model 3b' are presented as supplementary material.

Taken together, results for both models suggest that most of the variance common to all externalizing symptoms is accounted by a general psychopathology (P) factor rather than an externalizing factor in this sample of adolescents. Once the P factor was modelled, the externalizing factor captured variance specific to substance misuse symptoms (in model 3a), rather than the variance shared across all externalizing disorders. Thus, although referred to as a specific externalizing (EXT) factor, this factor in model 3a really represents a substance use and low ADHD factor.

How stable is this structure from early (14 years) to middle (16 years) adolescence? At 14 years the two specific factor bi-factor model of psychopathology (model 3a) also fit the data well (X² (42)=53.30, CFI=1.00, RMSEA=.011; SRMR=.016; BIC= 61037.87; Adj BIC= 60793.22), resulting in very similar loadings to those found at 16 years, with just slightly lower loadings for most internalizing symptoms on the P factor (see table S4 in supplementary material). Correlations between factors at 14 years and 16 years showed that factors were largely stable over two years, with non-significant or only small longitudinal correlations across factors (see bottom of Table 3). However, although largely stable across time, these factors were not found to be metrically invariant over time. That is, when factor loadings were constrained to be equal across time (*i.e.*, weak factorial invariance) the model fit was significantly worsened relative to when they were freely estimated (X²diff= 146.68, DFdiff=24, p<.001). Thus, after an inspection of the factor loading at 14 and 16 years, the loadings that did not demonstrate configural

invariance (*i.e.*, that differed across time) were freed, to test whether partial factorial invariance could be met. The model in which loadings for all internalizing indicators, except for eating disorders, number of drugs used and tobacco use were allowed to be freely estimated over time did not differ significantly from the base, freely estimated model (X²diff= 17.13, DFdiff=10, p=.072), indicating that the loadings for ADHD, CD, ODD, drinking problems and eating disorder did not differ across time. Taken together, these results suggest that while scores within factors were stable across time and very little longitudinal association existed across factors, and the P factor bi-factor structure fits well at both 14 and 16 years, the size of the contribution of psychopathology symptoms or indicators to the P factor may vary across development, with internalizing symptoms and drug and tobacco use becoming stronger with increasing age.

Unadjusted personality and cognitive correlates of psychopathology

All predictor models with covariates showed good model fit (see note under Table 4). Table 4 presents correlations between covariates and the bi-factor model of psychopathology from model 3a (associations with factors from model 3B' can be found in Table S5 in supplementary material). Results showed that after controlling for multiple testing, common variance across psychopathology (P factor) was significantly associated with high levels of impulsivity, novelty-seeking, neuroticism, hopelessness, sensation-seeking and extraversion, and low levels of agreeableness and conscientiousness. The P factor was also associated with high delay-discounting, low verbal and performance IQ, low working memory (spatial and verbal), poor response inhibition and risk-taking. Unique variance for EXT (SU and low ADHD) symptoms was significantly associated with high novelty-seeking, sensation-seeking and extraversion, high verbal IQ and high

risk-taking. In contrast, unique variance for INT symptoms was associated with high neuroticism, hopelessness and anxiety sensitivity, low novelty-seeking and extraversion, high conscientiousness, high attention (as measured by digit-span forward) and an attentional bias toward negatively-valenced verbal stimuli.

Adjusted associations between personality, cognitive correlates and psychopathology

In order to examine whether associations between personality, cognitive correlates and psychopathology factors remained once the effects of other personality and cognitive characteristics were adjusted for, different path analyses were conducted where correlations between correlates and factors were substituted by regression paths in the models. That is, for example, one model was conducted where all cognitive characteristics were entered as predictors and the psychopathology factor being regressed on all cognitive correlates simultaneously. Because of high correlations between some personality traits across measures (e.g., r=.47 between neuroticism and hopelessness, see supplementary table S2) and for ease of comparability with previous findings, separate models were conducted in which psychopathology factors were regressed on personality traits from each personality measure. Adjusted associations between correlates at 14 years and psychopathology factors at 16 years (see the second part of Table 4), showed that the P factor was predicted by high impulsivity, novelty-seeking, extraversion, hopelessness and neuroticism, and low agreeableness, conscientiousness and anxiety sensitivity. The P factor was also predicted by low spatial IQ, high delay discounting and poor response inhibition. The specific EXT (SU and low ADHD) factor was predicted by high novelty-seeking, sensation seeking and extraversion, and high verbal IQ and high

risk-taking (gambling task). Finally, the specific INT factor was associated with high neuroticism, hopelessness, conscientiousness, anxiety sensitivity, and low novelty-seeking and extraversion. The INT factor was also associated with an attentional bias toward negatively-valenced verbal stimuli. In these models, personality traits explained 8% to 15% of the variance of the P factor, 3% to 7% of the variance of the EXT (SU low ADHD) factor and 4% to 14% of the variance of the INT factor. Cognitive correlates explained 6%, 2% and 5% of the variance of the P factor, EXT (SU and low ADHD) factor and INT factor, respectively.

Discussion

The first objectives of the current study were to model the structure of psychopathology in a community sample of European adolescents, and to examine the stability of psychopathology symptoms from early to middle adolescence (14 to 16 years). Findings demonstrated that a general psychopathology (P) factor can be modelled in this community adolescent sample, as well as either a) two specific externalising and internalising factors or b) three specific ADHD/CD/ODD, substance use and internalising factors, providing further support for a spectrum and latent trait model of psychopathology (e.g., Lahey et al., 2012; 2015; Caspi et al., 2014; Murray et al., 2016). This study contributed to the literature by extending previous bifactor models to include eating disorders and a wider range of substance use symptoms. This study also showed that a bi-factor structure of psychopathology was stable across early-to-middle adolescence, a period characterised by substantial change and the onset of many disorders, but that the contributions made by different psychopathology symptoms to the P factor changed across development. Indeed, longitudinal factorial invariance analyses

conducted in the present sample showed that loadings for internalizing symptoms, as well as drug use and tobacco use, became stronger with age.

Of note, although like other studies we found that a bi-factor model of psychopathology, with either two or three specific factors fit the data well, our findings differ from previously reported P factor models in the following ways: a) externalizing symptoms loaded more strongly on the P factor in this study, rather than internalizing symptom, which has been the case for many studies modelling the P factor (*e.g.*, Lahey et al., 2012, Caspi et al., 2012); and b) when only two specific EXT and INT factors were modelled, the EXT factor included significant positive loadings for SU variables but negative loadings for ADHD and ODD. These discrepancies could reflect differences across sample demographics, measures used and/or developmental stage. Future studies on this sample could examine the structure of psychopathology using different indicators (*e.g.*, symptom scores) and test for factorial invariance across countries to aid in confirming these hypotheses. That said, this study joins the fast growing literature confirming that, regardless of symptoms/disorders measured, there is substantial variance shared amongst these that can be captured by a general psychopathology factor.

Interestingly, the P factor in this study accounted for the common variance across externalising symptoms that was previously attributed to a general externalising factor in another IMAGEN study focusing specifically on the structure of externalising symptoms (Castellanos-Ryan et al., 2014). This finding highlights that modeling the structure of psychopathology based on a broad range of symptoms may clarify the nature, antecedents, and implications of liabilities to multiple psychiatric problems, which may have been incompletely captured by narrower analyses (e.g., modelling just the EXT or

INT spectrums). Our results suggest that non substance-related externalising problems (*i.e.*, CD, ADHD, and ODD) may not have more in common with each other and with substance use problems than the general liability to psychopathology shared with internalising problems.

Another objective was to identify some of the common and unique personality and cognitive correlates of general and specific psychopathology factors. Findings showed that after controlling for multiple testing, common variance across psychopathology was generally related to most personality measures, with the exception of openness to experience and anxiety sensitivity, in theoretically expected ways and in line with previous findings. Namely, the P factor was associated positively with neuroticism, hopelessness, impulsivity, novelty-seeking and negatively with agreeableness and conscientiousness (Carragher et al., 2016; Caspi et al., 2014; Tackett et al., 2013). These associations remained largely unchanged after adjusting for other personality traits in the model, suggesting that the general liability to psychopathology may be characterised by a dysregulated personality profile involving high negative affect, low positive affect and poor behavioural control.

In terms of cognitive correlates, unadjusted findings also replicate the pattern of results suggesting that the P factor was associated with poor general cognitive functioning (Caspi et al., 2014). Adjusted associations showed that high delay discounting, poor response inhibition, and low performance IQ were uniquely associated with the general psychopathology factor in this sample of adolescents. These findings are consistent with those of Caspi et al. (2014) and Lahey et al. (2015) identifying low performance IQ and poor executive function as important correlates of a general psychopathology factor.

Delay discounting has not previously been examined as a correlate of a general liability to psychopathology within a spectrum or bi-factor methodology framework, but this finding echoes studies identifying poor delay discounting as an important non-specific risk factor for psychopathology (or trans-disease mechanism; *e.g.*, Bickel, Koffarnus, Moody & Wilson, 2014; Castellanos-Ryan et al., 2014).

Correlates of specific factors were also identified, with high sensation-seeking, high verbal IQ, and risk-taking being related to variance specific to substance misuse, and high neuroticism, hopelessness, anxiety-sensitivity, conscientiousness and agreeableness but low novelty seeking and extraversion, as well as an attentional bias toward negativelyvalenced verbal stimuli being associated with variance specific to internalizing disorders. These findings of dissociation in the cognitive profiles of specific substance use factors from general externalizing, or in this case psychopathology factors, are consistent with previous analyses of this and other adolescent samples (Castellanos-Ryan & Conrod, 2012; Castellanos-Ryan et al., 2014) showing that sensation-seeking and individual differences in reward responding and impulsive choice (as assessed by risk-taking on a gambling task) specifically predict vulnerability to substance use in adolescence, whereas measures of response inhibition or "impulsive action" predict variance in other externalizing problems and/or their co-occurrence. The result that neuroticism, hopelessness, anxiety-sensitivity and low extraversion are generally associated with internalizing symptoms is consistent with previous studies and reviews (Kahn et al., 2005; Kotov et al., 2010; Malouff, Thorsteinsson & Schutte, 2005; Naragon-Gainey, 2010). This is the first study that we know of to identify a cognitive correlate for the common variance across internalizing disorders in a bifactor framework. Attentional

biases toward negative stimuli have been consistently associated with depression (*e.g.*, Joorman & Quinn, 2014), but our study is the first to show that such biases represent a trans-diagnostic risk factor for depression, multiple types of anxious symptoms, and eating disorder symptoms - but not externalizing problems. Although this effect has been proposed in many heuristic models of psychopathology, it has been difficult to demonstrate until very recently, likely due to the lack of studies with sufficiently large and richly characterized samples.

The strengths of this study include the large sample size within a homogeneous population of 14-year olds assessed prospectively using a well-validated structured psychiatric assessment involving child and parent reports. This richly phenotyped sample also provides a unique opportunity to investigate the phenotypic structure of a variety of different forms of psychopathology and the cognitive correlates within this model. However, there are also a number of limitations to this population-based approach. First, although the sample is ethnically homogeneous (Caucasian), and thus findings could generalize only to a Caucasian adolescent population, there is heterogeneity in that adolescents are recruited from different cities in Germany, France, Ireland and the UK. In order to control for differences in these subsamples of adolescents, complex random effects analyses controlling for testing site as a cluster variable were conducted. However, it is important to note that findings on this combined sample may not relate to a specific population of reference in the usual sense. Future studies should test the factorial and other invariance across countries to determine how representative these findings are for each of the subsamples of adolescents. Second, although a wide range of psychopathology symptoms and measures were included in the current paper (including

eating disorders that had yet to be tested), the current study did not have any validated measures of psychosis, mania or schizophrenia (included in some shape or form in Carragher et al., 2016; Caspi et al., 2014; Forbush and Watson, 2013; Laceulle et al., 2015; Wright and Simms, 2015). Thus, while we call this dimension the P factor, to be consistent with other literature (e.g., Caspi et al., 2014), we also caution that without inclusion of information on thought disorder, the P factor presented in our final model is incomplete, and for the moment might be better referred to as a general behavioural/emotional dysregulation dimension. Additional follow-ups would help us to understand how these latent factors transform as adolescents transition to adulthood and may begin to experience psychiatric problems more typically seen in adulthood. Third, our analytic strategy did not allow us to model more specific subfactors of internalising problems (e.g., the distinction between "fear" and "distress"; Krueger, 1999). Future studies should investigate the relevance of these subfactors in a complete bi-factor framework, as we have done for externalising problems (e.g., Castellanos-Ryan et al., 2014). In addition, we note that that in the current analyses we did not include exactly the same externalizing indicators as those included in our previously reported externalizing factor structure on this same sample (Castellanos-Ryan et al., 2014), resulting in somewhat different findings (i.e., in the current analyses the model including a specific ADHD/CD/ODD factor fit well, but did not capture a significant amount of variance). In our previous analysis, more variable self and parent reported "screen" ratings of ADHD and CD symptoms, as assessed by the Strengths and Difficulties Questionnaire, were also included in the analysis, which were the indicators that loaded the strongest on a specific ADHD/CD factor. It appears that when more variable screen data from the whole

population are used (rather than diagnoses based on a screened sample), a general externalizing dimension can be meaningfully distinguished from an ADHD/CD-specific factor, including on cognitive and neural correlates. When focusing analyses on data representing likelihood of psychiatric diagnosis, this general externalizing factor is lost, and instead a general psychopathology factor is identified, on which ADHD and CD diagnoses dominate, but concurrent substance use and internalizing psychopathology is also captured.

Conclusions and clinical implications

The findings on the personality and cognitive correlates of the P factor (a) provide additional support for the criterion validity of a general psychopathology factor, and demonstrate that the general psychopathology factor is not a spurious artifact of measurement error, and (b) offer important clues as to the psychobiological processes that may underlie the general and specific factors of psychopathology. Indeed, the current study identifies common personality and cognitive correlates underlying different dimensions of psychopathology which can inform research on the etiology of mental disorders and eventually inform the development of new intervention strategies that better address comorbidity in clinical practice by targeting trans-diagnostic risk factors. The findings reported herein suggest that traits related to disinhibition/impulsivity, low agreeableness and high negative affect, as well as high delay discounting and poor response inhibition, might underlie common vulnerability to psychopathology. Finally, targets for specific patterns of psychopathology were confirmed or newly identified and overlap with some of the dimensions identified within the Research Domain Criteria framework (Insel et al., 2010). Specifically, targeting sensation seeking/reward

sensitivity/impulsive choice appears a relevant target for specific risk for substance misuse, and managing anxiety sensitivity/attentional biases to negative stimuli might be helpful in reducing risk for internalizing psychopathology. These conclusions perfectly aligne with new findings on trait-targeted interventions reported in randomized controlled trials (Conrod et al, 2013; O'Leary-Barrett et al, 2013; Olthuis, Watt, Mackinnon, & Stewart, 2014, 2015). These authors have all shown that cognitive behavioral interventions targeting these personality and cognitive profiles are effective in treating or preventing externalizing and internalizing psychopathology and substance misuse. It is worth noting that the trait-focused interventions across these studies all involved very different treatment delivery methods, including face-to-face and distance delivered (e.g., telephone) individual interventions, and school-based group interventions. This approach might also prove effective and efficient in reducing concurrent problems in general mental health clinics where clients often present with a variety of mental health concerns, and may not meet full diagnostic criteria. The dimensional approach to treatment presents advantages for these cases, where interventions can be offered rapidly, based on brief personality, cognitive or psychiatric screens, and might provide more clarity on primary source of impairment amongst many concurrent diagnoses. Furthermore, the dimensional approach has a number of practical and organizational advantages for service providers relative to the categorical approach as services can be arranged around limited set of key dimensions of risk that are relevant to multiple diagnostic categories (e.g., neuroticism, impulsivity and improving response inhibition). Finally, while there is much enthusiasm for a shift to more dimensional approaches in psychopathology research and treatment, there remain a number of very important avenues of investigation

that can only be addressed within the context of large and richly phenotyped and genotyped datasets, such as the question of how these dimensions interact to further influence risk for psychopathology. Despite relatively consistent findings on the bi-factor model of psychopathology, more studies are needed to understand environmental and genetic contributions to these risk dimensions, and their interactions. Both twin modeling and molecular genetic studies will help to identify the biological basis of these risk dimensions, as will large neuroimaging studies.

References

- Achenbach, T. M., & Edelbrock, C. S. (1981). Behavioral problems and competencies reported by parents of normal and disturbed children aged four through sixteen.

 Monographs of the Society of Research on Child Development, 46, 1-82.
- Asparouhov, T., & Muthén, B. (2005). Multivariate statistical modeling with survey data.
- Bar-Haim, Y., Lamy, D., Pergamin, L., Bakermans-Kranenburg, M. J., Van Ijzendoorn, M. H. (2007) Threat-related attentional bias in anxious and nonanxious individuals: A meta-analytic study. *Psychological Bulletin*. *133*(1):1-24.
- Bickel, W. K., Koffarnus, M. N., Moody, L., & Wilson, A. G. (2014). The behavioral and neuro-economic process of temporal discounting: A candidate behavioral marker of addiction. *Neuropharmacology*, 76 Pt B, 518-527.
- Carragher, N., Teesson, M., Sunderland, M., Newton, N.C., Krueger, R.F., Conrod, P.J., Barrett, E.L., Champion, K.E., Nair, N.K., & Slade, T. (2016). The structure of adolescent psychopathology: a symptom-level analysis. *Psychological Medicine*, 46, 981–994.
- Caspi, A., Houts, R. M., Belsky, D. W., Goldman-Mellor, S. J., Harrington, H., Israel, S., Meier, M. H., Ramrakha, S., Shalev, I., Poulton, R., & Moffitt, T. E. (2014). The P factor: one general psychopathology factor in the structure of psychiatric disorders? *Clinical Psychological Science*, 2(2), 119-137.
- Castellanos-Ryan, N., & Conrod, P. J. (2012). Personality and substance misuse:

 Evidence for a four factor model of vulnerability. In J. Vester, K. Brady, E.

 Strain, M. Galanter & P. J. Conrod (Eds.), *Drug Abuse and Addiction in Medical Illness* (Vol. 1 & 2.): Humana/Spring Press.

- Castellanos-Ryan, N., O'Leary-Barrett, M., Sully, L., & Conrod, P. J. (2013). Sensitivity and specificity of a brief personality screening instrument in predicting future substance use, emotional, and behavioral problems: 18-month predictive validity of the substance use risk profile scale. *Alcoholism: Clinical and Experimental Research*, *37*, E281-290.
- Castellanos-Ryan, N., Rubia, K., & Conrod, P. J. (2011). Response inhibition and reward response bias mediate the predictive relationships between impulsivity and sensation seeking and common and unique variance in conduct disorder and substance misuse. *Alcoholism: Clinical and Experimental Research*, *35*, 140-155.
- Castellanos-Ryan, N., Struve, M., Whelan, R., Banaschewski, T., Barker, G. J., Bokde,
 A. L., Bromberg, U., Buchel, C., Flor, H., Fauth-Buhler, M., Frouin, V., Gallinat,
 J., Gowland, P., Heinz, A., Lawrence, C., Martinot, J. L., Nees, F., Paus, T.,
 Pausova, Z., Rietschel, M., Robbins, T. W., Smolka, M. N., Schumann, G.,
 Garavan, H., Conrod, P. J., & Consortium, I. (2014). Neural and cognitive
 correlates of the common and specific variance across externalising problems in
 young adolescence. *American Journal of Psychiatry*, 171, 1310-1319.
- Cloninger, C. R., Przybeck, T. R., Svrakic, D. M., & Wetzel, R. D. (1994). *The*Temperament and Character Inventory (TCI): A guide to its development and use.

 Washington University, St. Louis: MI: Center for the Psychobiology of Personality.
- Conrod, P. J., O'Leary-Barrett, M., Newton, N., Topper, L., Castellanos-Ryan, N., Mackie, C., & Girard, A. (2013). Effectiveness of a selective, personality-targeted

- prevention program for adolescent alcohol use and misuse: a cluster randomized controlled trial. *JAMA Psychiatry*, 70, 334-342.
- Costa Jr, P. T., & McCrae, R. R. (1992). Normal Personality Assessment in Clinical Practice: The NEO Personality Inventory. *Psychological Assessment*, *4*, 5-13.
- Ferreri, F., Lapp, L. K., & Peretti, C. S. (2011). Current research on cognitive aspects of anxiety disorders. *Current Opinions in Psychiatry*, 24(1), 49-54.
- Forbush, K. T., & Watson, D. (2013). The structure of common and uncommon mental disorders. *Psychological Medicine*, *43*, 97–108.
- Goodman, A., Heiervang, E., Collishaw, S., & Goodman, R. (2011) The 'DAWBA bands' as an ordered-categorical measure of child mental health: description and validation in British and Norwegian samples. *Social Psychiatry and Psychiatric Epidemiology*, 46, 521-532.
- Groth-Marnat, G., & Baker, S. (2003). Digit Span as a measure of everyday attention: a study of ecological validity. *Perception and Motor Skills*, 97, 1209-1218.
- Hibell, B., Anderson, B., Bjarnason, T., Kokkevi, A., Morgan, M., & Narusk, A. (1997).The 1995 ESPAD report. Alcohol and other drug use among students in 26European countries. In T. S. C. f. I. o. A. a. O. Drugs (Ed.). Stockholm, Sweden.
- Hobson, C. W., Scott, S., & Rubia, K. (2011). Investigation of cool and hot executive function in ODD/CD independently of ADHD. *Journal of Child Psychology and Psychiatry*, 52, 1035–1043.
- Hu, L., & Bentler, P. M. (1999). Cutoff criteria for fit indexes in covariance structure analysis: Conventional criteria versus new alternatives. *Structural Equation Modeling*, 6, 1-55.

- Insel, T., Cuthbert, B., Garvey, M., Heinssen, R., Pine, D. S., Quinn, K., Sanislow, C., & Wang, P. (2010). Research domain criteria (RDoC): toward a new classification framework for research on mental disorders. *American Journal of Psychiatry*, 167, 748-751.
- Kaufman, J. N., Ross, T. J., Stein, E. A., & Garavan, H. (2003). Cingulate hypoactivity in cocaine users during a GONOGO task as revealed by event-related functional magnetic resonance imaging. *Journal of Neuroscience*, 23, 7839–7843.
- Kendler, K. S., Prescott, C. A., Myers, J., & Neale, M. C. (2003). The structure of genetic and environmental risk factors for common psychiatric and substance use disorders in men and women. *Archives of General Psychiatry*, 60, 929-937.
- Kessler, R. C., Berglund, P., Demler, O., Jin, R., Merikangas, K. R., & Walters, E. E. (2005). Lifetime prevalence and age-of-onset distributions of DSM-IV disorders in the National Comorbidity Survey Replication. *Archives of General Psychiatry*, 62, 593-602.
- Kahn, A. M., Jacobson, K. C., Gardner, C. O., Prescott, C. A., & Kendler, K. S. (2005).
 Personality and comorbidity of common psychiatric disorders. British Journal of Psychiatry, 186, 190–196.
- Kirby, K. N., Petry, N. M., & Bickel, W. K. (1999). Heroin addicts have higher discount rates for delayed rewards than non-drug-using controls. *Journal of Experimental Psychology General*, 128, 78-87.
- Kotov, R., Gamez, W., Schmidt, F., & Watson, D. (2010). Linking "big" personality traits to anxiety, depressive, and substance use disorders: a meta-analysis.

 *Psychological Bulletin, 136(5), 768-821.

- Kotov, R. Leong, S.H., Mojtabai, R., Eckardt Erlanger, A.C, Fochtmann, L.J.,
 Constantino, E., Carlson, G.A., & Bromet, E.J. (2013) Boundaries of
 Schizoaffective Disorder Revisiting Kraepelin. *JAMA Psychiatry*.70(12), 1276-1286.
- Krank, M., Stewart, S. H., O'Connor, R., Woicik, P. B., Wall, A. M., & Conrod, P. J. (2011). Structural, concurrent, and predictive validity of the Substance Use Risk Profile Scale in early adolescence. *Addictive Behaviors*, *36*, 37-46.
- Krueger, R. F. (1999). The structure of common mental disorders. *Archives of General Psychiatry*, *56*, 921-926.
- Kuo, W. H., Gallo, J. J., & Eaton, W. W. (2004). Hopelessness, depression, substance disorder, and suicidality--a 13-year community-based study. *Social Psychiatry* and Psychiatric Epidemiology, 39(6), 497-501.
- Laceulle, O.M., Vollebergh, W.A.M., & Ormel, J. (in press). The Structure of
 Psychopathology in Adolescence: Replication of a General Psychopathology
 Factor in the TRAILS Study. *Clinical Psychological Science*, *DOI*10.1177/2167702614560750,1–11.
- Lahey, B. B., Applegate, B., Hakes, J. K., Zald, D. H., Hariri, A. R., & Rathouz, P. J. (2012). Is there a general factor of prevalent psychopathology during adulthood? *Journal of Abnormal Psychology*, 121, 971-977.
- Lahey, B.B., Rathouz, P.J., Keenan, K., Stepp, S.D., Loeber, R., & Hipwell, A.E. (2015). Criterion validity of the general factor of psychopathology in a prospective study of girls. *Journal of Child Psychology and Psychiatry*, 56, 415-422.

- Lahey, B. B., Van Hulle, C. A., Singh, A. L., Waldman, I. D., & Rathouz, P. J. (2011).

 Higher-order genetic and environmental structure of prevalent forms of child and adolescent psychopathology. *Archives of General Psychiatry*, 68, 181-189.
- Lijffijt, M., Kenemans, J. L., Verbaten, M. N., & van Engeland, H. (2005). A metaanalytic review of stopping performance in attention-deficit/hyperactivity disorder: Deficient inhibitory motor control? *Journal of Abnormal Psychology*, 114, 216–222.
- Lockenhoff, C. E., Terracciano, A., Patriciu, N. S., Eaton, W. W., & Costa, P. T., Jr. (2009). Self-reported extremely adverse life events and longitudinal changes in five-factor model personality traits in an urban sample. *Journal of Trauma and Stress*, 22, 53-59.
- Malouff, J.M., Thorsteinsson, E. B., & Schutte, N. S. (2005). The relationship between the five-factor model of personality and symptoms of clinical disorders: A meta-analysis. *Journal of Psychopathology and Behavioural Assessment*, 27, 101-114.
- Murray, A.L., Eisner, M., & Ribeaud, D. (2016). The Development of the General Factor of Psychopathology 'p Factor' Through Childhood and Adolescence. *Journal of Abnormal Child Psychology*, DOI 10.1007/s10802-016-0132-1
- Muthén, L. K., & Muthén, B. O. (2010). *Mplus User's Guide* (6 ed.). Los Angeles: Muthén & Muthén.
- Nagl, M., Jacobi, C., Paul, M., Beesdo-Baum, K., Hofler, M., Lieb, R., & Wittchen, H. U.
 (2016). Prevalence, incidence, and natural course of anorexia and bulimia nervosa among adolescents and young adults. *European Child & Adolescent Psychiatry*.
 Advance online publication.

- Naragon-Gainey, K. (2010). Meta-analysis of the relations of anxiety sensitivity to the depressive and anxiety disorders. *Psychological Bulletin*, *136*, 128-150.
- Nelson, C. A., Furtado, E. A., Fox, N. A., & Zeanah, C. H. (2009). The deprived human brain: developmental deficits among institutionalized Romanian children—and later improvements—strengthen the case for individualized care. *American Scientist*, 97, 222–229.
- Newman, J. P., & Wallace, J. F. (1993). Diverse pathways to deficient self-regulation: Implications for disinhibitory psychopathology in children. *Clinical Psychology Review*, *13*, 699-720.
- Noordhof, A., Krueger, R. F., Ormel, J., Oldehinkel, A.J., & Hartman, C.A. (2015).

 Integrating Autism-Related Symptoms into the Dimensional Internalizing and Externalizing Model of Psychopathology. The TRAILS Study. *Journal of Abnormal Psychology*, 43, 577-587.
- O'Leary-Barrett, M., Topper, L., Al-Khudhairy, N., Pihl, R. O., Castellanos-Ryan, N., Mackie, C. J., & Conrod, P. J. (2013). Two-year impact of personality-targeted, teacher-delivered interventions on youth internalising and externalising problems: a cluster-randomized trial. *Journal of the American Academy of Child and Adolescent Psychiatry*, 52, 911-920.
- Olley, A., Malhi, G., & Sachdev, P. (2007). Memory and executive functioning in obsessive-compulsive disorder: a selective review. *Journal of Affective Disorders*, 104(1-3), 15-23.

- Olthuis, J. V., Watt, M. C., Mackinnon, S. P., & Stewart, S. H. (2014). Telephone-delivered cognitive behavioral therapy for high anxiety sensitivity: a randomized controlled trial. *Journal of Consulting and Clinical Psychology*, 82, 1005-1022.
- Olthuis, J. V., Watt, M. C., Mackinnon, S. P., & Stewart, S. H. (2015). CBT for high anxiety sensitivity: Alcohol outcomes. *Addictive Behaviors*, 46C, 19-24.
- Oosterlaan, J., & Sergeant, J. A. (1998). Response inhibition and response re-engagement in attentiondeficit/hyperactivity disorder, disruptive, anxious and normal children. Behavioral Brain Research, 94, 33–43.
- Patalay, P., Fonagy, P., Deighton, J., Belsky, J., Vostanis, P., & Wolpert, M. (2015). A general psychopathology factor in early adolescence. *British Journal of Psychiatry*, 207, 15-22.
- Peckham, A. D., McHugh, R. K., & Otto, M. W. (2010). A meta-analysis of the magnitude of biased attention in depression. *Depression and Anxiety*, 27(12), 1135-1142.
- Peeters, M., Monshouwer, K., Janssen, T., Wiers, R. W., & Vollebergh, W. A. (2014).

 Working memory and alcohol use in at-risk adolescents: A 2-year follow-up.

 Alcoholism: Clinical and Experimental Research, 38, 1176–1183.
- Pettersson, E., Lahey, B., Larsson, H., Lundström, S., & Lichtenstein, P. (2015). The general factor of psychopathology in childhood and maladaptive outcomes in young adulthood: a Swedish population twin study. *Behavior Genetics*, 45, 679-679.

- Polak, A. R., Witteveen, A. B., Reitsma, J. B., & Olff, M. (2012). The role of executive function in posttraumatic stress disorder: a systematic review. *Journal of Affective Disorders*, *141*(1), 11-21.
- Reynolds, C. R. (1997). Forward and backward memory span should not be combined for clinical analysis. *Archives of Clinical Neuropsychology*, *12*, 29-40.
- Saunders, J. B., Aasland, O. G., Babor, T. F., de la Fuente, J. R., & Grant, M. (1993).

 Development of the Alcohol Use Disorders Identification Test (AUDIT): WHO collaborative project on early detection of persons with harmful alcohol consumption--II. *Addiction*, 88, 791-804.
- Schumann, G., Loth, E., Banaschewski, T., Barbot, A., Barker, G., Buchel, C., Conrod, P.
 J., Dalley, J. W., Flor, H., Gallinat, J., Garavan, H., Heinz, A., Itterman, B.,
 Lathrop, M., Mallik, C., Mann, K., Martinot, J. L., Paus, T., Poline, J. B.,
 Robbins, T. W., Rietschel, M., Reed, L., Smolka, M., Spanagel, R., Speiser, C.,
 Stephens, D. N., Strohle, A., Struve, M., & consortium, I. (2010). The IMAGEN
 study: reinforcement-related behaviour in normal brain function and
 psychopathology. *Molecular Psychiatry*, 15, 1128-1139.
- Snyder, H. R. (2013). Major depressive disorder is associated with broad impairments on neuropsychological measures of executive function: a meta-analysis and review.

 *Psychological Bulletin, 139(1), 81-132.
- Stice E, Marti CN, Shaw H, Jaconis M (2009) An 8-year longitudinal study of the natural history of threshold, subthreshold, and partial eating disorders from a community sample of adolescents. *Journal of Abnormal Psychology*, 118:587–597

- Tackett, J. L., Lahey, B. B., van Hulle, C., Waldman, I., Krueger, R. F., & Rathouz, P. J. (2013). Common genetic influences on negative emotionality and a general psychopathology factor in childhood and adolescence. *Journal of Abnormal Psychology*, 122, 1142-1153.
- Wagner, S., Muller, C., Helmreich, I., Huss, M., & Tadic, A. (2015). A meta-analysis of cognitive functions in children and adolescents with major depressive disorder.

 European Child & Adolescent Psychiatry, 24(1), 5-19.
- Wechsler, D. (2003). Wechsler Intelligence Scale for Children—Fourth edition (WISC-IV). San Antonio, TX: The Psychological Association.
- Widiger, T. A., & Costa, P. T., Jr. (1994). Personality and personality disorders. *Journal of Abnormal Psychology*, 103(1), 78-91.
- Woicik, P. A., Stewart, S. H., Pihl, R. O., & Conrod, P. J. (2009). The Substance Use Risk Profile Scale: a scale measuring traits linked to reinforcement-specific substance use profiles. *Addictive Behaviors*, *34*, 1042-1055.
- Wright, A. G. C., & Simms, L. J. (2015) A metastructural model of mental disorders and pathological personality traits. *Psychological Medicine*, 45, 2309–2319
- Young, S. E., Friedman, N. P., Miyake, A., Willcutt, E. G., Corley, R. P., Haberstick, B.
 C., & Hewitt, J. K. (2009). Behavioral disinhibition: Liability for externalising spectrum disorders and its genetic and environmental relation to response inhibition across adolescence. *Journal of Abnormal Psychology*, 118, 117–130

Table 1. Descriptives and correlations between psychopathology symptoms at 16 years and demographic measures.

		1	2	3	4	5	6	7	8	9	10	11	12
1.	ADHD band	1.00											
2.	CD band	0.39	1.00										
3.	ODD band	0.43	0.41	1.00									
4.	Alcohol problems	0.02	0.20	0.06	1.00								
5.	Number of drugs	0.13	0.26	0.06	0.35	1.00							
6.	Smoking Frequency	0.16	0.30	0.10	0.37	0.56	1.00						
7.	Gen Anxiety band	0.19	0.21	0.10	0.07	0.15	0.12	1.00					
8.	Depression band	0.30	0.31	0.15	0.12	0.22	0.18	0.49	1.00				
9.	Eating band	0.08	0.12	0.10	0.09	0.08	0.09	0.32	0.27	1.00			
10.	OCD band	0.25	0.21	-0.01	0.05	0.13	0.09	0.45	0.37	0.23	1.00		
11.	Panic and Phobias	0.24	0.17	0.16	0.01	0.05	0.05	0.44	0.37	0.28	0.33	1.00	
12.	Social Phobia	0.10	0.10	0.10	-0.03	0.02	0.00	0.27	0.19	0.18	0.24	0.27	1.00
13.	Gender	-0.08	0.00	0.03	-0.02	-0.06	0.01	0.22	0.18	0.31	0.10	0.26	0.09
14.	Language (Eng vs Other)	-0.05	-0.05	0.02	0.08	-0.02	-0.15	-0.02	-0.01	0.09	0.01	-0.08	0.11
Mea	n	0.41	1.31	0.54	0.62	0.37	1.79	0.65	0.63	0.58	0.38	0.31	1.18
SD		0.79	0.71	1.05	1.19	0.77	2.31	0.99	0.93	0.73	0.66	0.40	0.72
Rang	ge	0-5	0-5	0-5	0-8	0-7	0-7	0-5	0-5	0-5	0-5	0-5	0-5
Likely diagnosis or case±, % (n)		3 (52)	4 (71)	7 (120)	14 (218)	26 (415)	48 (768)	8 (136)	4 (67)	1 (13)	1 (16)	1 (16)	2 (27)

Bold indicates significant at p<.05. ADHD – Attention Deficit Hyperactivity Disorder; CD – Conduct Disorder; ODD – Oppositional Defiant Disorder; Gen – general; OCD Obsessive Compulsive Disorder. *Percentages based on observed (no–n-missing) data at 16 years (N=1603).

Table 2. Fit indices for structural equation models of psychopathology at 16 years

Model	X^2	df	CFI	RMSEA	SRMR	BIC	Adj BIC
Model 1: one factor	2448.59	54	.00	.141	.080	50108.17	49901.66
Model 2a: Correlated two subfactors (EXT, INT)	1032.37	53	.59	.091	.066	49189.26	49355.97
Model 2b: Correlated three subfactors (EXT, SU, INT)	216.64	51	.93	.038	.035	48994.67	48778.62
Model 3a: Bi-factor two specific (EXT, INT)	175.98	42	.94	.038	.022	48937.07	48692.43
Model 3b: Bi-factor three specific (EXT, SU, INT)	277.59	42	.90	.050	.026	48957.72	48713.08
Model 3a':revised bi-factor two specific factors with correlation	170.91	41	.94	0.38	.021	48938.24	48690.42
Model 3b': revised Bifactor three specific factors with correlations	170.52	40	.94	.038	.021	48937.92	48690.10

Tests of goodness of fit included the Chi-square and Comparative Fit Indices (CFI), the Standardized Root Mean Square Residual (SRMR) and the Root Mean Square Error of Approximation (RMSEA). Hu and Bentler (1999) suggest the following guidelines for interpreting Goodness-of-Fit Indices: SRMR and values close to or below .08, RMSEA values close to or below .06 and CFI close to or above .90 indicate acceptable model fit. Models were compared using the Bayesian information criterion (BIC) and the sample adjusted (adj) BIC, frequently used to compare non-nested models. Smaller values on both these measures indicate a better fitting model while penalizing for increasing model complexity.

Table 3. Standardized factor loadings and significance levels for the bi-factor model 3a at 16 and correlations between factors at 14 and 16 years

	P facto	r at 16	EXT (SU ADHD) fa		INT factor at 16		
	Loading	p	Loading	P	Loading	P	
ADHD band	0.64	0.000	-0.25	0.008			
CD band	0.65	0.000	0.01	0.994			
ODD band	0.58	0.000	-0.24	0.073			
Drinking related problems	0.26	0.010	0.43	0.000			
Number of drugs used	0.42	0.000	0.60	0.000			
Smoking frequency	0.45	0.000	0.60	0.000			
General Anxiety band	0.32	0.000			0.60	0.000	
Depression band	0.46	0.000			0.46	0.000	
Social Phobia band	0.13	0.007			0.34	0.000	
Panic and other phobias	0.27	0.002			0.56	0.000	
Eating Disorder band	0.19	0.000			0.28	0.000	
OCD band	0.29	0.000			0.57	0.000	
Correlations between factors	r	p	r	P	r	p	
P factor at 14	0.73	0.000	0.03	0.865	-0.09	0.033	
EXT (SU) factor at 14	0.00	0.988	0.62	0.000	0.03	0.660	
INT factor at 14	0.03	0.059	0.02	0.543	0.50	0.000	

Legend: P – psychopathology; ADHD – Attention Deficit Hyperactivity Disorder; CD – Conduct Disorder; SU – Substance use; EXT - externalizing psychopathology; INT – internalizing psychopathology; band: computer-generated likelihood that the individual suffers from that disorder using DSM-IV-TR criteria; SP – self reported; PR – parent reported; Loading – estimated standardized factor loadings; p – 2-tailed significance level.

Table 4. General and specific associations between psychopathology factors (at 16 years) and personality and cognition (at 14 years)

	Unadjusted Associations (bivariate correlations)								Adjusted Associations (regression paths)						
_	P factor		EXT (SU and low ADHD) factor		' INT	INT factor		P factor		EXT (SU and low ADHD) factor		INT factor			
	r	p	r	P	r	p		β	p	β	p	β	p		
Personality Measures a	at 14 years:	:													
Anxiety Sensitivity	-0.03	0.173	-0.03	0.321	0.14	0.003		-0.06	0.034	-0.03	0.331	0.10	0.046		
Hopelessness	0.18	0.000	0.03	0.624	0.22	0.000		0.14	0.000	0.05	0.404	0.20	0.000		
Impulsivity	0.34	0.000	0.05	0.320	0.06	0.057		0.31	0.000	0.04	0.602	0.02	0.569		
Sensation-seeking	0.12	0.007	0.13	0.003	-0.06	0.279		0.08	0.064	0.14	0.001	-0.02	0.683		
							R^2	0.15	0.000	0.03	0.079	0.06	0.009		
Novelty-seeking	0.27	0.000	0.26	0.000	-0.15	0.000		0.29	0.000	0.27	0.000	-0.19	0.000		
							R^2	0.08	0.011	0.07	0.041	0.04	0.007		
Neuroticism	0.18	0.000	0.03	0.519	0.33	0.000		0.12	0.000	0.02	0.734	0.34	0.000		
Extraversion	0.08	0.001	0.11	0.000	-0.17	0.001		0.19	0.000	0.14	0.001	-0.09	0.011		
Openness	-0.04	0.420	0.05	0.269	0.04	0.310		0.01	0.816	0.06	0.082	0.01	0.815		
Agreeableness	-0.27	0.000	-0.04	0.550	-0.05	0.290		-0.22	0.000	-0.06	0.378	0.03	0.499		
Conscientiousness	-0.25	0.000	-0.07	0.351	0.06	0.001		-0.18	0.000	-0.08	0.269	0.15	0.000		
							R^2	0.14	0.000	0.03	0.246	0.14	0.000		
Cognitive Measures at	14 years:														
Verbal IQ	-0.10	0.014	0.12	0.008	0.18	0.120		-0.03	0.484	0.13	0.027	0.19	0.170		
Performance IQ	-0.14	0.000	0.04	0.232	0.06	0.077		-0.07	0.025	-0.03	0.540	-0.03	0.478		
DS forward	-0.04	0.039	0.04	0.401	0.06	0.043		0.05	0.249	-0.01	0.955	0.04	0.160		
DS backward	-0.10	0.018	0.05	0.117	0.00	0.998		-0.07	0.216	0.03	0.454	-0.06	0.262		
Delay discounting	0.14	0.000	0.00	0.942	0.01	0.613		0.10	0.015	0.02	0.572	0.04	0.306		
Risk-taking (CGT)	0.07	0.013	0.06	0.045	-0.01	0.843		0.06	0.093	0.06	0.040	0.00	0.990		
RI (commission)	0.14	0.000	-0.01	0.880	-0.04	0.336		0.09	0.041	0.03	0.444	0.00	0.975		
AAB (Pos-Neg Om)	-0.07	0.068	0.03	0.328	-0.10	0.026		-0.07	0.050	0.04	0.213	-0.09	0.012		
Spatial WM	0.10	0.000	-0.05	0.232	-0.08	0.205		0.02	0.568	-0.02	0.572	-0.05	0.402		
							R^2	0.06	0.003	0.02	0.098	0.05	0.427		

P: Psychopathology; EXT: externalizing; SU: substance use; INT: internalizing; IQ: Intelligence Quotient; DS: digit span; CGT: Cambridge Gambling Task; RI: Response Inhibition; AAB:

Affective Attentional bias; Om: Omission errors; WM: Working Memory. Bold indicates significant after controlling for multiple testing. Standardized coefficients provided: r: correlations and

 β : standardised regression paths or betas; R²: variance explained in adjusted association models only. Model Fit: X² (132)=643.51, CFI=.92, RMSEA=.042, SRMR=.023, for model with all personality correlates (unadjusted associations); X² (78)=423.98, CFI=.93, RMSEA=.045, SRMR=.025, for model with adjusted associations for SURPS personality traits (regression paths); X² (51)=598.23, CFI=.93, RMSEA=.070, SRMR=.027, for model with TCI novelty-seeking only (regression path); X² (87)=697.25, CFI=.92, RMSEA=.056, SRMR=.027, for model with adjusted associations for NEO-FFI personality traits (regression paths); X² (124)=350.67, CFI=.95, RMSEA=.029; SRMR=.022, for model with all cognitive measures (unadjusted associations); X² (126)=357.89, CFI=.93, RMSEA=.029; SRMR=.022, for model with all cognitive measures (adjusted associations).

Figure 1. Bi-factor model 3a of Psychopathology

