

CAHIER 8131

Investment, Taxation and Econometric Policy  
Evaluation : Some Evidence on the  
Lucas Critique\*

by

Jean-Marie Dufour

April 1979  
May 1981  
May 1982

\*The author wishes to thank Arnold Zellner, Nicholas Kiefer, Robert E. Lucas, Jr, Houston Stokes and the members of the University of Chicago Econometrics and Statistics Colloquium for many helpful comments. This work was supported by Grants from the Social Sciences and Humanities Research Council of Canada and by a F.C.A.C. Grant from the Government of Québec. It is a revised version of Chapter 3 of the author's Ph.D. Dissertation at the University of Chicago (1979).

---

Ce cahier est publié conjointement par le Département de science économique et par le Centre de recherche en développement économique de l'Université de Montréal.

Cette étude a été publiée grâce à une subvention du fonds FCAC pour l'aide et le soutien à la recherche.

## ABSTRACT

The aggregate investment schedule is a relationship for which the issue of stability over time is of special importance for macroeconomic policy. In particular, one may wish to use this schedule to study the impact of various policy measures such as changes in the nominal corporate tax rate, changes in depreciation allowances, investment tax credits, etc. An ingenious formulation of an investment function making possible such studies is due to Hall and Jorgenson (1967). The importance of using, for such policy simulations, a model which exhibits a good stability over time is easy to understand. In particular, the parameters should be invariant with respect to the policy changes themselves. This point has been forcefully stressed by Lucas (1976). This author argues the since parameters in econometric relationships reflect economic agent's decision rules and these integrate knowledge about policies, changes in policies are likely to induce shifts in the parameters. Lucas describes three cases where such phenomena could be observed, one of which involves the Hall-Jorgenson (1967) model of investment demand (and taxation). In particular, as regards the impact of investment tax credits with this model, Lucas makes two kinds of prediction: first, if the model is implemented under an assumption of static expectations (vs rational expectations) and estimated from a period during which policy rules changes appreciably, it is likely to exhibit parameter instability related to these policy changes; second, the impact of tax credits is likely to be heavily under-estimated.

The paper presents empirical evidence on both these effects on the basis of a version of this model estimated from U.S. data (1956-1972) by Gordon and Jorgenson (1976) and extensively used in policy simulations; fluctuations of investment tax credits over the period considered suggest the Lucas effects are likely to be observable in this case. For this purpose, we use recursive stability analysis, an exploratory methodology introduced by Brown, Durbin and Evans (1975) and extended by Dufour (1979, 1982). This method is based on a process of recursive estimation of the model considered with, among other things, an analysis of associated prediction errors; it has the advantage of making very weak assumptions concerning the type of instability to be detected and indicates the direction of the prediction errors induced. The main finding is a discontinuity associated with the first imposition of the tax credit (1964-66); furthermore, the observed shift induced a strong phenomenon of unprediction of investment, thus bringing support for Lucas' hypothesis.

## RESUME

La fonction d'investissement est une relation dont la stabilité temporelle a une importance considérable pour les politiques macroéconomiques. En particulier, on peut se servir d'une telle relation afin d'étudier l'impact de diverses mesures telles des changements dans les taux de taxation du revenu des compagnies ou dans les taux de dépréciation, les crédits d'impôt à l'investissement, etc. Une formulation particulièrement ingénieuse de cette relation et permettant de telles études a été proposée notamment par Hall et Jorgenson (1967).

Il est facile de comprendre l'importance, pour de telles simulations, d'utiliser un modèle dont les coefficients peuvent être considérés comme stables dans le temps. Lucas (1976) a spécialement insisté sur ce problème. Cet auteur suggère que des changements dans les politiques peuvent induire des changements dans les coefficients: ceci provient du fait que les coefficients des modèles économétriques reflètent des règles de décision optimales qui tiennent compte des politiques gouvernementales. Lucas décrit trois cas où on est susceptible d'observer de tels phénomènes: l'un de ces cas est le modèle de demande d'investissement de Hall et Jorgenson (1967). Considérant le cas particulier des crédits d'investissement, Lucas fait deux types de prévision: premièrement, si le modèle est estimé sous une hypothèse d'attentes statiques et à partir de données d'une période au cours de laquelle la règle de politique pertinente a changé, on peut s'attendre à ce que le modèle démontre une instabilité reliée à ces changements; deuxièmement, l'impact d'un crédit d'investissement temporaire sera sous-estimé.

Dans ce contexte, nous présentons des résultats sur ces deux effets à partir d'une version du modèle originalement estimée avec des données américaines (1956-72) par Gordon et Jorgenson (1976) et qui a été utilisée pour faire des simulations de politiques. Comme on peut observer plusieurs fluctuations dans la politique touchant les crédits d'investissement, nous anticipons d'observer les effets suggérés par Lucas dans ce cas. Pour cette fin, nous utilisons "l'analyse récursive de la stabilité" telle qu'introduite par Brown, Durbin et Evans (1975) et généralisée par Dufour (1979, 1982). Cette méthode est basée sur un processus récursif d'estimation et l'analyse de diverses statistiques résultantes: les avantages de l'approche sont qu'aucune hypothèse n'est faite sur le type d'instabilité à détecter et qu'elle fournit des indications sur la direction des erreurs de prévision. A l'aide de cette méthode, nous avons détecté une discontinuité qu'on peut associer très précisément avec la première imposition du crédit d'investissement (1964-66); de plus, le changement observé produit un phénomène important de sous-prévision de l'investissement. Ces deux observations semblent donc supporter l'hypothèse de Lucas.

## 1. Introduction

The aggregate investment schedule is a relationship for which the issue of stability over time is of special importance for macroeconomic policy. In particular, one may wish to use this schedule to study the impact of various policy measures such as changes in the nominal corporate tax rate, changes in depreciation allowances, investment tax credits, etc. An ingenious formulation of an investment function making possible such studies is due to Hall and Jorgenson (1967). Also, a particularly extensive simulation based on this model, aimed at studying the impact of investment tax credits in the United States over the period 1960-1985 was made by Gordon and Jorgenson (1976).

The importance of using, for such policy simulations, a model which exhibits a good stability over time is easy to understand. In particular, the parameters should be invariant with respect to the policy changes themselves. This point has been forcefully stressed by Lucas (1976). This author argues that since parameters in econometric relationships reflect economic agents' decision rules and these integrate knowledge about policies, changes in policies are likely to induce shifts in the parameters. Lucas describes three cases where such phenomena could be observed : the first one is linked to the aggregate consumption function, the second to the Hall-Jorgenson (1967) model of investment demand (and taxation) and the third to the Phillips curve. We will concentrate here on the second case in order to obtain some empirical evidence on the instability issue involved.

The argument of Lucas is that the effect of a change in the rate of an investment tax credit depends crucially on expectations concerning future changes in this rate. The impact of a given change in the rate of the

tax credit will differ depending on expectations about future changes in the rate. Or, in other words, the response coefficient to a change in the rate of the tax credit depends on expectations about future changes in this rate. In particular, after developing a simple investment model, Lucas shows that the impact of a given change may be substantially bigger if it is thought to be transitory rather than permanent (once-and-for all)<sup>1</sup>. Thus, assuming the changes in the investment tax credit were considered permanent by the relevant economic agents while they were in fact thought transitory, may lead to appreciably underestimate the impact of the tax credit.

At the empirical level, if one wants to forecast accurately the effect of a proposed change in the tax credit, it is necessary that:

- (1) the correct assumptions concerning expectations about future changes in the tax credit (following the proposed change) be used;
- (2) the model be specified and estimated using the correct expectational assumptions over the historical period used for estimation.

Hall and Jorgenson (1967) as well as Gordon and Jorgenson (1976) assumed changes in tax rates were viewed as permanent. We will devote our attention here primarily to the second study. During the sampling period used for the estimation of the investment function of Gordon and Jorgenson (1956-72), five major changes in the tax credit took place. The tax credit was originally introduced to stimulate investment in 1962. Then "the effectiveness of the tax credit was increased substantially in 1964

---

<sup>1</sup>More precisely, assuming the tax credit follows a Markovian scheme (which includes as special cases a permanent credit and a frequently imposed but always transitory credit), Lucas (1976) shows that the impact of the tax credit on investment can be much bigger if it is thought transitory rather than permanent. Indeed, under reasonable values of the parameters, the ratio of effects may be in the range of 4 to 7.

with the repeal of the Long Amendment<sup>1</sup>. The investment tax credit was suspended in 1966-67 and repealed in 1969 in order to reduce the level of investment. The tax credit was re-enacted in 1971 to stimulate investment expenditures<sup>2</sup>. Thus, given this apparent instability of policy, it would not be surprising (if we follow Lucas' argument) to observe parameter instability in the Gordon-Jorgenson model (unless expectations effectively obeyed the scheme implicitly assumed by Hall and Jorgenson). Also, if we estimate the model recursively (adding observations gradually), as suggested by Brown, Durbin and Evans (1975) and Dufour (1979, 1982), we would expect the introduction of the investment tax credit to be associated with under-predictions of investment expenditures, since it is argued that the assumption of static expectations is likely to lead to underestimate the impact of the tax credit.

The precise model considered by Gordon and Jorgenson (1976) is of the form (for quarterly data) :

$$(1.1) \quad IPDE58_t = \alpha + \delta K_t + \sum_{i=0}^6 \beta_i V_{t-i} + u_t .$$

$IPDE58_t$  is real investment (1958 dollars) in producer's durable equipment (during period  $t$ ),  $K_t$  is gross, beginning of period, real capital stock of producer's durable equipment,  $V_t$  is a proxy for desired capital stock

---

<sup>1</sup>The Long Amendment forbade to use for depreciation purposes that part of the cost of a capital asset financed by the tax credit.

<sup>2</sup>From Gordon and Jorgenson (1971, p. 278). We list the "effective tax credits" (1961-72), as measured by these authors, in Table 7. The "effective tax credit" could be non-zero for periods longer than the nominal credit because, once the credit was suspended or repealed, firms could still use a credit to which they were entitled but did not use when it was in force.

given by

$$(1.2) \quad V_t = (\text{PGNP}_{t-2}) (\text{GNP58}_{t-1}) / C_{t-2},$$

where  $\text{GNP58}_t$  is real gross national product (1958 dollars)<sup>1</sup>,  $\text{PGNP}_t$  is the GNP price deflator and  $C_t$  is the rental cost of capital;  $u_t$  is a random disturbance. The cost of capital  $C_t$  is defined as

$$(1.3) \quad C_t = \text{PIPDE}_t \left[ 0.138 + R_t (1 - U_t) \right] \left[ 1 - U_t Z_t - \text{TC}_t + Y_t Z_t \text{TC}_t U_t \right] / (1 - U_t),$$

where  $\text{PIPDE}_t$  is the price deflator for investment in producer's durable equipment, 0.138 is the depreciation rate on producer's durable equipment as calculated by Christensen and Jorgenson (1969),  $U_t$  is the nominal corporate tax rate,  $R_t$  is the interest rate on new issues of High-Grade corporate bonds,  $Z_t$  is the present discounted value of depreciation allowances,  $\text{TC}_t$  is the effective tax credit and  $Y_t$  equals one during those years in which the Long Amendment applied and zero otherwise.

In order to estimate this model, Gordon and Jorgenson (1976) used a second degree Almon polynomial lag structure constrained to pass through zero after seven periods, i.e. it was assumed that

$$(1.4) \quad \beta_i = a_0 - a_1 i - a_2 i^2, \quad i = 0, 1, \dots, 7$$

with  $\beta_7 = a_0 - 7a_1 - 49a_2 = 0$ , so that there are effectively only two free parameters in the distributed lag over  $V_t$ . Under these conditions, the relation to be estimated takes the form:

$$(1.5) \quad \text{IPDE58}_t = \alpha + \delta K_t + a_1 W_{1t} + a_2 W_{2t} + u_t,$$

<sup>1</sup>IPDE58<sub>t</sub> and GNP58<sub>t</sub> are seasonally adjusted and measured at annual rates

where

$$(1.6) \quad W_{1t} = \sum_{i=0}^6 (7-i)v_{t-i} \quad , \quad W_{2t} = \sum_{i=0}^6 (49-i^2)v_{t-i}$$

Furthermore, since the original Durbin-Watson statistic was 0.7554, a first-order autoregressive transform was used ( $\hat{\rho} = .6223$ ) and the following result was obtained (based on the period 1956-72<sup>1</sup>):

$$(1.7) \quad \begin{aligned} \text{IPDE58}_t = & -9.656 + 0.0572 K_t + 0.00181 v_t + 0.00218 v_{t-1} \\ & (1.522) \quad (0.0163) \quad (0.00071) \quad (0.00033) \\ & + 0.00233 v_{t-2} + 0.00228 v_{t-3} + 0.00202 v_{t-4} \\ & (0.00019) \quad (0.00031) \quad (0.00038) \\ & + 0.00156 v_{t-5} + 0.00088 v_{t-6} \\ & (0.00036) \quad (0.00023) \\ R^2 = & 0.9577 \quad , \quad DW = 1.9788 \end{aligned}$$

In this paper, the stability over time of the above model is analyzed. For this purpose, an "exploratory" methodology aimed at being sensitive to a wide variety of instability patterns is used; it is based on estimating recursively the model under study and considering associated paths of coefficient estimates and prediction errors. An especially interesting aspect of this approach for our problem is that it can give us

---

<sup>1</sup>Effective observations, not including those observations which are "lost" via the lagging process and the autoregressive transformation. The standard errors are given in parenthesis.  $R^2$  is the multiple correlation coefficient and D.W. the Durbin-Watson statistic (both for the transformed model).



information on the timings of parameter shifts and the directions of resulting prediction errors, one of the implications of Lucas' conjecture for this model. For a detailed description of the methodology followed and its statistical bases, the reader is referred to Dufour (1982). Another purpose of the study is precisely to illustrate the working and usefulness of the recursive approach to the analysis of the stability of econometric models.<sup>2</sup>

Along these lines thus, we present, in Section 2, the results of three different recursive estimation experiments relating to model (1.5) over the period 1956-1972 (and using the data of Gordon and Jorgenson, 1976). First, we estimate (1.5) recursively by ordinary least squares. Second, since Gordon and Jorgenson (1976) corrected the model for autocorrelation (which, in some cases, may be a rather ad hoc correction for a problem of parameter instability), it is important that we look how the conclusions may be affected after making such a correction; thus we do a similar experiment on the correspondingly transformed model (using  $\hat{\rho} = .6223$ , the same estimate of  $\rho$  as Gordon and Jorgenson) :

$$(1.8) \quad IPDE58_t(\hat{\rho}) = \alpha(1-\hat{\rho}) + \delta K_t(\hat{\rho}) + a_1 W_{1t}(\hat{\rho}) + a_2 W_{2t}(\hat{\rho}) + \epsilon_t^*$$

where  $IPDE58_t(\hat{\rho}) = IPDE58_t - \hat{\rho} IPDE58_{t-1}$ ,  $K_t(\hat{\rho}) = K_t - \hat{\rho} K_{t-1}$ , etc.<sup>1</sup>

Finally, we must deal with an extra problem : the capital stock  $K_t$  cannot rigorously be considered as being non-stochastic and independent of the full set of the disturbances  $u_t$  for it is a function of past investment. Thus the

---

<sup>1</sup>See Dufour (1982, Section 2.5) for a discussion of this procedure.

<sup>2</sup>Another illustration, which concerns the stability of the demand for money during the German hyperinflation, may be found in Dufour (1981b).

regressor  $K_t$  may be viewed as a form of lagged dependent variable and the tests performed in the two first experiments cannot be considered exact. As suggested in Dufour (1982, Section 2.5), what can be done in such a case is to get rid of the troublesome regressor  $K_t(\hat{\rho})$  by subtracting  $\hat{\delta}K_t(\hat{\rho})$  on both sides of (1.6) where  $\hat{\delta}$  is the estimate of  $\delta$  based on the full sample; we thus consider the regression :

$$(1.9) \quad \text{IPDE58}_t(\hat{\rho}) - \hat{\delta}K_t(\hat{\rho}) = \alpha(1-\hat{\rho}) + a_1W_{1t}(\hat{\rho}) + a_2W_{2t}(\hat{\rho}) + \epsilon_t^*$$

where  $\hat{\delta} = 0.0572$  and  $\hat{\rho} = .6223$ , and we perform the recursive estimation experiment on the remaining coefficients<sup>1</sup>.

---

<sup>1</sup>Of course, this third experiment involves losing some of the advantages of "recursivity" (we cease to estimate  $\delta$  recursively), hence quite probably a loss of power. But it appears necessary in the present circumstances as a cross-check of the results obtained without taking into account the presence of the lagged dependent variable.

## 2. Recursive Stability Analysis

We first estimate equation (1.5) recursively without any transformation (1956/I-1972/IV)<sup>1</sup>. The recursive estimates obtained are listed in Table 1 and graphed in Figures 1A-1D; the recursive residuals (one to four and eight-steps ahead) are listed in Table 2A, with a number of test statistics in Table 2B<sup>2</sup>, and they are graphed in Figures 2A-2E.

If we look first at the recursive estimates, we can distinguish at least four phases (for all coefficients): the first phase (say, up to 1961/I) is characterized by wide fluctuations (and somewhat "weird" values, especially at the very beginning<sup>3</sup>) and a rough trend (upward for  $\alpha$  and  $a_1$ , downward for  $\delta$  and  $a_2$ ); then we can observe a period of relative stability showing no trend (1961/II-1963/III), except for  $\delta$  which trends upward from 1962/IV; third (1963/IV-1966/IV), we notice a very definite period of trend (downward for  $\alpha$  and  $a_2$ , upward for  $\delta$  and  $a_1$ ), during which all coefficients change sign<sup>4</sup> and fourth, over the rest of the sample period (1967/I-1972/IV),  $a_1$  and  $a_2$  move in the direction opposite to the one followed before while  $\alpha$  and  $\delta$  seem to stabilize. The fourth quarter of 1963 and the last quarter of 1966 clearly appear to be breaking points. Let us now look at the (one-step ahead) recursive residuals. Although there is no systematic tendency to over or underpredict over the full period (as indicated by the global location tests in Table 2B), we can observe a run of 13 consecutive underpredictions from 1963/IV to 1966/IV, a very surprising outcome if the assumed model is correct (the probability of

---

<sup>1</sup>Of course, given that  $K_t$  is a form of lagged dependent variable and if disturbances are autocorrelated, least squares coefficient estimates will be inconsistent. Nevertheless, the appearance of "autocorrelation" may be a symptom of an instability problem and thus an experiment without such a correction seems indicated. In any case, it will allow us to illustrate how a misspecification may lead to the observation of parameter instability in a recursive estimation experiment.

<sup>2</sup>The test statistics in Table 2B, as well as those in Tables 4B and 6 are based on the (one-step ahead) recursive residuals. We report systematically three categories of tests (general tests, run tests and serial dependence tests which can be compared and cross-checked [see Dufour (1982, Section 4)]).

<sup>3</sup>This is not too surprising since, at the beginning, few observations are used for estimation.

<sup>4</sup>This is, by itself, a somewhat preoccupying behaviour.

obtaining at least one run of this length or more is .0065, under the null hypothesis). The total number of runs (of over- or under-predictions) is extremely small (16) in relation to the sample size and there is strong evidence of serial dependence (at least up to a distance of 3 quarters).

Indeed the striking features of the trajectory of the recursive residuals consist in a first period exhibiting a tendency to overpredict (negative residuals) up to 1963/III<sup>1</sup> followed by a long run of 13 consecutive under-predictions (1963/IV-1966/IV), a "breaking point" between 1966/IV and 1967/I, another run of 9 under-predictions (1967/IV-1969/IV), while the sequel looks relatively "random". We can also observe that the two, three and four-steps ahead recursive residuals show basically the same pattern, in fact a substantially more definite (or "cleaner") pattern. It is quite interesting to compare the observed trajectory of the recursive residuals with the movement in the effective investment tax credit<sup>2</sup>. The long run of under-predictions starts (1963/IV) roughly with the repeal of the Long Amendment (1964/I) and extends as long as the effective tax credit is non-zero (till 1966/IV); then we can note a discontinuity in the series (1967/I) which takes place with the suspension of the tax credit, while the following run of under-predictions (1967/IV-1969/IV) can be associated (although more weakly than in the previous case) with the reimposition of the tax credit (1967/II-1969/I)<sup>3</sup>. With respect to the same issue it is also instructive to compute t-statistics to test the null hypothesis of a zero mean (from the one-step ahead recursive residuals) over the separate subperiods corresponding to the different

---

<sup>1</sup> This phenomenon is also indicated by the CUSUM test (see Figure 2F). Note also that the CUSUM of Squares test is not significant (at level .05).

<sup>2</sup> See Table 7 for a listing of variables  $TC_t$  (effective tax credit rate) and  $U_t$  (dummy for Long Amendment) from 1961/I to 1972/IV.

phases of the tax credit (as given in Table 7)<sup>1</sup>. We have the following set of results:

Period <sup>2</sup>	t	p-value <sup>3</sup>
1962/I -1966/III	2.553	.0200
1964/I -1966/III	8.834	.00000251
1967/II-1969/I	1.724	.128
1971/II-1972/IV	1.127	.303
Remainder <sup>4</sup>	-3.790	.000705

It is remarkable that each period where the effective tax credit is non-zero corresponds to a positive t-value (indicating a tendency to under-predict) while the period in which it does not apply produces a negative t-value. The indication is particularly strong (significant) for the first application of the tax credit (especially after the repeal of the Long Amendment). Thus, estimating recursively equation (1.5) by ordinary least squares (and making abstraction of the fact that the capital stock  $K_t$  may be viewed as a form of lagged dependent variable), we find several signs of instability. These, in particular, point to the presence of a substantial shift contemporaneous with the first imposition of the investment tax credit, especially after the repeal of the Long Amendment. Furthermore, the instability involved induced systematic under-prediction of

<sup>1</sup>This is justified by the fact that the (one-step ahead) recursive residuals are i.i.d.  $N[0, \sigma^2]$  under the null hypothesis [see Dufour (1982, Section 4.3)].

<sup>2</sup>1962/I-1966/III corresponds to the first application of the tax credit; 1964/I-1966/III is the same period after the repeal of the Long Amendment; 1967/II-1969/I corresponds to the second application and 1971/II-1972/IV to the third.

<sup>3</sup>Marginal significance level.

<sup>4</sup>1957/I-1961/IV, 1966/IV-1967/I, 1969/II-1971/I.

the level of investment expenditures over this period. In other respects, the two other applications of the tax credit do not seem to have induced a significant effect (although still loosely in the same direction).

Let us now consider the results of a similar experiment applied to equation (1.8), i.e. to model (1.5) after correction for autocorrelation (using  $\hat{\rho} = .6223$ ). The recursive estimates are listed in Table 3 and graphed in Figures 3A-3D; the recursive residuals (one to four and eight-steps ahead) are listed in Table 4A, with a number of test statistics in Table 4B, and they are graphed in Figures 4A-4E.<sup>1</sup> With respect to the recursive estimates, we still distinguish four phases corresponding basically to the same subperiods: first (1957/I-1961/I) wide fluctuations with rough trends (upward for  $\alpha$  and  $a_2$ , downward for  $a_1$ ); second (1961/II-1963/III), a period of relative stability showing no trend (except in the case of  $\delta$  which starts to go up near 1961/IV; thirdly (1963/IV-1966/IV), a very neat trend for all coefficients (downward for  $\alpha$  and  $a_2$ , upward for  $\delta$  and  $a_1$ ) during which all coefficients change sign; fourth (1967/I-1972/IV), a period where all coefficients seem to stabilize. If we then consider the (one-step ahead) recursive residuals (Figure 4A), the pattern appears more "random" than without the transformation (compare with Figure 2A). Global location tests and serial correlation tests are not significant at standard levels (say .10). Nevertheless, a tendency to overpredict is still visible in the earlier period (up to 1963/II) and, especially, a run of 14 consecutive under-predictions from 1963/III to 1966/IV, followed by a sudden drop (1967/I), as observed in the preceding case.<sup>2</sup> The 1967/IV-1969/IV run (of under-predictions) disappears. These observations are very eloquently confirmed when considering the several-

---

<sup>1</sup>The recursive residuals obtained in this way are not, of course, exact (for the true value of  $\rho$  is unknown). Nevertheless, since  $\hat{\rho}$  is a consistent estimate of  $\rho$ , their use is justified in large samples. See Dufour (1982, Section 2.5).

<sup>2</sup>By the way, it is interesting to compare the residuals in Figure 4A (recursive) with the corresponding generalized least squares residuals in

steps ahead recursive residuals (Figures 4B-4E). There are thus continuing signs of instability, in association particularly with the first application of the tax credit (especially after the repeal of the Long Amendment). If we compute, as we did previously, t-statistics over the separate subperiods corresponding to the different phases of the tax credit, we obtain :

Period	t	p-value
1962/I -1966/III	2.178	.0429
1964/I -1966/III	6.066	.0000812
1967/II-1969/I	1.130	.256
1971/II-1972/IV	.194	.853
Remainder <sup>1</sup>	-1.839	.0762

As in the first experiment, we can see that the t-values over each period of application of the tax credit are positive, while, for the rest, we get a negative t-value. Moreover, the t-value for the first period of application is significant (at level .04) and very strongly significant (at level .00008) if the period where the Long Amendment applied is excluded.

Finally, we consider the results of a recursive estimation experiment on equation (1.9), in order to take into account the fact that  $K_t$  is truly a form of lagged dependent variable. The recursive estimates are listed in Table 5 and graphed in Figures 5A-5C; the

---

<sup>1</sup>1957/I-1961/IV, 1966/IV-1967/I, 1969/II-1971/I.

recursive residuals (one to four and eight-steps ahead) are listed in Table 6A, with a number of test statistics in Table 6B, and they are graphed in Figures 6A-6E. With respect to the recursive estimates, we can still observe more or less the same four phases: first (1956/IV-1961/I) wide fluctuations with rough trends (upward for  $\alpha$  and  $a_2$ , downward for  $a_1$ ); second (1961/II-1963/II), a period of relative stability; third (1963/III-1966/IV), a clear trend (downward for  $\alpha$  and  $a_2$ , upward for  $a_1$ ); fourth (1967/I-1972/IV), a period where all coefficients seem to stabilize. With respect to the (one-step ahead) recursive residuals (see Figure 6A), we note now that none of the test statistics in Table 6B nor the CUSUM and CUSUM of Squares test (Figures 6F and 6G) are significant (at level .05). In particular, the longest run test is not conclusive<sup>1</sup>. Nevertheless, the several-steps ahead recursive residuals (Figures 6B-6E) do not seem to be affected in the same way and still exhibit basically the same pattern as in the previous experiment; in particular, the two and three-steps ahead recursive residuals show continuous runs of under-predictions covering the period 1963/III-1966/IV. Indeed, the similitude between Figure 4A and Figure 6A (showing one-step ahead recursive residuals) is itself striking: we note a tendency to over-predict up to 1963/II and a tendency to underpredict over the period 1963/III-1966/IV, while the rest looks relatively "random". If we compute t-values over the separate subperiods corresponding to the separate phases of the tax credit, we find results basically similar to the ones obtained before:

---

<sup>1</sup>Two residuals, in the middle of the longest run previously observed (1963/III-1966/IV) were lowered hence "cutting" the run.



Period	t	p-value
1962/I -1966/III	2.197	.0414
1964/I -1966/III	4.697	.000653
1967/II-1969/I	.957	.370
1971/II-1972/IV	.105	.920
Remainder <sup>1</sup>	-1.944	.0613

The t-value attached to 1962/I-1966/III (first application of the tax credit) is positive and significant at level .04 while, for the period 1964/I-1966/III (after the repeal of the Long Amendment), it is significant at level .00065. The contrast between the periods of applications of the tax credit (which produce positive t-values) and the remainder of the period (which produces a negative t-value) is again to be noted.

Thus, although in a less conclusive form than with the two previous experiments, we can still observe a phenomenon of under-prediction associated with the first imposition of the tax credit (especially after the repeal of the Long Amendment). For the two other applications of the tax credit, we do not observe significant effects (although the corresponding t-values are positive and thus also indicate a tendency to over-predict).

---

<sup>1</sup>1956/IV-1961/IV, 1966/IV-1967/I, 1969/II-1971/I.

### 3. Conclusions

The results obtained in this recursive stability analysis are not as clear and definite as those obtained, for example, for the German demand for money (Dufour, 1982). They are confused, in particular, by the presence of a regressor (the capital stock) which contains lagged values of the dependent variable. Nevertheless, one feature remains constant throughout the three experiments performed : there appears to be a discontinuity associated with the introduction of the first investment tax credit (1962/I-1966/III), especially after the repeal of the Long Amendment (1964/I); furthermore, the discontinuity is of a type that induces under-prediction of investment, a behaviour in contrast with the performance of the model before 1962 (which rather runs in the direction of over-prediction). This phenomenon of underprediction appears quite in line with Lucas' forecast. There is also some indication of a tendency to overpredict investment for the two other periods where the tax credit was in force (1967/II-1969/I and 1971/II-1972/IV), as suggested by the corresponding t-values, but these are too small to be considered significant. Indeed, the latter part of the sample (1967/II-1972/IV) exhibits few signs of instability.

Thus, on the whole, we find some evidence of the type of instability suggested by Lucas (1976), although it appears difficult to qualify this evidence as being very "strong" or "neat". Of course, one could try to explain the instability observed by a misspecification other than the one pointed out by Lucas (e.g. the Almon lag scheme used may be wrong) : in any event, whatever the "true" problem may be, it certainly appears desirable to know about its existence.

Gordon-Jorgenson Model: Recursive Estimates (OLSQ)

1956/I-1972/IV

Quarter <sup>1</sup>	a	b	a <sub>1</sub>	a <sub>2</sub>
5604.00	-.326	1.609	-.183	.188
5701.00	-.355	1.793	-.199	.206
5702.00	-.250	.932	-.113	.142
5703.00	-.250	1.169	-.145	.142
5704.00	.103	-.271	-.447	.447
5801.00	-.201	-.300	-.380	.478
5802.00	-.213	-.522	-.787	.603
5803.00	-.213	-.522	-.769	.603
5804.00	-.113	-.189	-.283	.649
5901.00	-.112	-.144	-.219	.649
5902.00	-.127	-.165	-.246	.277
5903.00	-.366	-.338	-.105	.915
5904.00	-.57	-.476	-.553	.120
6001.00	-.390	-.949	-.464	.111
6002.00	-.399	-.964	-.462	.110
6003.00	-.27	-.126	-.452	.986
6004.00	-.12	-.160	-.541	.723
6101.00	-.77	-.207	-.792	.199
6102.00	-.20	-.232	-.115	.372
6103.00	-.27	-.239	-.114	.465
6104.00	-.299	-.240	-.113	.489
6201.00	-.288	-.241	-.113	.490
6202.00	-.288	-.241	-.114	.482
6203.00	-.288	-.241	-.119	.533
6204.00	-.300	-.212	-.755	.130
6301.00	-.31	-.194	-.113	.171
6302.00	-.31	-.191	-.125	.686
6303.00	-.31	-.194	-.104	.833
6401.00	-.300	-.196	-.734	.580
6402.00	-.288	-.196	-.410	.205
6403.00	-.27	-.194	-.186	.451
6404.00	-.26	-.191	-.748	.586
6501.00	-.23	-.180	-.202	.907
6502.00	-.21	-.169	-.458	.119
6503.00	-.17	-.146	-.891	.168
6504.00	-.12	-.118	-.130	.213
6601.00	-.7	-.836	-.168	.255
6602.00	-.2	-.460	-.195	.282
6603.00	-.2	-.904	-.224	.311
6604.00	-.6	-.171	-.250	.338
6701.00	-.3	-.738	-.223	.309
6702.00	-.2	-.127	-.205	.290
6703.00	-.1	-.183	-.197	.282
6704.00	-.2	-.929	-.191	.274
6801.00	-.5	-.243	-.152	.226
6802.00	-.6	-.287	-.146	.219
6803.00	-.7	-.374	-.134	.203
6804.00	-.7	-.414	-.129	.197
6901.00	-.8	-.530	-.118	.183
6902.00	-.8	-.553	-.116	.181
6903.00	-.9	-.577	-.114	.178
6904.00	-.9	-.592	-.117	.181
7001.00	-.8	-.553	-.104	.167
7002.00	-.9	-.536	-.994	.162
7003.00	-.9	-.536	-.997	.162
7004.00	-.7	-.420	-.121	.188
7101.00	-.8	-.410	-.123	.191
7102.00	-.8	-.468	-.107	.191
7103.00	-.8	-.494	-.975	.161
7104.00	-.9	-.571	-.622	.120
7201.00	-.9	-.611	-.405	.956
7202.00	-.9	-.609	-.420	.973
7203.00	-.9	-.593	-.541	.110
7204.00	-.9	-.592	-.545	.111

<sup>1</sup>End of sample (1956/I- )

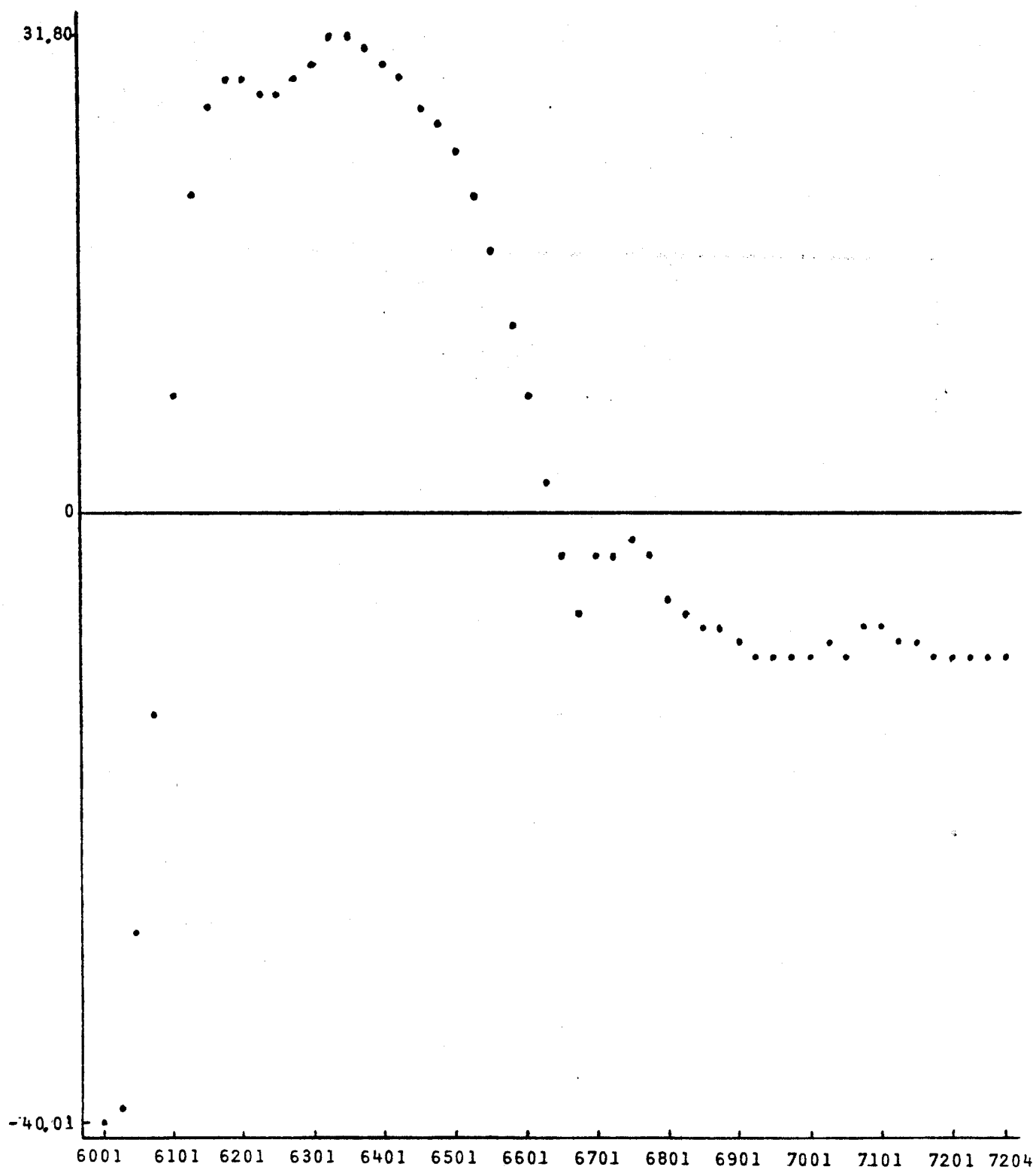


FIGURE 1A - Gordon-Jorgenson Model: Recursive Estimates of  $\alpha$  (OLSQ)<sup>1</sup>

<sup>1</sup>The values up to 1959/IV are excluded from the graph because most of them are "too big". Similar remark for Figures 1B-D, 3A-D, 5A-C.

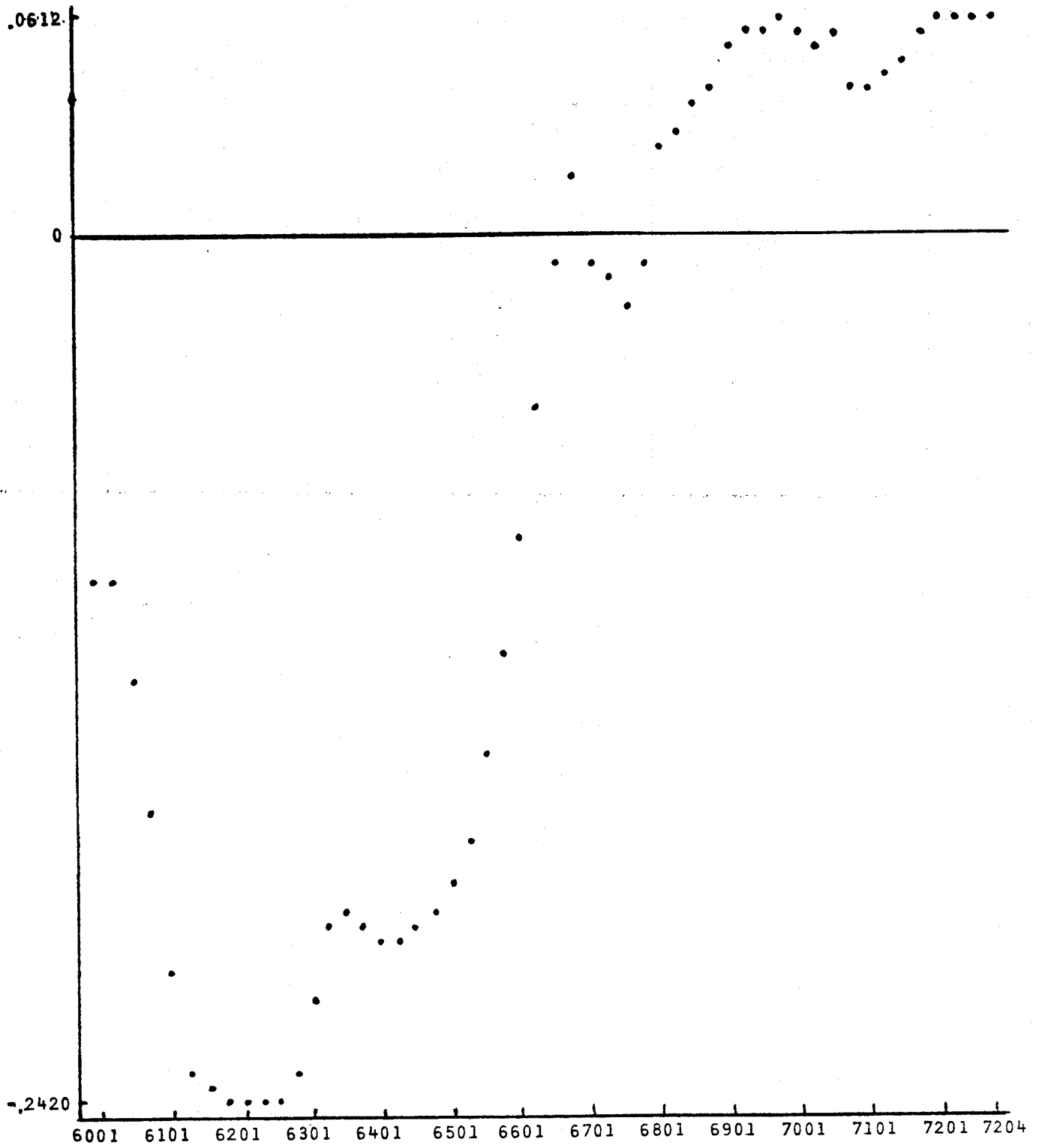


FIGURE 1B - Gordon-Jorgenson Model: Recursive Estimates of  $\delta$  (OLSQ)

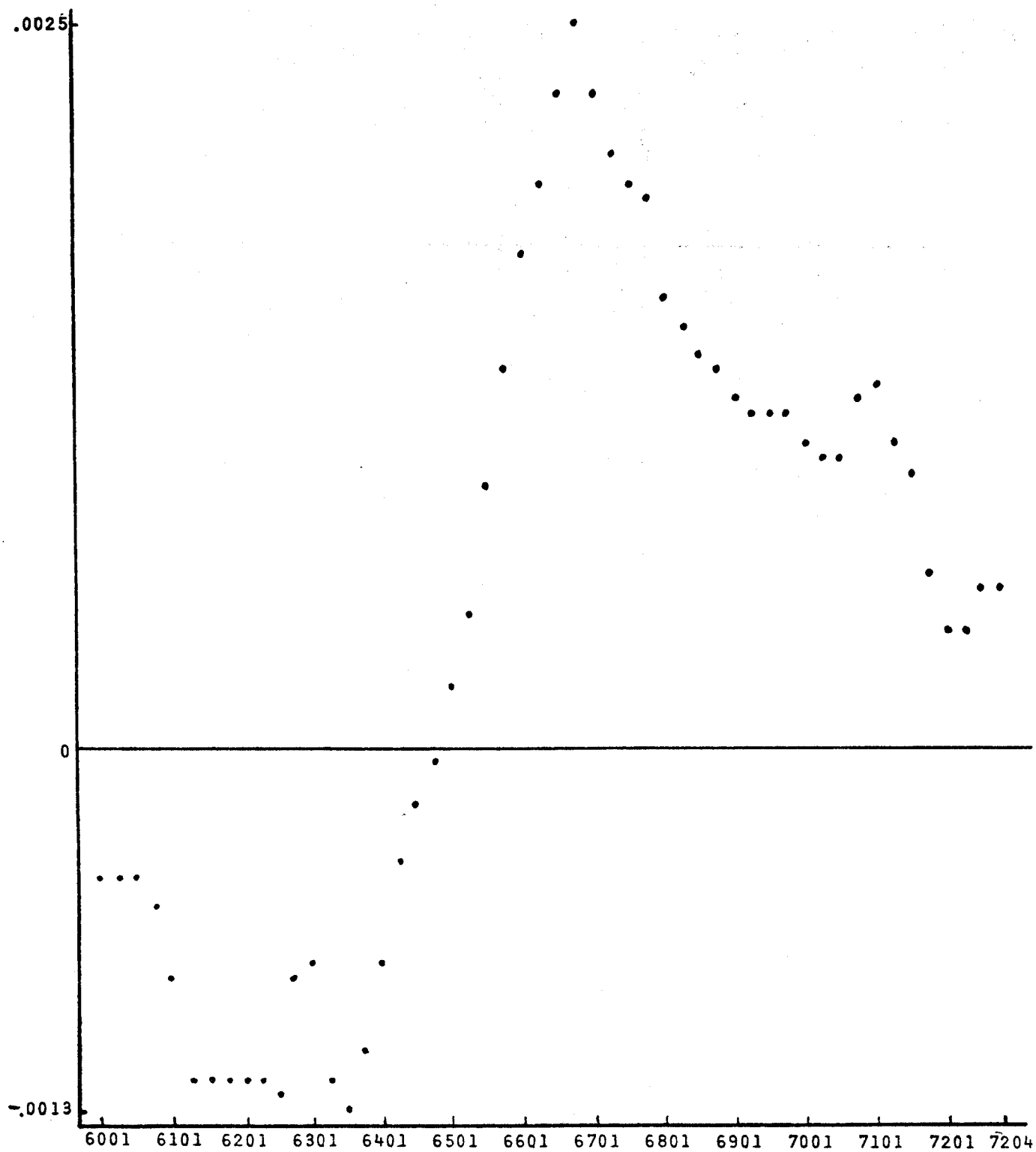


FIGURE 1C - Gordon-Jorgenson Model: Recursive Estimates of  $a_1$  (OLSQ)

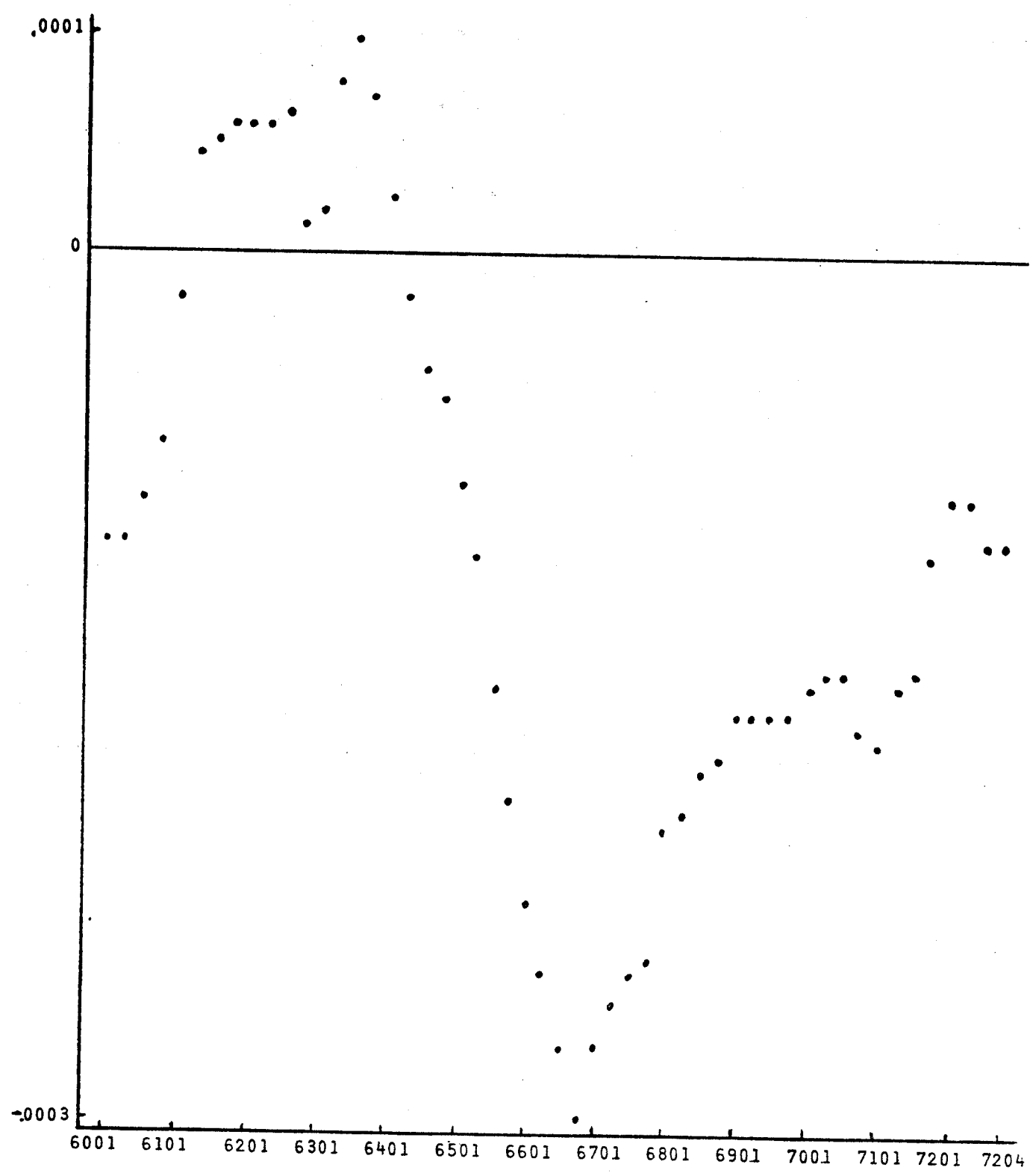


FIGURE 1D - Gordon-Jorgenson Model: Recursive Estimates of  $a_2$  (OLSQ)

TABLE 2A

Gordon-Jorgenson Model: Recursive Residuals (OLSQ)

1957/I-1972/IV

Quartèr <sup>1</sup>	RECF1 <sup>2</sup>	RECF2	RECF3	RECF4	RECF8
5701	00	0	0	0	0
5702	357818E-01	0.824488	0.585280E-01	0.	0.
5703	344723	0.251558	0.483592	0.	0.
5704	19354287	0.351524	0.764657	0.7771	0.
5801	2969770	0.167530	0.108739	0.105566	0.
5802	7966403	0.255509	0.677388	0.3999E-01	0.
5803	1621278	0.304827	0.817239	0.5022E-01	0.
5804	101153	0.398108	0.566173	0.794517	0.
5901	101771	0.654247	0.40776	0.818519	0.
6001	4146578	0.354248	0.721637	0.194521	0.
6002	105791	0.49498	0.023864	0.49889	0.
6003	04991	0.951927	0.210225	0.62681	0.
6101	04728	0.351484	0.51445	0.76681	0.
6102	508768	0.315846	0.19140	0.51373	0.
6103	58891E-01	0.00708	0.12519	0.3951E-02	0.
6201	398205	0.19461	0.66926	0.35415	0.
6202	127645	0.349657	0.10338	0.43185	0.
6301	298449	0.25531	0.55715	0.66507	0.
6302	155238	0.155312	0.95715	0.74981	0.
6303	307218	0.28401	0.42846	0.09381	0.
6401	098774	0.716995	0.52567	0.39139	0.
6402	071538	0.182995	0.95407	0.87434	0.
6403	105328	0.448835	0.69798	0.53078	0.
6501	379614	0.66012	0.48861	0.22079	0.
6502	380449	0.89013	0.15977	0.22079	0.
6503	360682	0.88758	0.62382	0.33043	0.
6601	317479	0.97036	0.47382	0.46369	0.
6602	287833	0.80225	0.84290E-01	0.60384	0.
6701	154483	0.20225	0.40257	0.51894	0.
6702	504857	0.139618	0.79457	0.45819	0.
6703	504139	0.256618	0.40335	0.61913	0.
6801	461751	0.66826	0.44189	0.56173	0.
6802	1251751	0.191364	0.01898	0.1842	0.
6803	651751	0.66826	0.44189	0.56173	0.
6804	1251751	0.191364	0.01898	0.1842	0.
6805	651751	0.66826	0.44189	0.56173	0.
6806	1251751	0.191364	0.01898	0.1842	0.
6807	651751	0.66826	0.44189	0.56173	0.
6808	1251751	0.191364	0.01898	0.1842	0.
6809	651751	0.66826	0.44189	0.56173	0.
6810	1251751	0.191364	0.01898	0.1842	0.



TABLE 2A(continued)

Quarter	RECF1	RECF2	RECF3	RECF4	RECF8
6803.00	.685086	.728844	1.37120	1.43665	.947518
6804.00	.420879	.618989	1.667951	1.32924	.271486
6901.00	1.60633	1.66055	1.792551	1.742221	1.69343
6902.00	.374977	.712641	.798389	.979769	1.41324
6903.00	.40651	.478478	.836154	.859727	1.75007
6904.00	.373479	.443668	.511301	.920745	1.574585E-01
7001.00	-1.26358	-1.13770	-1.751653	.932465	.292775
7003.00	1.568404	.765092	.451672	.699850	1.63588
7004.00	-2.77299	-2.85289	-2.62206	.479465	1.581686
7102.00	-2.171176	.923891	.160334	.278711	.539415E-01
7103.00	1.497430	.791176	.702334	.354265E-01	.202776
7104.00	1.84262	1.89770	2.12844	.696984	1.16619
7201.00	1.38766	1.96370	2.0827	1.23195	1.39959
7202.00	-1.19159	1.32281	.98081	2.10303	1.678776
7203.00	-1.134017	-1.32819	.849760	1.134454	.333327
7204.00	1.745724E-01	-1.291257	.309686	.558641E-01	.935315

<sup>1</sup>Quarter of the observation predicted.

<sup>2</sup>RECFk refers to the k-steps ahead recursive residuals (k = 1, 2, 3, 4, 8).

TABLE 2B

Gordon-Jorgenson Model (OLSQ): Test Statistics<sup>1</sup>

Number of residuals = 64

Global Location Tests <sup>2</sup>			p-values <sup>3</sup>				
	t-Test	.0619		.9506			
No of Positive Residuals	32		1.0000				
Wilcoxon Test	1053		.9307				
Runs Tests <sup>4</sup>	No of Runs	16	.000019				
	Length of the Longest Run	13	.0065				
Serial Correlation Tests <sup>5</sup>	Modified Von Neumann Ratio		.6779	≤ .002			
	Rank Tests						
	k	Signed-Rank Tests			Sign-Tests		
		$S_k$	$S'_k$	p-value	$S_k$	$S'_k$	p-value
	1	1735	4.977	.00000065	48	4.158	.000038
2	1421	3.116	.0018	41	2.540	.0151	
3	1284	2.431	.0150	37	1.664	.1237	
4	1091	1.296	.1951	33	.7746	.5190	
5	1041	1.177	.2390	33	.9113	.4350	
6	1095	1.854	.0637	35	1.576	.1480	
7	1058	1.839	.0659	35	1.722	.1112	
8	983	1.509	.1313	32	1.069	.3497	
9	1015	2.053	.0401	33	1.483	.1770	
10	958	1.856	.0635	32	1.361	.2203	
11	877	1.430	.1528	31	1.236	.2717	
12	807	1.075	.2825	30	1.109	.3317	

<sup>1</sup>Based on the one-step ahead recursive residuals.

<sup>2</sup>See Dufour (1982, Section 4.3). These are two-sided tests.

<sup>3</sup>Marginal significance levels.

<sup>4</sup>See Dufour (1982, Section 4.5). These are one-sided tests :  $P[R \leq 16] = .000019$  and  $P[L \geq 16] = .0065$ , where  $R$  = no of runs (of + 's or -'s) and  $L$  = length of the longest run.

<sup>5</sup> $S_k$  is a rank statistic for testing serial dependence [see Dufour (1982, Section 4.6)], where  $k$  is the lag used,  $a_k(r) = 1$  for the sign test and  $a_k(r) = r$  for the signed-rank test;  $S'_k = (S_k - N\mu_k)/\sigma_k$ , where  $\mu_k = E(S_k)$  and  $V(S_k) = \sigma_k^2$  under the null hypothesis. We consider here two-sided tests (against positive or negative serial dependence). For a more complete theory of these tests, see Dufour (1981).

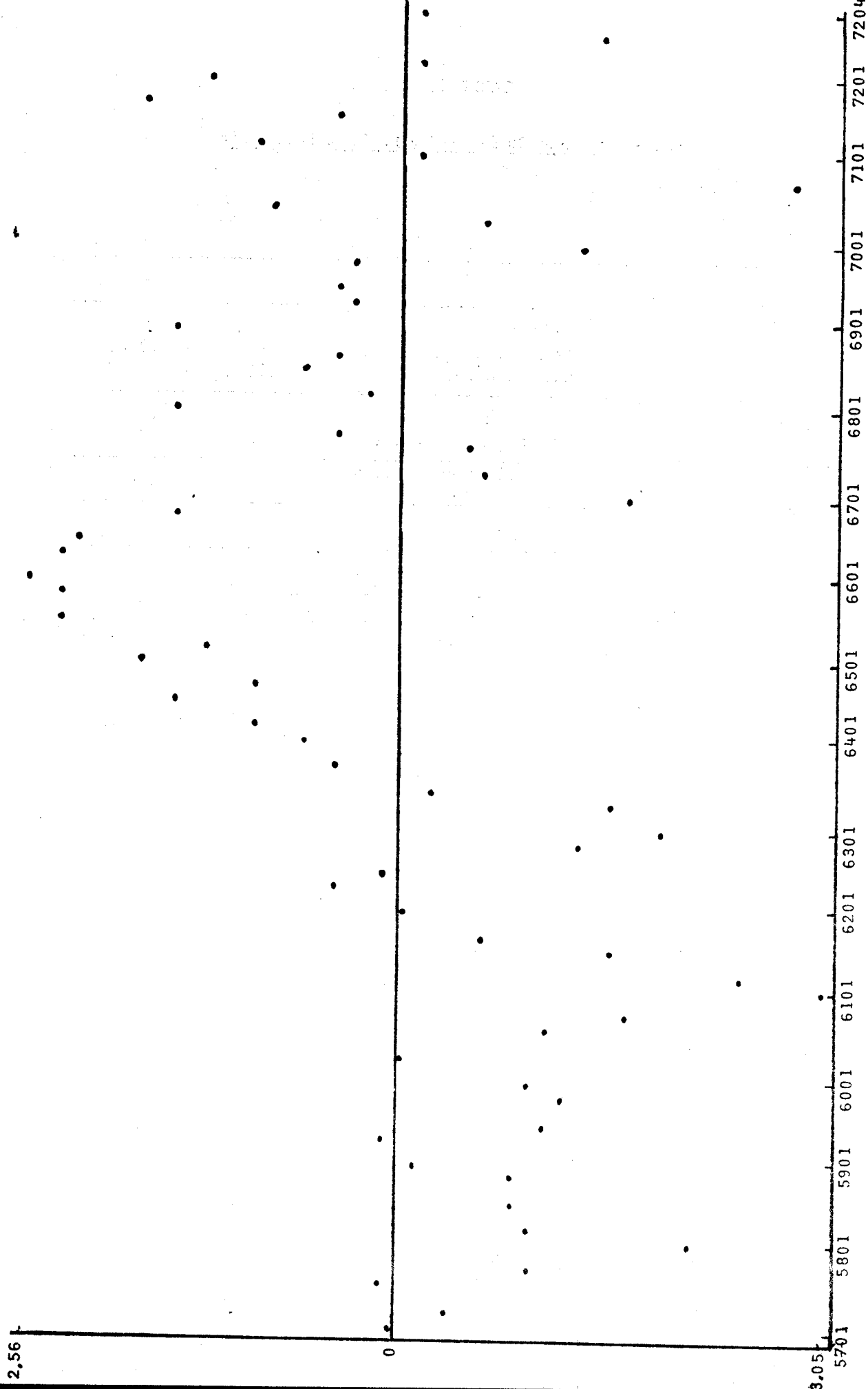


FIGURE 2A - Gordon-Jorgenson Model: One-Step Ahead Recursive Residuals (OLSQ)

3.10

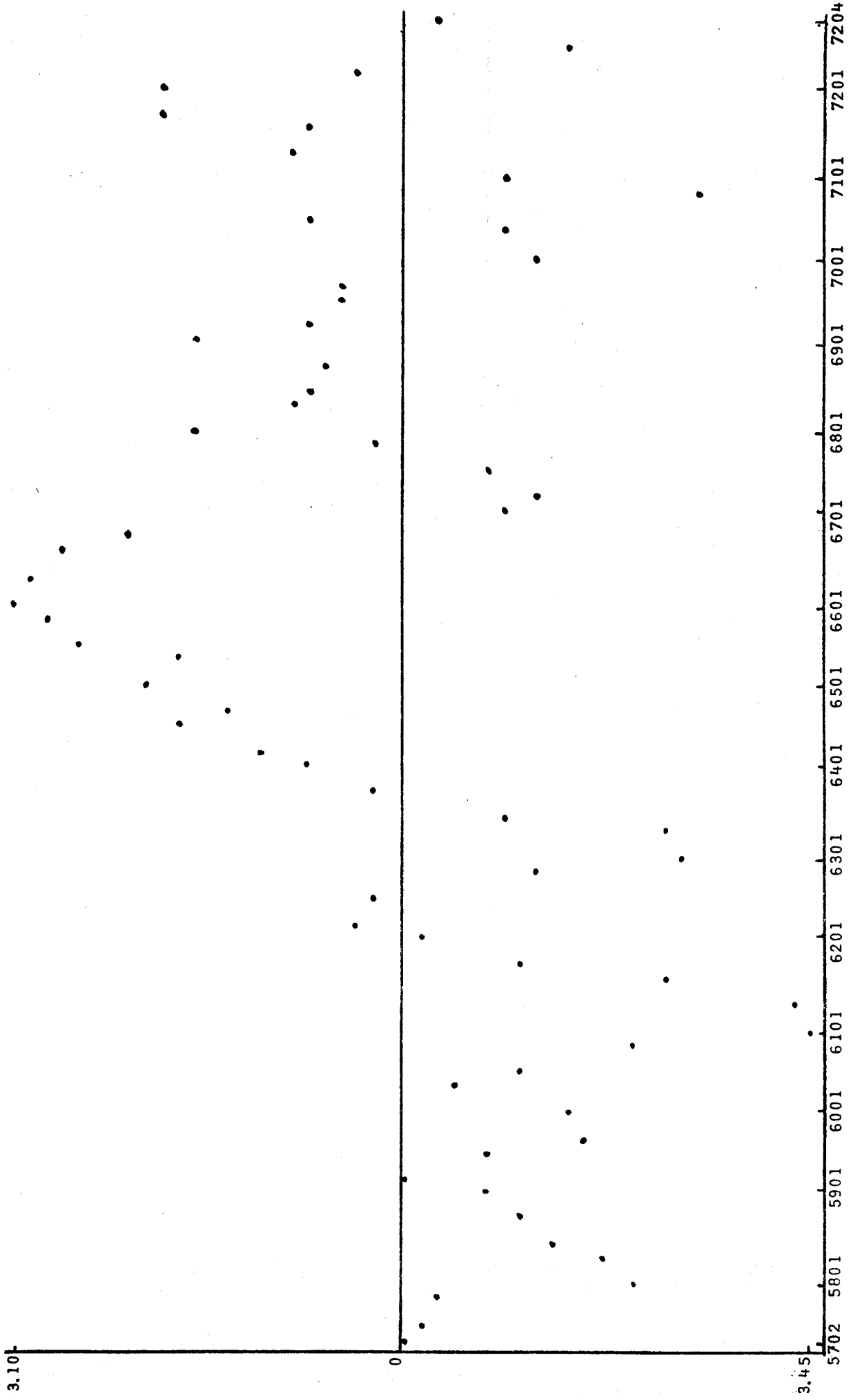


FIGURE 2B - Gordon-Jorgenson Model: Two-Steps Ahead Recursive Residuals (OLSQ)

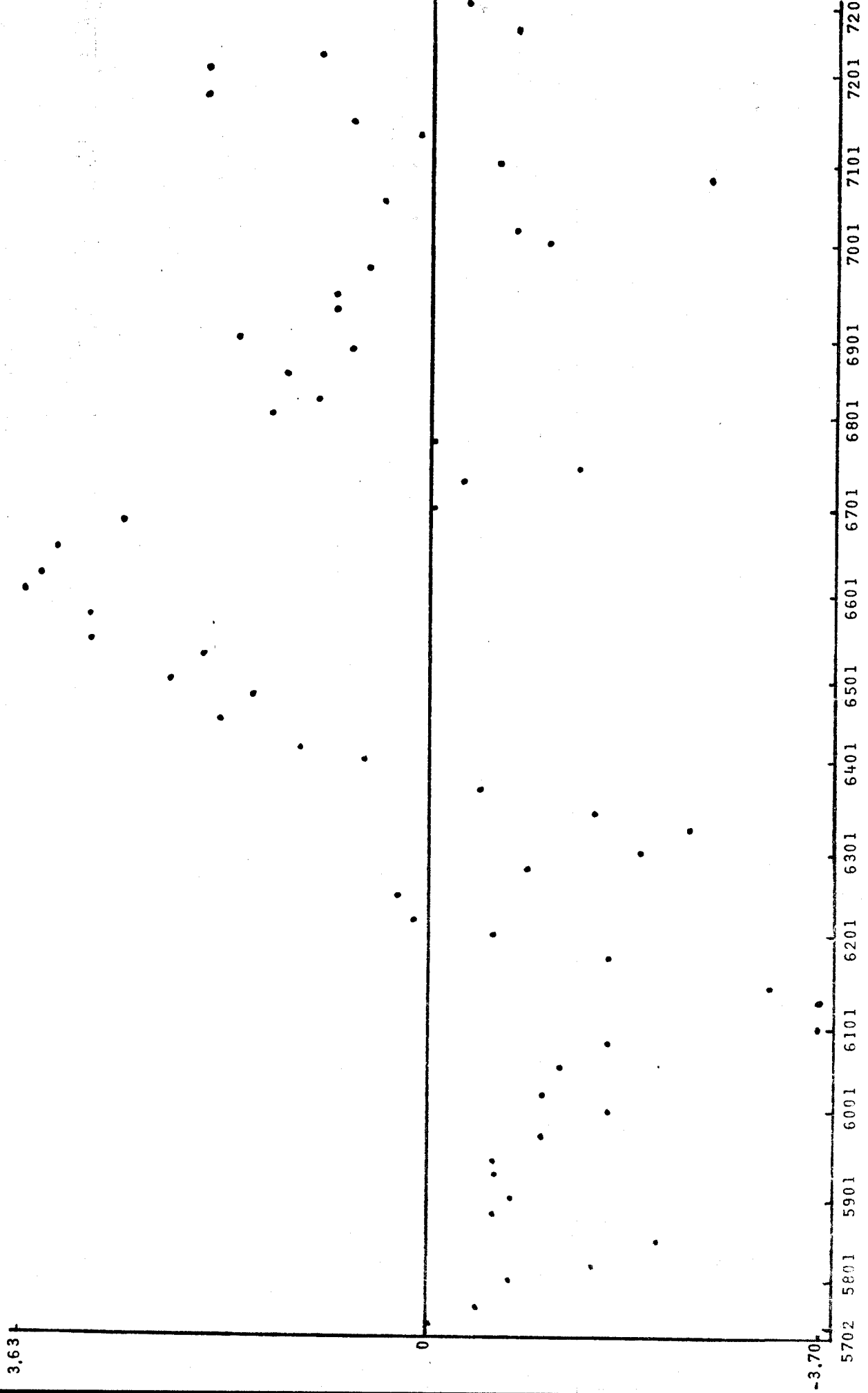


FIGURE 2C - Gordon-Jorgenson Model: Three-Steps Ahead Recursive Residuals (OLSQ)

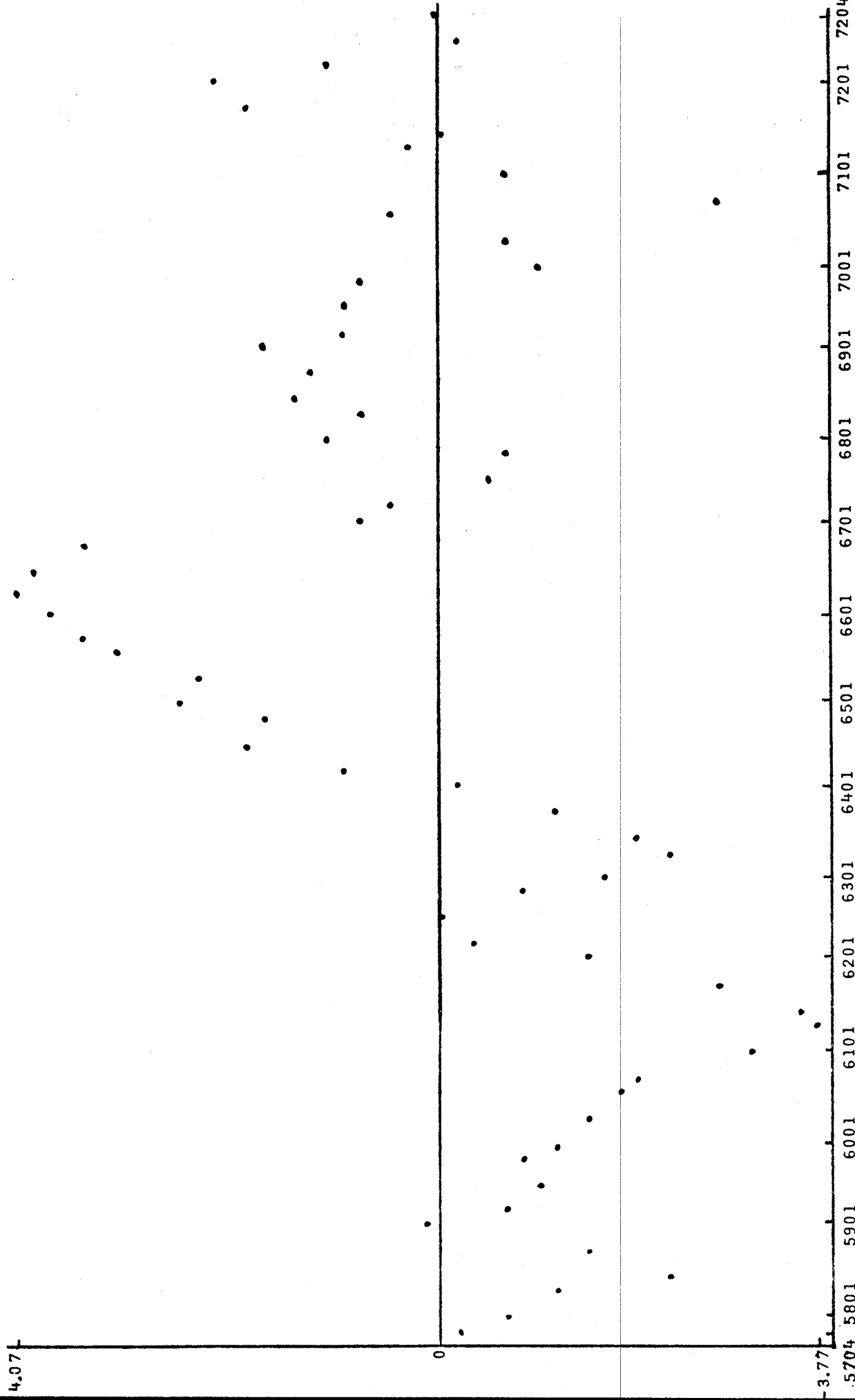


FIGURE 2D - Gordon-Jorgenson Model: Four-Steps Ahead Recursive Residuals (OLSQ)

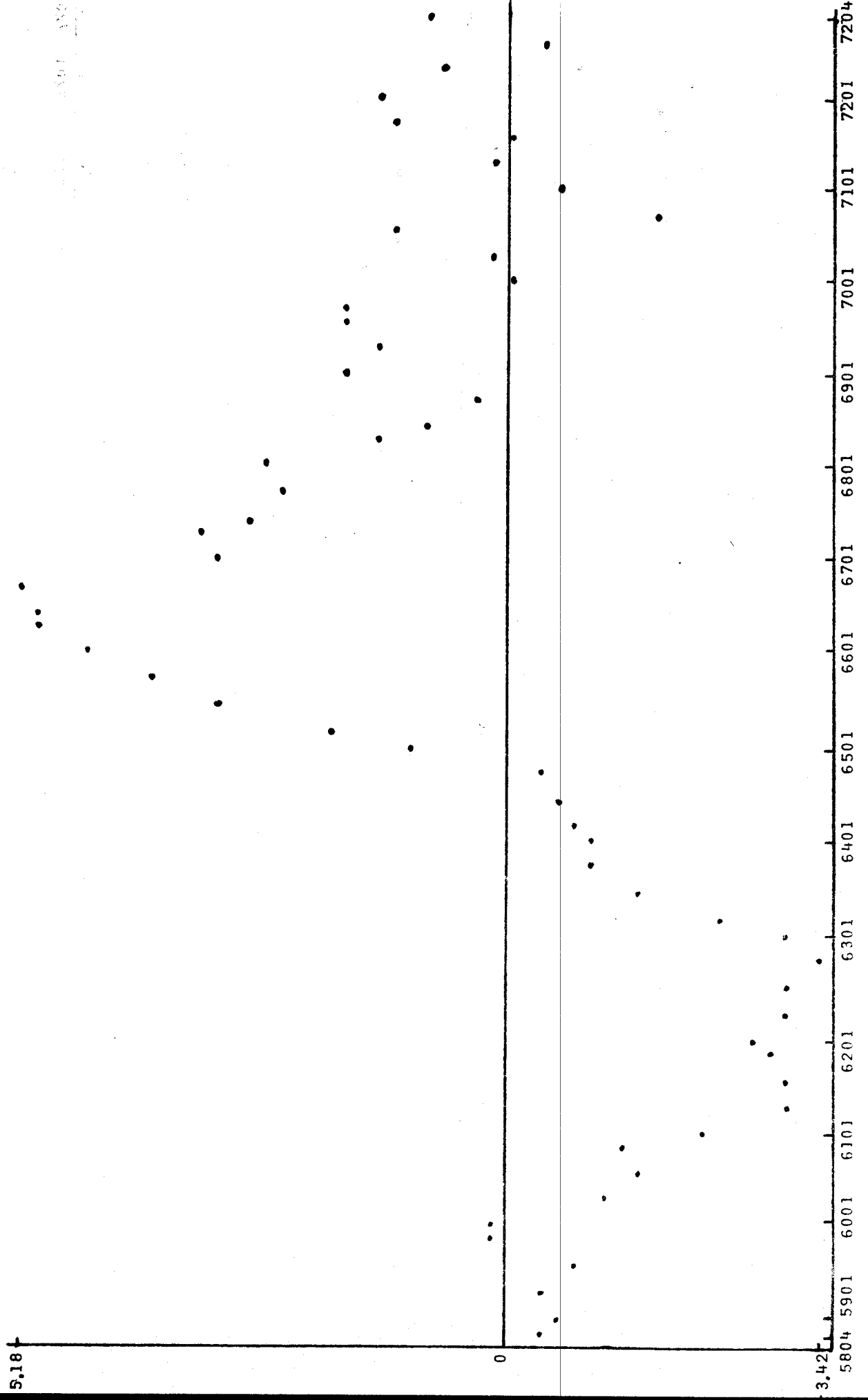


FIGURE 2E - Gordon-Jorgenson Model: Eight-Steps Ahead Recursive Residuals (OLSQ)

5 % SIGNIFICANCE LINE

5 % SIGNIFICANCE LINE

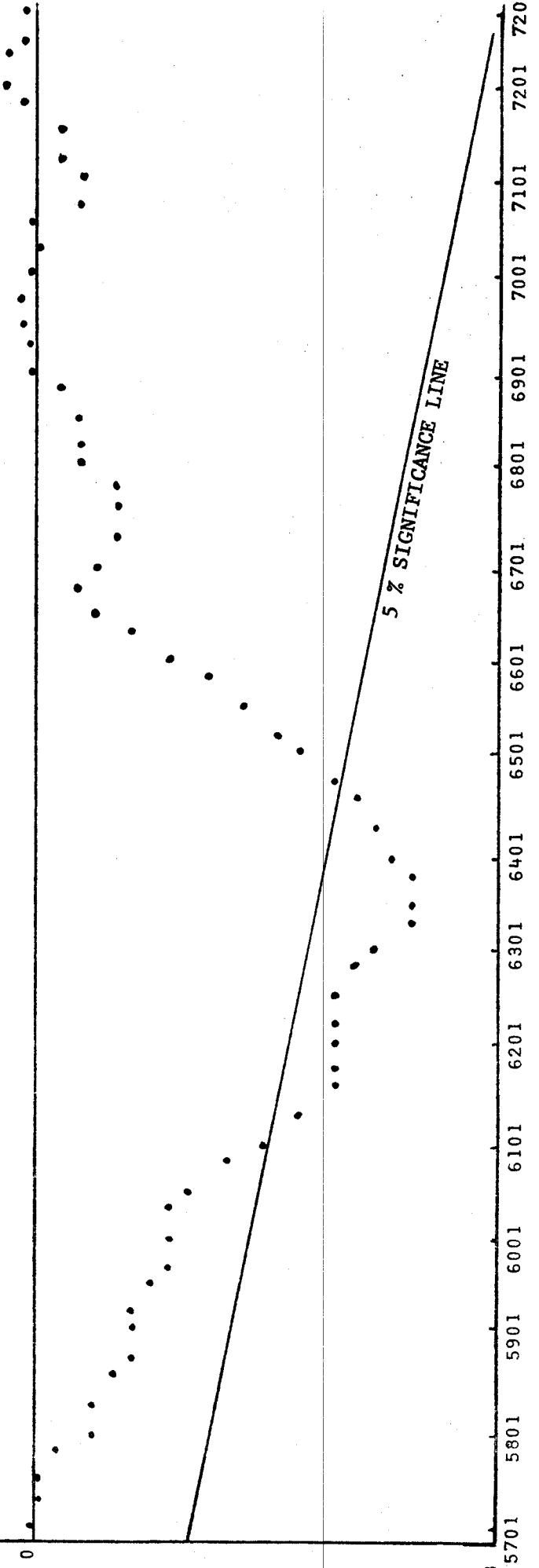


FIGURE 2F - Gordon-Jorgenson Model: CUSUM of Recursive Residuals (OLSQ)<sup>1</sup>

<sup>1</sup>Based on the one-step ahead recursive residuals.



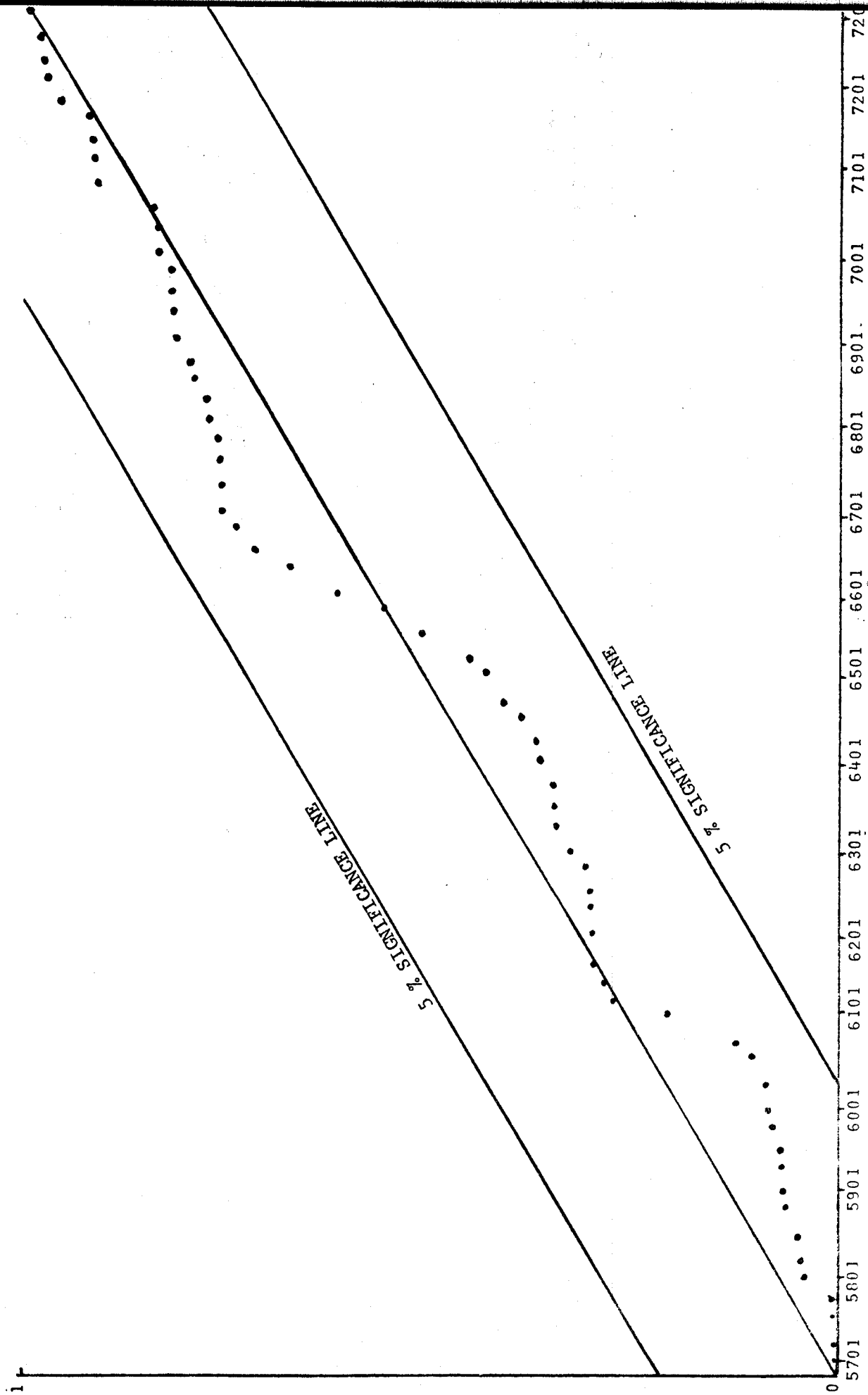


FIGURE 2C - Gordon-Jorgenson Model: CUSUM of Squares of Recursive Residuals (OLSQ)<sup>1</sup>

<sup>1</sup> Based on the one-step ahead recursive residuals.

Gordon-Jorgenson Model: Recursive Estimates ( $\hat{\rho} = .6223$ )

1956/I-1972/IV

Quarter	$\alpha$	$\delta$	$a_1$	$a_2$
5604.00	-1.363	-1.911	.518	-.719
5701.00	-.771	.242	.228	.799
5702.00	-.333	.133	.183	.348
5703.00	-.106	.066	.091	.144
5704.00	-.055	.033	.049	.072
5801.00	-.055	.033	.049	.072
5802.00	-.137	.091	.122	.183
5803.00	-.149	.099	.133	.199
5804.00	-.151	.101	.135	.201
5901.00	-.244	.801	.622	.217
5902.00	-.277	.873	.673	.249
5903.00	-.293	.860	.658	.244
5904.00	-.311	.823	.620	.251
6001.00	-.300	.842	.647	.221
6002.00	-.308	.846	.633	.221
6003.00	-.300	.705	.577	.337
6004.00	-.346	.755	.622	.344
6101.00	-.411	.781	.654	.233
6102.00	-.400	.702	.599	.233
6103.00	-.442	.730	.622	.268
6104.00	-.480	.750	.640	.268
6201.00	-.338	.299	.300	.263
6202.00	-.366	.296	.300	.263
6203.00	-.355	.275	.283	.258
6204.00	-.340	.241	.233	.182
6301.00	-.400	.222	.224	.140
6302.00	-.411	.222	.224	.140
6303.00	-.377	.220	.217	.147
6304.00	-.333	.213	.209	.127
6401.00	-.300	.201	.188	.111
6402.00	-.230	.181	.166	.093
6403.00	-.233	.166	.153	.083
6404.00	-.206	.166	.153	.075
6501.00	-.155	.117	.107	.055
6502.00	-.113	.105	.097	.045
6503.00	-.144	.413	.388	.289
6504.00	-.138	.639	.617	.897
6601.00	-.138	.533	.506	.897
6602.00	-.112	.820	.866	.103
6603.00	-.116	.114	.106	.103
6604.00	-.117	.125	.135	.103
6701.00	-.157	.265	.277	.125
6702.00	-.138	.464	.472	.640
6703.00	-.177	.364	.355	.145
6704.00	-.193	.408	.462	.147
6801.00	-.111	.731	.719	.139
6802.00	-.138	.466	.472	.914
6803.00	-.109	.543	.528	.128
6804.00	-.100	.538	.508	.128
6901.00	-.110	.698	.602	.117
6902.00	-.110	.625	.612	.117
6903.00	-.110	.644	.603	.119
6904.00	-.110	.672	.693	.117
7001.00	-.109	.555	.497	.127
7002.00	-.109	.585	.512	.107
7003.00	-.111	.743	.339	.108
7004.00	-.111	.283	.110	.867
7101.00	-.118	.445	.478	.180
7102.00	-.109	.540	.544	.141
7103.00	-.109	.514	.642	.132
7104.00	-.110	.596	.642	.128
7201.00	-.110	.606	.241	.785
7202.00	-.109	.584	.345	.719
7203.00	-.109	.570	.471	.903
7204.00	-.109	.572	.468	.103

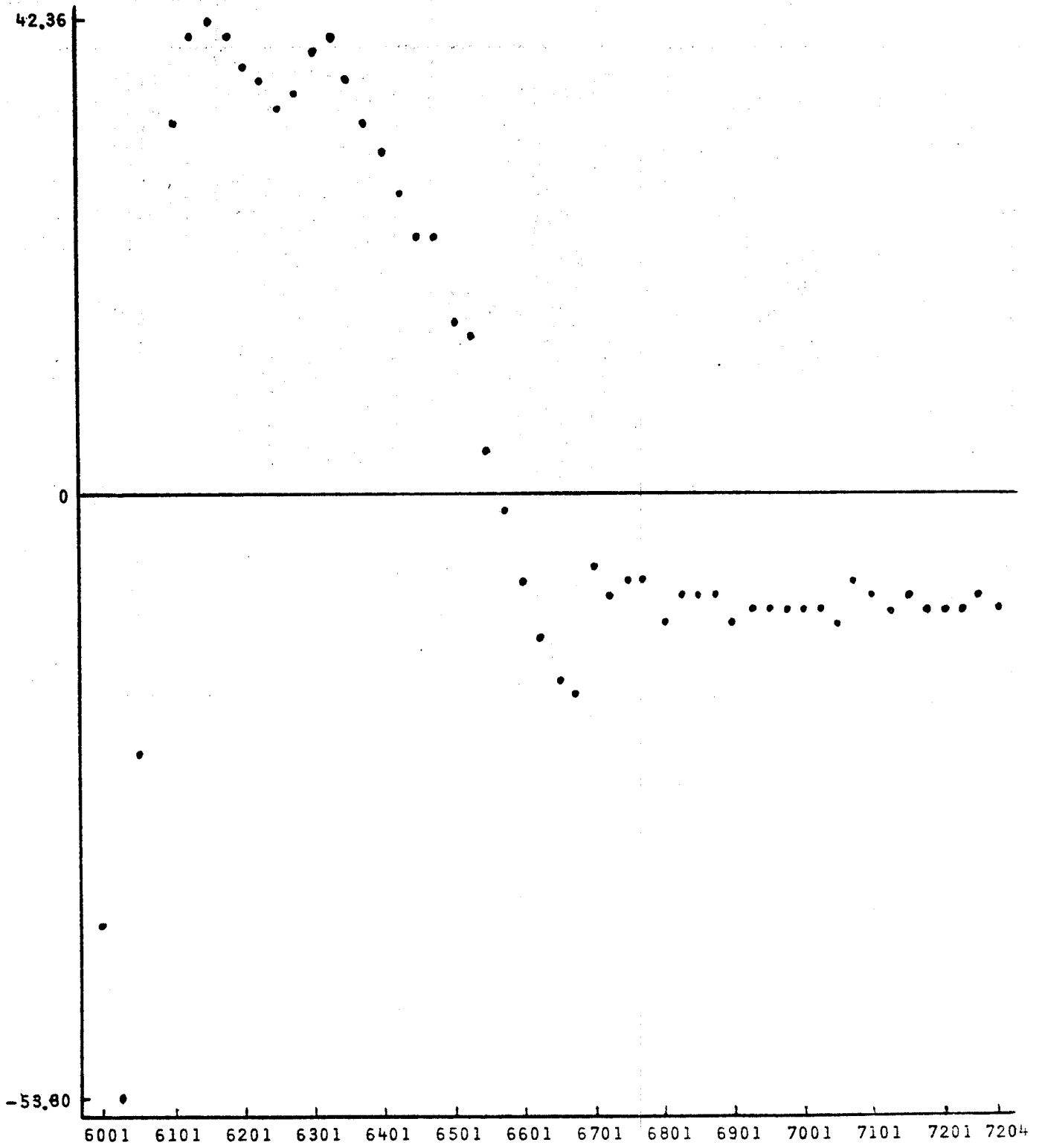


FIGURE 3A - Gordon-Jorgenson Model: Recursive Estimates of  $\alpha(\hat{\rho} = .6223)$

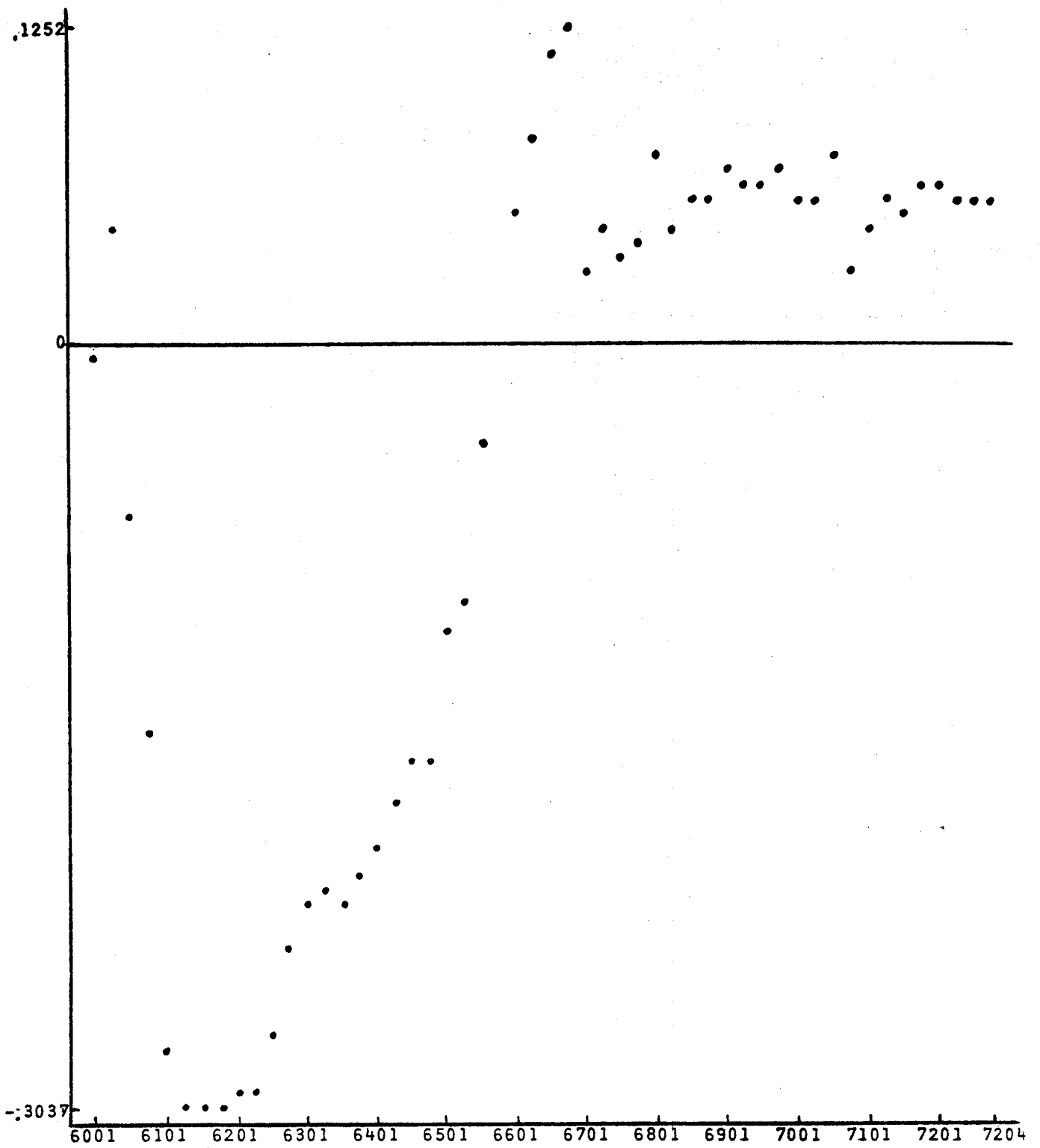


FIGURE 3B - Gordon-Jorgenson Model: Recursive Estimates of  $\delta(\hat{\rho} = .6223)$

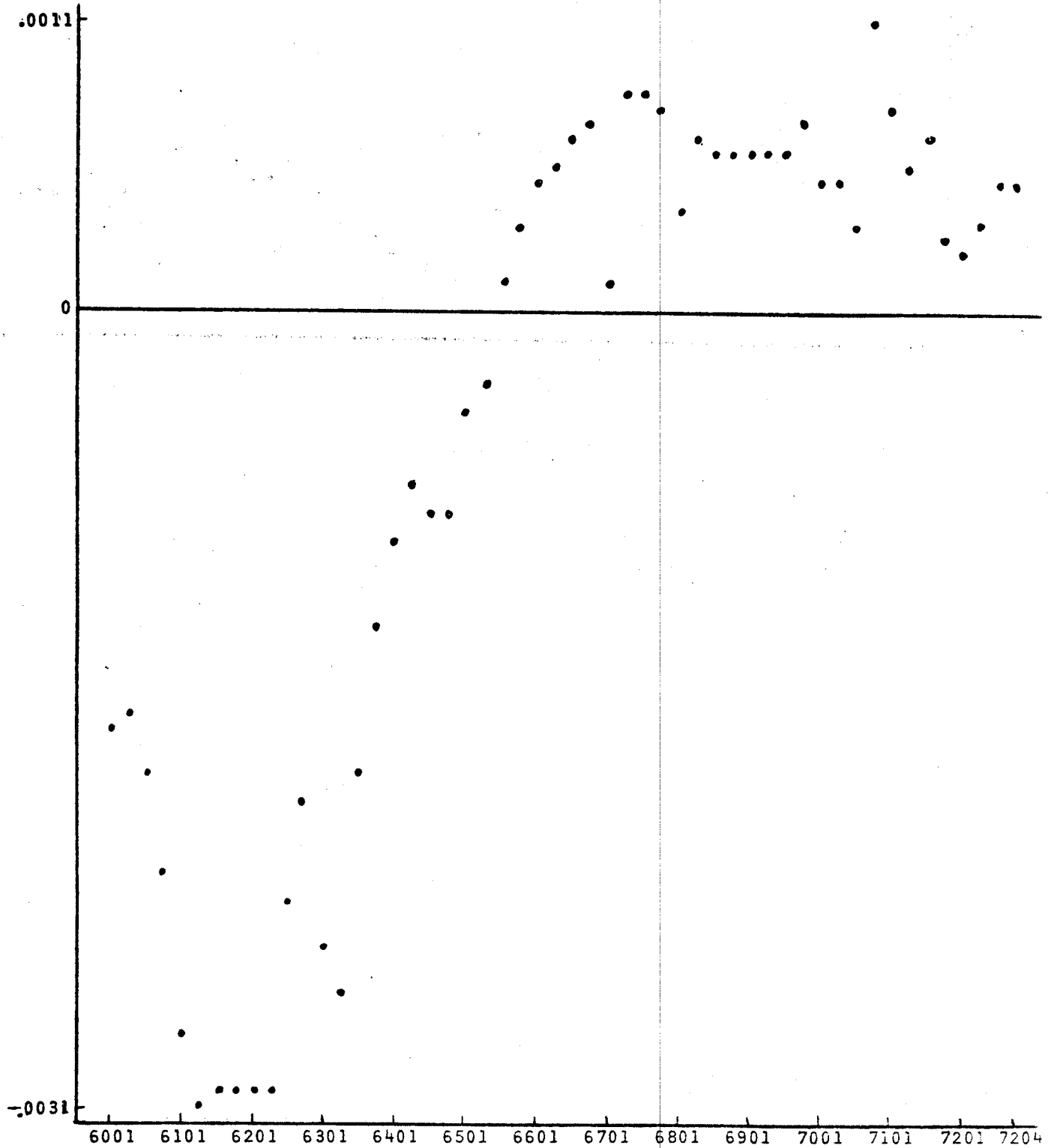


FIGURE 3C - Gordon-Jorgenson Model : Recursive Estimates of  $a_1$  ( $\hat{\rho} = .6223$ )

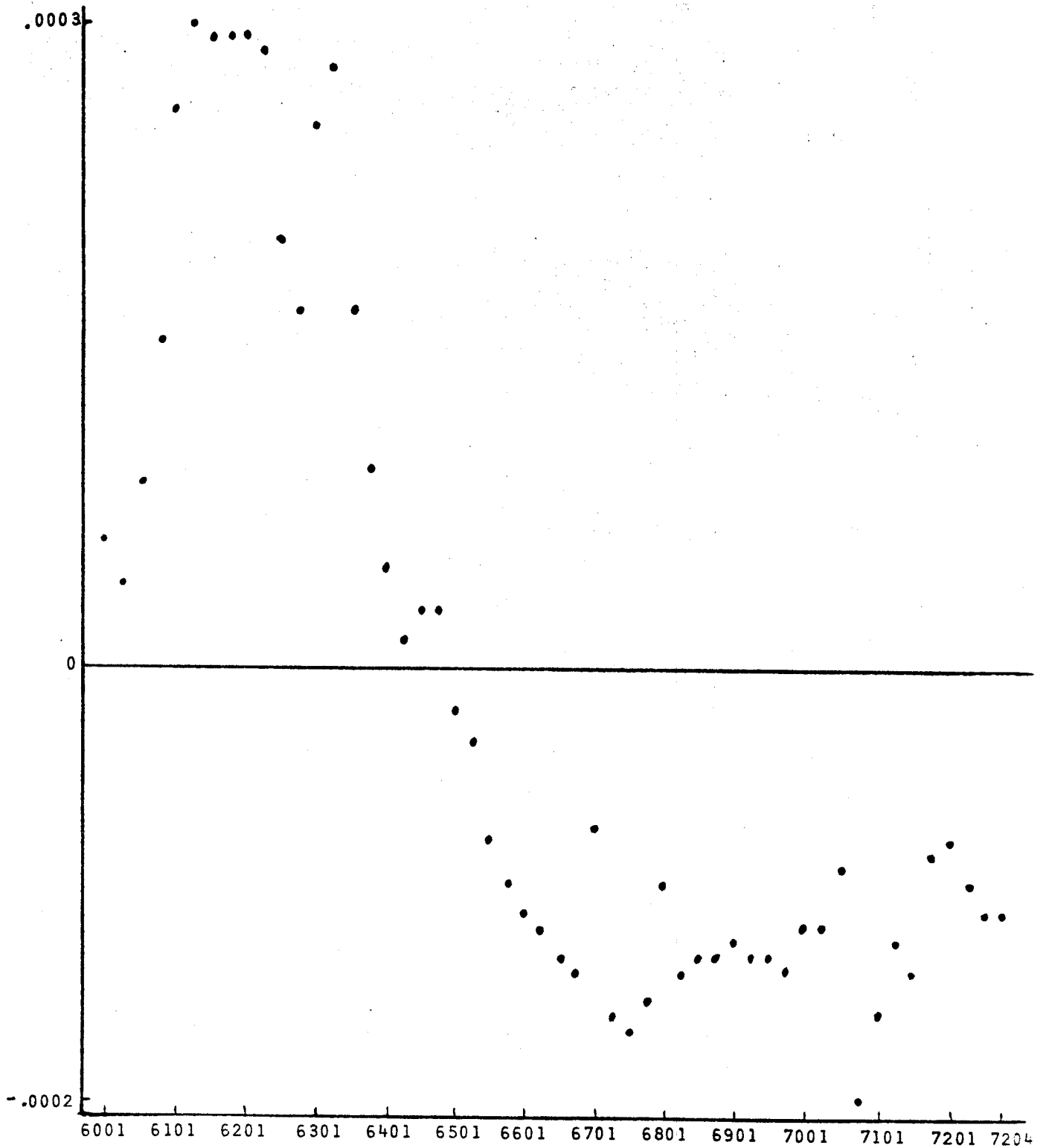


FIGURE 3D - Gordon-Jorgenson Model: Recursive Estimates of  $a_2$  ( $\hat{\rho} = .6223$ )

TABLE 4A

Gordon-Jorgenson Model : Recursive Residuals ( $\hat{\rho} = .6223$ )  
1957/I-1972/IV

Quarter	RECF1	RECF2	RECF3	RECF4	RECF8
5701	0	0	0	0	0
5702	8611672	774472E-01	859351	0	0
5703	5281268	800702	474460	0	0
5704	193988	2185038	481472	0	0
5801	583009	150935	419871	0	0
5802	546452	781836	797495	0	0
5803	279880	569645	788367	0	0
5804	3019404	1089220	406180	0	0
5901	486282	118943	907941	0	0
5902	231746	117947	250382	0	0
5903	556526	488567	1493385	0	0
6001	117772	282971	19348	0	0
6002	121116	43899	06434	0	0
6003	306238	36806	615306	0	0
6101	636879	139637	48306	0	0
6102	270189	139904	215653	0	0
6103	348389	403974	457858	0	0
6201	529047	408137	31109	0	0
6202	47811	55044	625796	0	0
6301	026633	671442	133797	0	0
6302	855462	966897	157065	0	0
6303	649300	780248	001528	0	0
6304	671528	813982	207152	0	0
6401	088899E-01	116602	265595	0	0
6402	1723879	3022320	465525	0	0
6403	1254325	555404	1584570	0	0
6404	327221	225097	44664	0	0
6501	992098	182505	110160	0	0
6502	608547	810909	109809	0	0
6503	647405	855234	031733	0	0
6504	241159	123413E-01	69170E-01	0	0
6601	080004	2442257E-02	357529	0	0
6602	1284050	804847	867431	0	0
6603	874075	804847	667871	0	0
6604	998884	574932	434692	0	0
6701	0	0	0	0	0
6702	0	0	0	0	0
6703	0	0	0	0	0
6704	0	0	0	0	0
6801	0	0	0	0	0
6802	0	0	0	0	0
6803	0	0	0	0	0
6804	0	0	0	0	0
6901	0	0	0	0	0
6902	0	0	0	0	0
6903	0	0	0	0	0
6904	0	0	0	0	0
7001	0	0	0	0	0
7002	0	0	0	0	0
7003	0	0	0	0	0
7004	0	0	0	0	0
7101	0	0	0	0	0
7102	0	0	0	0	0
7103	0	0	0	0	0
7104	0	0	0	0	0
7201	0	0	0	0	0
7202	0	0	0	0	0
7203	0	0	0	0	0
7204	0	0	0	0	0
7301	0	0	0	0	0
7302	0	0	0	0	0
7303	0	0	0	0	0
7304	0	0	0	0	0
7401	0	0	0	0	0
7402	0	0	0	0	0
7403	0	0	0	0	0
7404	0	0	0	0	0
7501	0	0	0	0	0
7502	0	0	0	0	0
7503	0	0	0	0	0
7504	0	0	0	0	0
7601	0	0	0	0	0
7602	0	0	0	0	0
7603	0	0	0	0	0
7604	0	0	0	0	0
7701	0	0	0	0	0
7702	0	0	0	0	0
7703	0	0	0	0	0
7704	0	0	0	0	0
7801	0	0	0	0	0
7802	0	0	0	0	0
7803	0	0	0	0	0
7804	0	0	0	0	0
7901	0	0	0	0	0
7902	0	0	0	0	0
7903	0	0	0	0	0
7904	0	0	0	0	0
8001	0	0	0	0	0
8002	0	0	0	0	0
8003	0	0	0	0	0
8004	0	0	0	0	0
8101	0	0	0	0	0
8102	0	0	0	0	0
8103	0	0	0	0	0
8104	0	0	0	0	0
8201	0	0	0	0	0
8202	0	0	0	0	0
8203	0	0	0	0	0
8204	0	0	0	0	0
8301	0	0	0	0	0
8302	0	0	0	0	0
8303	0	0	0	0	0
8304	0	0	0	0	0
8401	0	0	0	0	0
8402	0	0	0	0	0
8403	0	0	0	0	0
8404	0	0	0	0	0
8501	0	0	0	0	0
8502	0	0	0	0	0
8503	0	0	0	0	0
8504	0	0	0	0	0
8601	0	0	0	0	0
8602	0	0	0	0	0
8603	0	0	0	0	0
8604	0	0	0	0	0
8701	0	0	0	0	0
8702	0	0	0	0	0
8703	0	0	0	0	0
8704	0	0	0	0	0
8801	0	0	0	0	0
8802	0	0	0	0	0
8803	0	0	0	0	0
8804	0	0	0	0	0
8901	0	0	0	0	0
8902	0	0	0	0	0
8903	0	0	0	0	0
8904	0	0	0	0	0
9001	0	0	0	0	0
9002	0	0	0	0	0
9003	0	0	0	0	0
9004	0	0	0	0	0
9101	0	0	0	0	0
9102	0	0	0	0	0
9103	0	0	0	0	0
9104	0	0	0	0	0
9201	0	0	0	0	0
9202	0	0	0	0	0
9203	0	0	0	0	0
9204	0	0	0	0	0
9301	0	0	0	0	0
9302	0	0	0	0	0
9303	0	0	0	0	0
9304	0	0	0	0	0
9401	0	0	0	0	0
9402	0	0	0	0	0
9403	0	0	0	0	0
9404	0	0	0	0	0
9501	0	0	0	0	0
9502	0	0	0	0	0
9503	0	0	0	0	0
9504	0	0	0	0	0
9601	0	0	0	0	0
9602	0	0	0	0	0
9603	0	0	0	0	0
9604	0	0	0	0	0
9701	0	0	0	0	0
9702	0	0	0	0	0
9703	0	0	0	0	0
9704	0	0	0	0	0
9801	0	0	0	0	0
9802	0	0	0	0	0
9803	0	0	0	0	0
9804	0	0	0	0	0
9901	0	0	0	0	0
9902	0	0	0	0	0
9903	0	0	0	0	0
9904	0	0	0	0	0
0001	0	0	0	0	0
0002	0	0	0	0	0
0003	0	0	0	0	0
0004	0	0	0	0	0

TABLE 4A(continued)

Quarter	REC F1	REC F2	REC F3	REC F4	REC F8
6803.00	401749	872024E-01	41R159	411499	335805
6804.00	.356371E-01	.552811E-01	.235565	.110127	.724393
6901.00	-1.37995	1.353016	-1.40427	1.107190	-1.726603
6902.00	-1.606598	-1.394172E-01	.352814	1.283992	-1.184358
6903.00	.153207	.368905	.243380	-.283922	.560953
6904.00	.347819	-1.16810	-1.243207	.517522	-1.56895
7001.00	-1.32623	1.356228E-01	1.1654E-01	-1.22040	1.097516
7002.00	1.304025	-1.356228E-01	.919330	.974578	1.306772
7003.00	1.244120	3.223429	2.754485	3.10273	2.461683
7004.00	-1.438177	-3.05759	-2.754485	-.389009	-.441684
7101.00	1.966352	1.508676E-01	.259426	.594522	.942505
7102.00	-1.325133	-1.308676E-01	3.482201	-.682476	-.432824
7103.00	3.221860	1.665032	1.537283	1.776837	4.328276
7104.00	1.291062	1.665032	.537283	.617685	1.419566
7201.00	-1.032200	-1.664441	-1.49443	-1.431900	-1.919566
7202.00	1.436947	1.505677	1.402631	.431900	1.919566



TABLE 4B

Gordon-Jorgenson Model ( $\hat{\rho} = .6223$ ): Test Statistics

Number of residuals = 64

Global Location Tests			p-value				
	t-Test		-.1203	.9042			
No of Positive Residuals		38	.1686				
Wilcoxon Test		1126	.5652				
Runs Tests	No of Runs		29	.2250			
	Length of the Longest Run		14	.0032			
Serial Correlation Tests	Modified Von Neumann Ratio		1.967	$\geq .10$			
	Rank Tests						
	k	Signed-Rank Tests		Sign-Tests			
		$S_k$	$S'_k$	p-value	$S_k$	$S'_k$	p-value
	1	1161	1.047	.2949	35	.8819	.4500
	2	1103	.8869	.3751	36	1.270	.2529
	3	1114	1.210	.2262	36	1.408	.2000
	4	789	-.9276	.3536	26	-1.033	.3663
	5	897	.0906	.9278	32	.6509	.6029
	6	1126	2.094	.0362	36	1.838	.0869
	7	1092	2.109	.0349	37	2.252	.0331
	8	787	-.0897	.9285	30	.5345	.6889
	9	870	.8379	.4021	28	.1348	1.0000
	10	710	-.2798	.7796	26	-.2720	.8919
	11	578	-1.217	.2235	24	-.6868	.5831
	12	696	.0638	.9492	25	-.2774	.8899

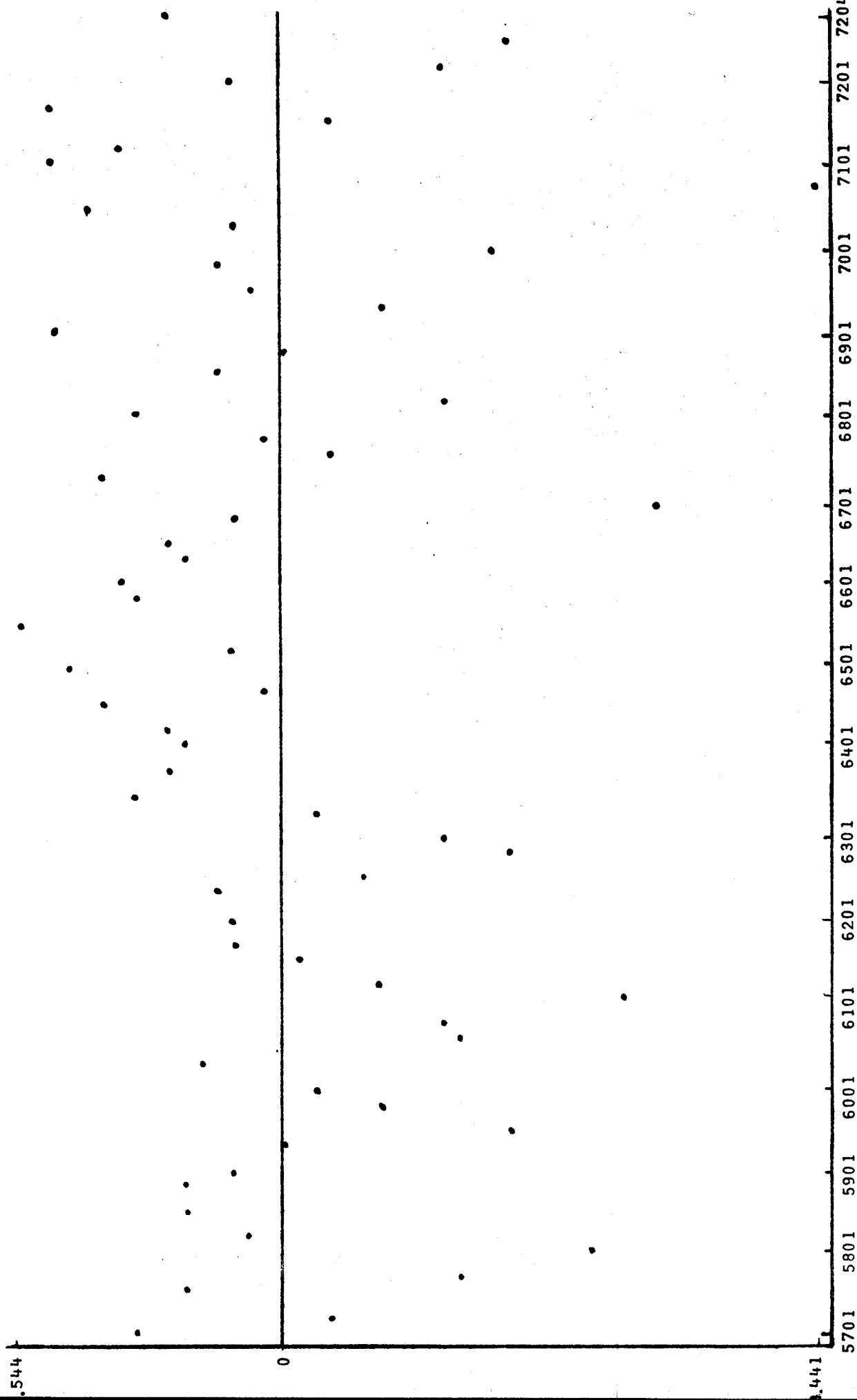


FIGURE 4A - Gordon-Jorgenson Model: One-Step Ahead Recursive Residuals ( $\hat{\rho} = .6223$ )

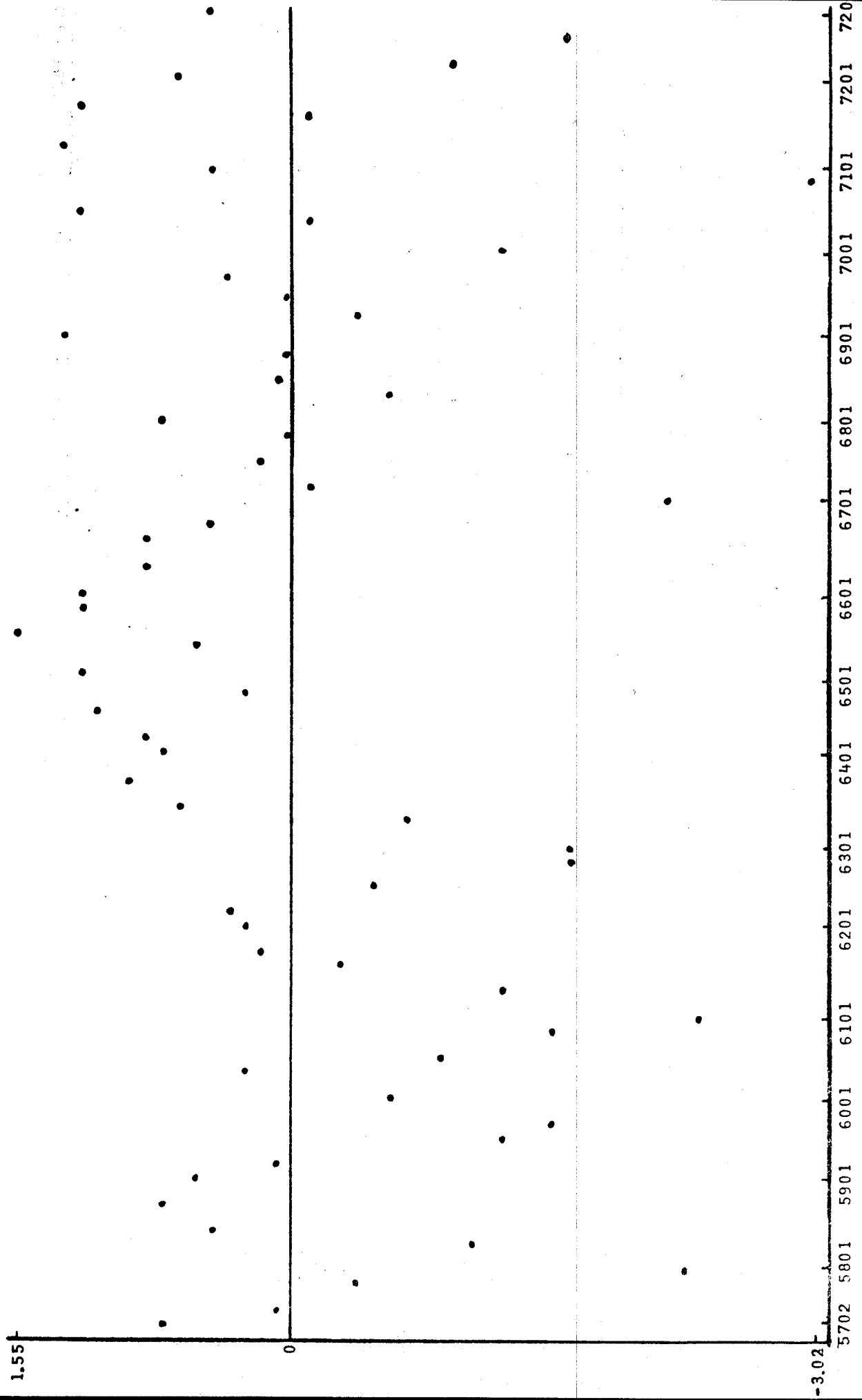
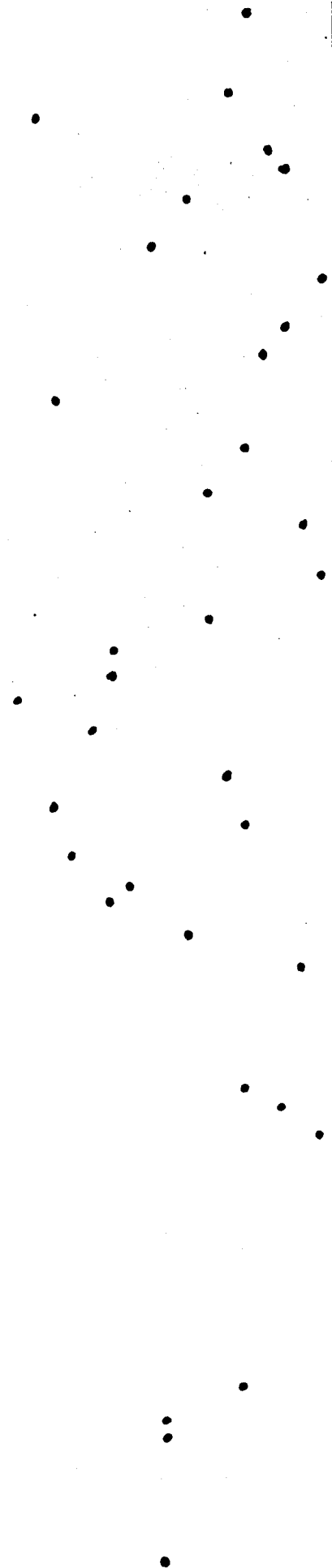
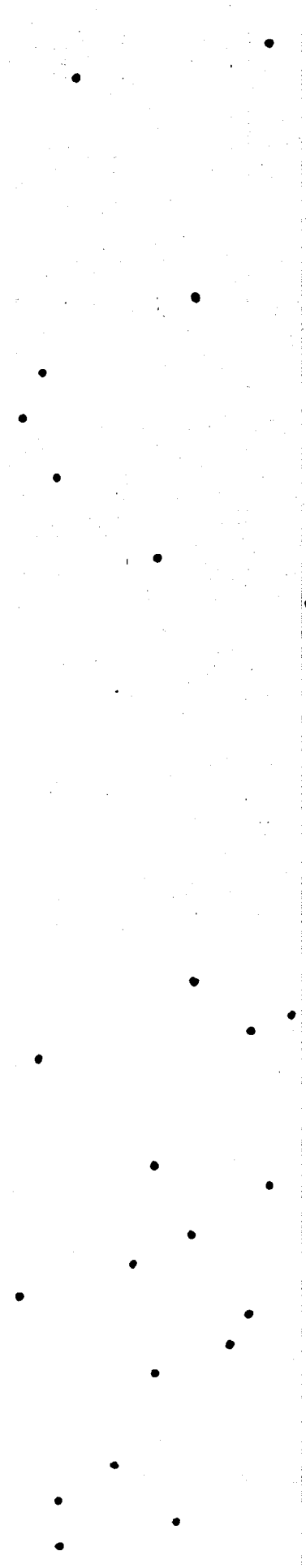


FIGURE 4B - Gordon-Jorgenson Model: Two-Steps Ahead Recursive Residuals ( $\hat{\rho} = .6223$ )

1.85



0



2.89

5703 5801 5901 6001 6101 6201 6301 6401 6501 6601 6701 6801 6901 7001 7101 7201 7204

FIGURE 4C - Gordon-Jorgensen Model: Three-Steps Ahead Recursive Residuals ( $\hat{\rho} = .6233$ )

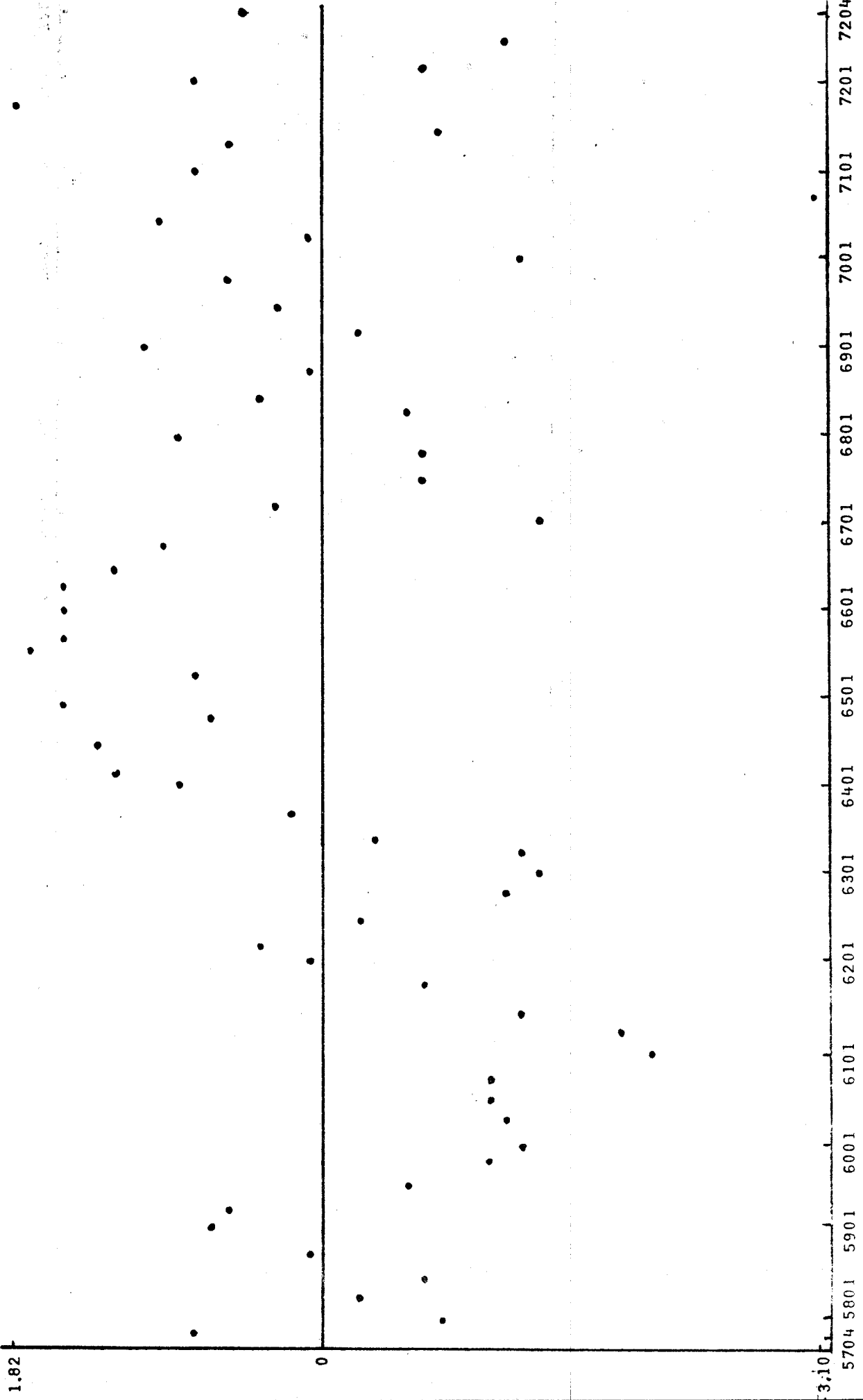


FIGURE 4D - Gordon-Jorgenson Model: Four-Steps Ahead Recursive Residuals ( $\hat{\rho} = .6223$ )

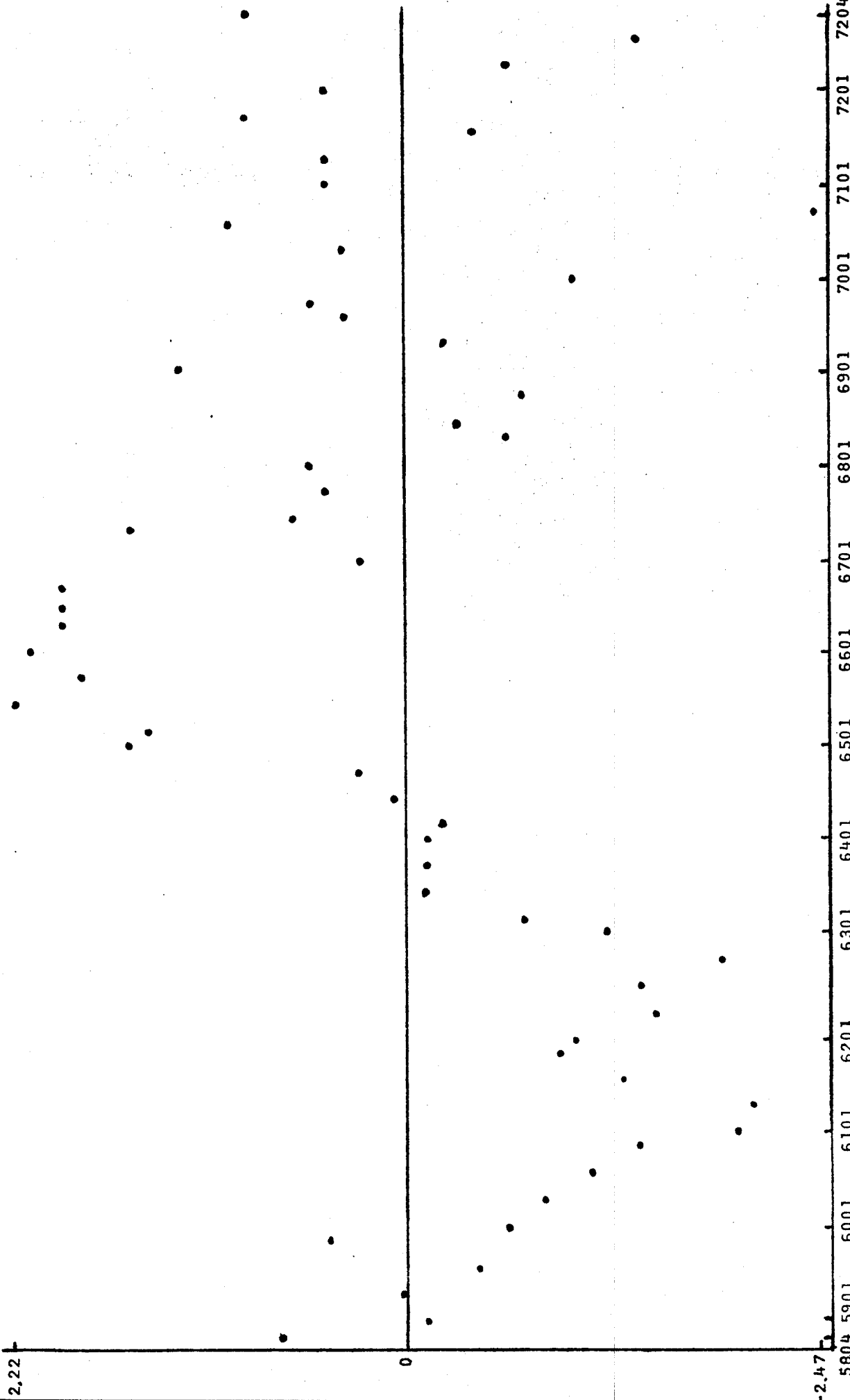


FIGURE 4E - Gordon-Jorgenson Model: Eight-Steps Ahead Recursive Residuals ( $\hat{\rho} = .6223$ )

22.8

5 % SIGNIFICANCE LINE

0

5 % SIGNIFICANCE LINE

-22.8

5701 5801 5901 6001 6101 6201 6301 6401 6501 6601 6701 6801 6901 7001 7101 7201 7200

FIGURE 4F - Gordon-Jorgenson Model: CUSUM of Recursive Residuals ( $\hat{\rho} = .6223$ )

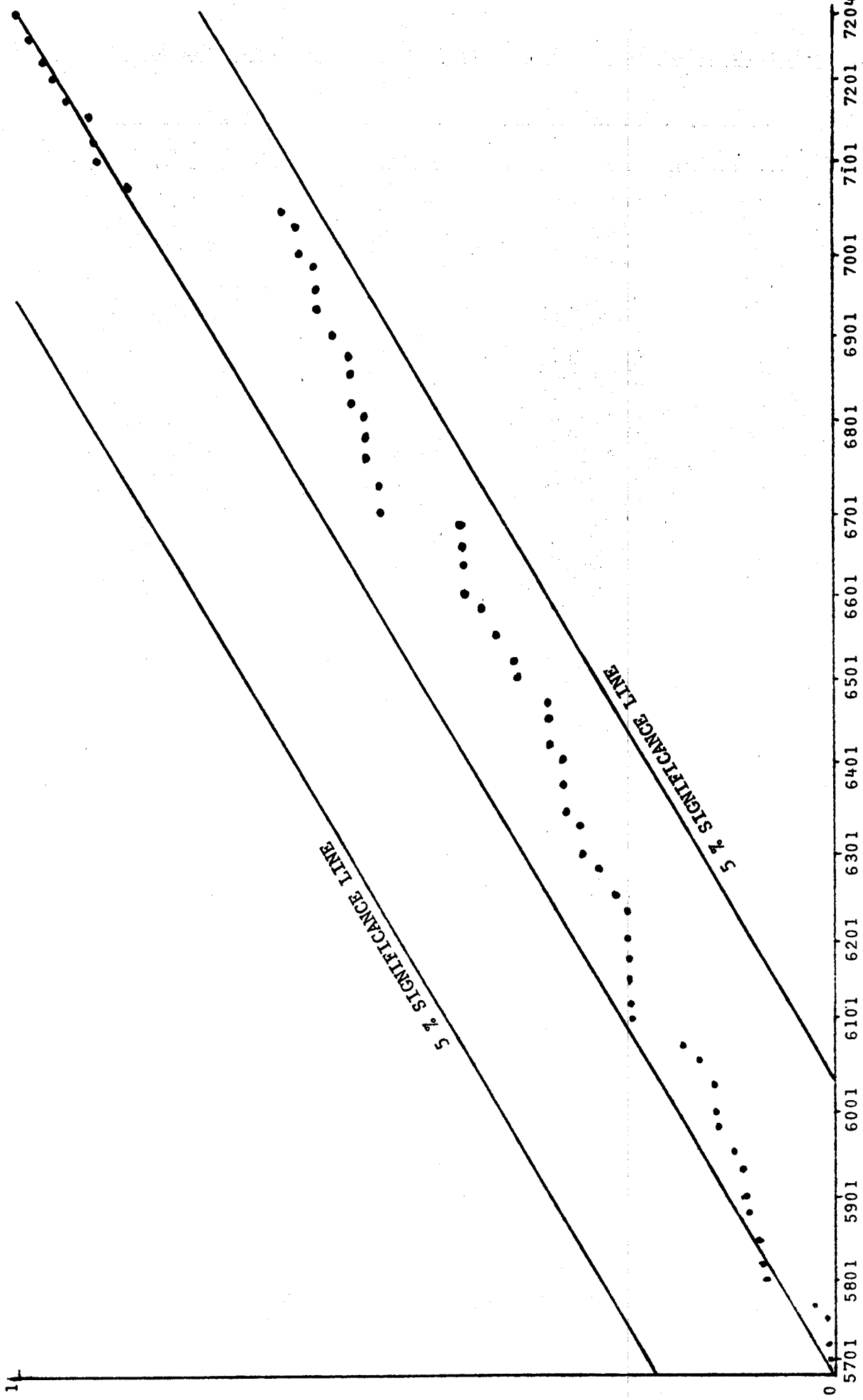


FIGURE 4G - Gordon-Jorgenson Model: CUSUM of Squares of Recursive Residuals ( $\hat{\rho} = .6223$ )



TABLE 5

Gordon-Jorgenson Model: Recursive Estimates ( $\hat{\rho} = .6223$ , Capital Subtracted)

1956/I-1922/IV

Quarter	a	a <sub>1</sub>	a <sub>2</sub>
5603.00	-269.217	.255279E-01	.337676E-02
5604.00	-533.578	.279458E-01	.414223E-02
5701.00	-44.24818	.595371E-01	.379051E-02
5702.00	-9.255772	.570352E-01	.685720E-03
5703.00	68.83340	.410830E-01	.140310E-03
5704.00	68.58204	.613426E-01	.707420E-03
5801.00	-63.26884	.733955E-01	.977030E-03
5802.00	-81.49223	.529594E-01	.778160E-03
5803.00	-75.31455	.387347E-01	.608640E-03
5804.00	-67.0780	.364840E-01	.203870E-03
5901.00	-65.3075	.119684E-01	.273990E-04
5902.00	-66.8731	.238700E-01	.101450E-03
5903.00	-60.6049	.133530E-01	.297000E-05
5904.00	-55.3328	.136920E-01	.111900E-04
6001.00	-51.7038	.144892E-01	.271500E-04
6002.00	-55.6916	.154230E-01	.297400E-04
6003.00	-43.7917	.133791E-01	.302200E-04
6004.00	-31.6055	.144037E-01	.653000E-04
6101.00	-11.5303	.167787E-01	.130640E-03
6102.00	-4.31984	.200291E-01	.180750E-03
6103.00	-2.84977	.191119E-01	.173390E-03
6104.00	-3.73363	.192502E-01	.173220E-03
6201.00	-4.59414	.193330E-01	.172460E-03
6202.00	-5.98786	.194906E-01	.171580E-03
6203.00	-5.76906	.180056E-01	.155480E-03
6204.00	-9.81220	.149145E-01	.130120E-03
6301.00	4.34310	.200030E-01	.196676E-03
6302.00	5.01822	.209752E-01	.208850E-03
6303.00	.470875	.123765E-01	.104630E-03
6304.00	-1.67775	.781920E-01	.498900E-04
6401.00	-2.77298	.547340E-01	.217400E-04
6402.00	-3.60712	.405850E-01	.444000E-05
6403.00	-4.65327	.541100E-01	.176000E-04
6404.00	-4.44952	.554150E-01	.194200E-04
6501.00	-5.00437	.268871E-01	.150300E-04
6502.00	-5.25985	.279230E-01	.137800E-04
6503.00	-7.45749	.686700E-01	.554100E-04
6504.00	-8.19962	.251700E-01	.771400E-04
6601.00	-9.08444	.455502E-01	.101370E-03
6602.00	-9.62948	.574710E-01	.115670E-03
6603.00	-10.3707	.855280E-01	.148250E-03
6604.00	-10.88145	.108389E-01	.174520E-03
6701.00	-8.85886	.135900E-01	.350800E-04
6702.00	-9.63780	.729460E-01	.132960E-03
6703.00	-9.48714	.597720E-01	.118000E-03
6704.00	-9.45833	.605080E-01	.118770E-03
6801.00	-9.73482	.518540E-01	.109610E-03
6802.00	-9.43976	.591200E-01	.117190E-03
6803.00	-9.54382	.585340E-01	.116720E-03
6804.00	-9.52379	.582340E-01	.116350E-03
6901.00	-9.99664	.714830E-01	.131930E-03
6902.00	-9.83204	.664940E-01	.125080E-03
6903.00	-9.89208	.678930E-01	.127740E-03
6904.00	-9.98831	.821540E-01	.143830E-03
7001.00	-9.77337	.473430E-01	.104620E-03
7002.00	-9.80918	.536640E-01	.111720E-03
7003.00	-9.96651	.653850E-01	.125100E-03
7004.00	-9.56351	.558620E-01	.113700E-03
7101.00	-9.67842	.545990E-01	.112520E-03
7102.00	-9.78678	.491540E-01	.106650E-03
7103.00	-9.70090	.568400E-01	.115070E-03
7201.00	-10.0124	.261710E-01	.814000E-04
7202.00	-10.0879	.2033940E-01	.750800E-04
7203.00	-9.84532	.352510E-01	.912600E-04
7204.00	-9.46094	.470720E-01	.103850E-03
7205.00	-9.65569	.468860E-01	.103920E-03

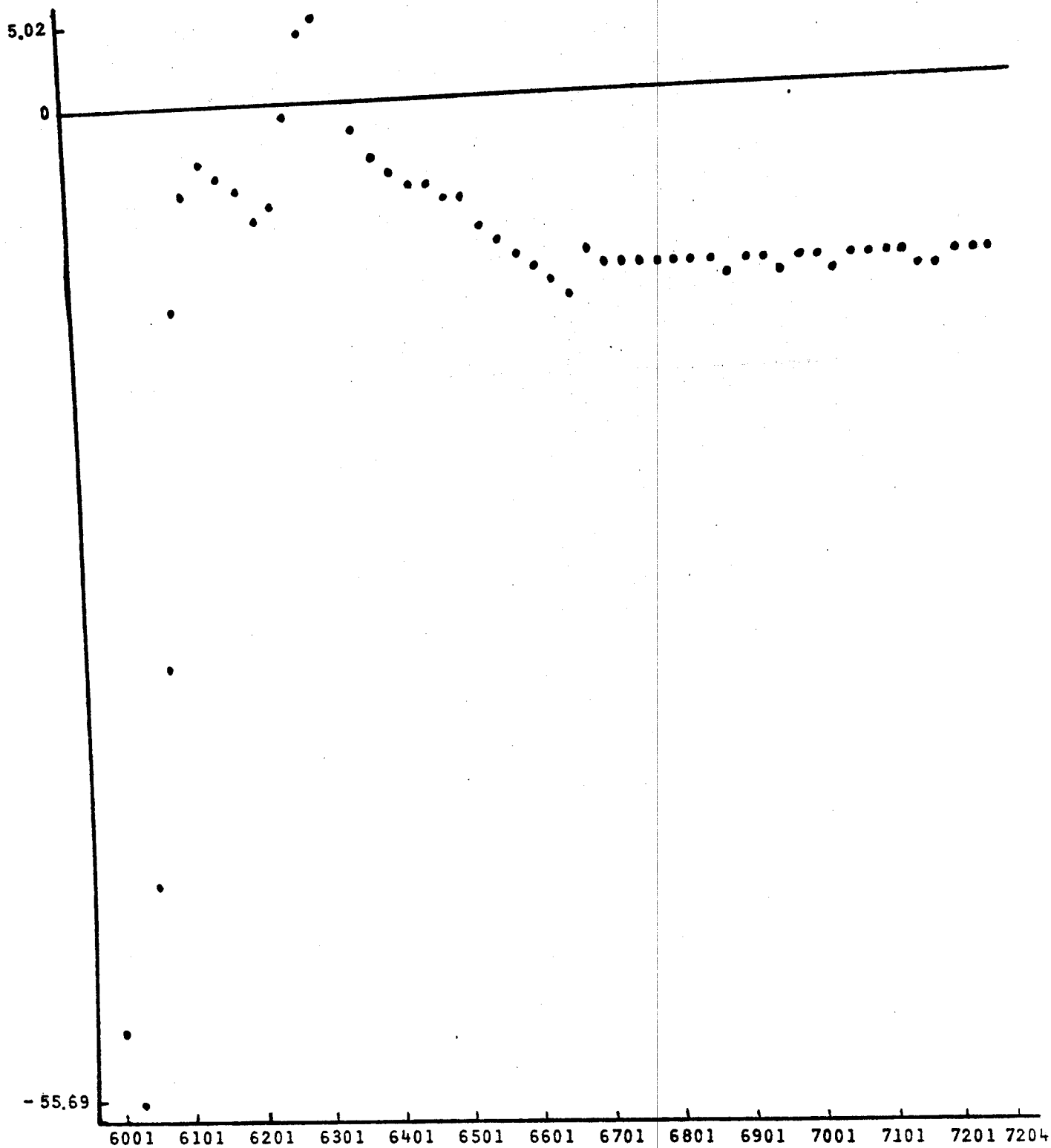


FIGURE 5A - Gordon-Jorgenson Model: Recursive Estimates of  $\alpha(\hat{\rho} = .6223, \text{Capital Subtracted})$

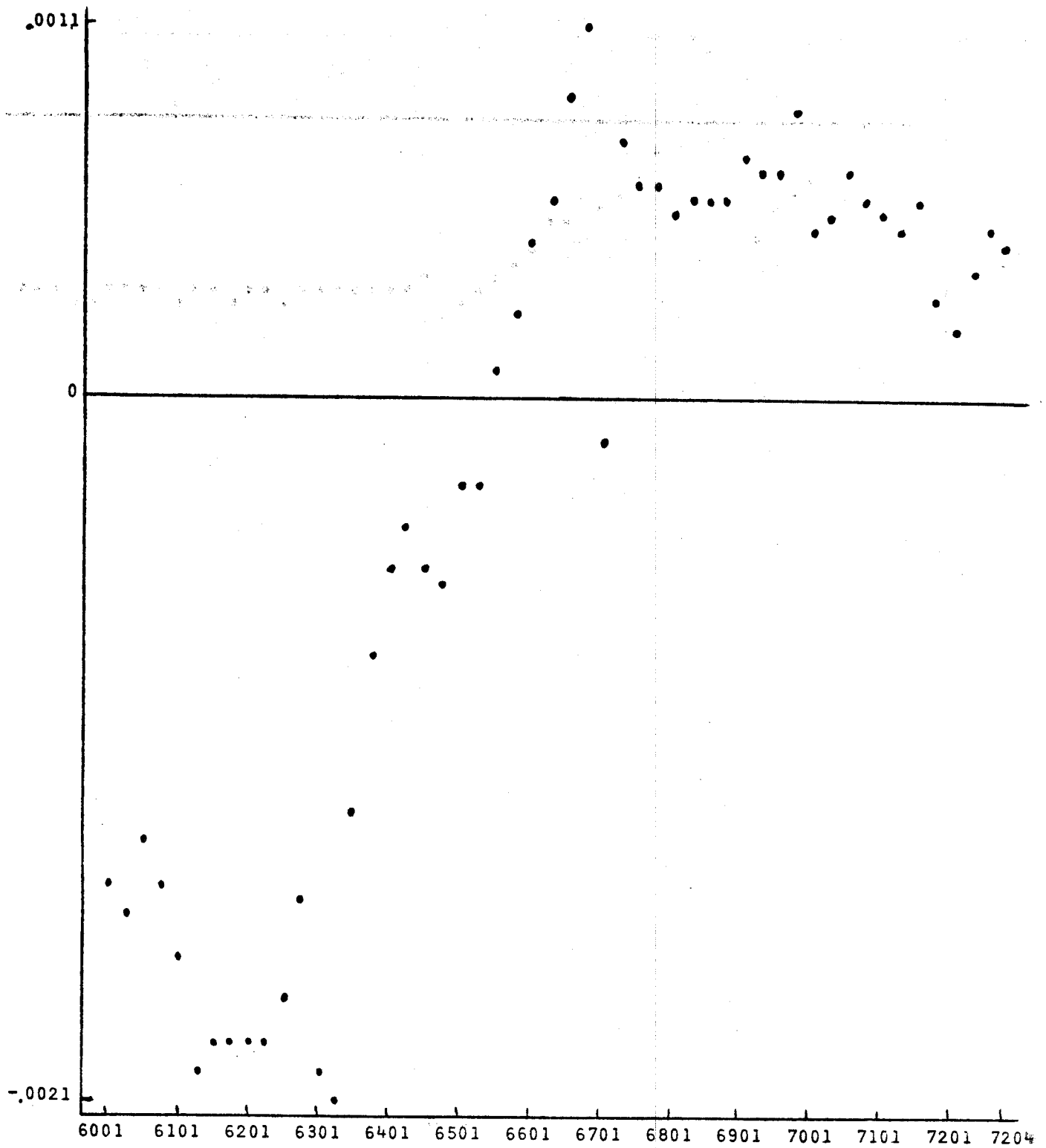


FIGURE 5B - Gordon-Jorgenson Model: Recursive Estimates of  $a_1$  ( $\hat{\rho} = .6223$ , Capital Subtracted)

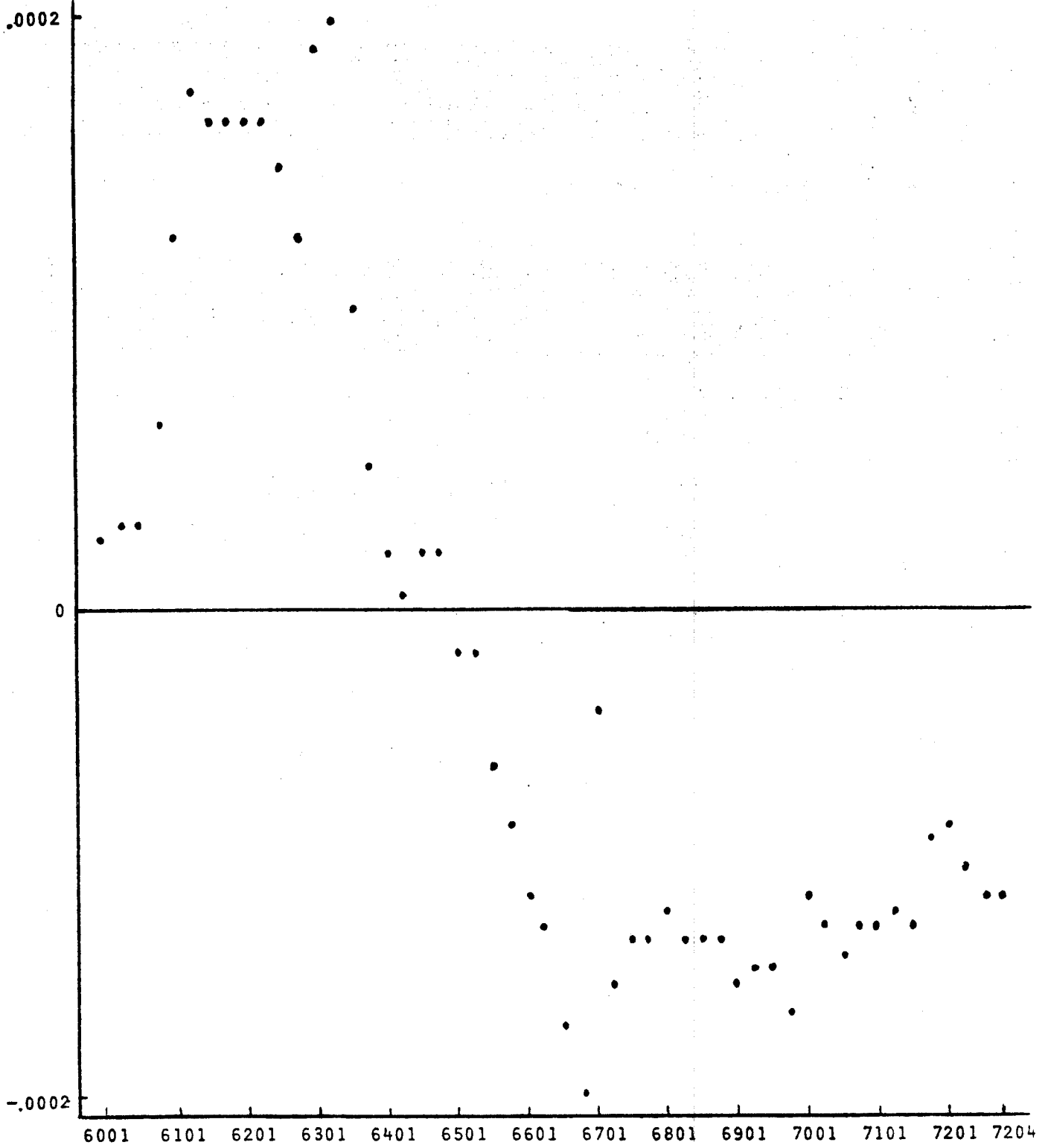


FIGURE 5C - Gordon-Jorgenson Model: Recursive Estimates of  $a_2$  ( $\hat{\rho} = .6223$ , Capital Subtracted)

TABLE 6A

Gordon-Jorgenson Model: Recursive Residuals ( $\hat{p} = .6223$ , Capital Subtracted)  
1956/IV-1972/IV

Quarter	REC F1	REC F2	REC F3	REC F4	REC F8
5604.00	419694	0	0	0	0
5701.00	854156	181309	0	0	0
5703.00	237042	711647	58997	316754	0
5704.00	652186	123677	321946	676919	0
5801.00	109345	-226273	96953	798197	0
5802.00	183353	-167592E-01	674265	220077	0
5803.00	199495	618592	802342	187077	44858
5804.00	668033	908778	022342	191831	7389661E-01
5901.00	718484	902349	148881	814338	2546298E-01
5902.00	768923	498644	27798	273334	181193E-01
5903.00	561219	696917	571752	475021E-01	168041E-01
5904.00	335644	433087	17292	534351	7086614
6002.00	120159	113671	17088	17725	8384772
6003.00	154716	126712	53554	531155	745542
6101.00	201239	801929	02292	801949	1745872
6102.00	143053	530339	87215	659991	163972
6104.00	143053	745738E-01	165770	849991	139219
6201.00	154716	186748	110270	699991	339209
6202.00	271637	302910	314655E-01	237760	136043
6204.00	116647	16343	97316	678074	522470
6301.00	1917977	130901	1883896	874827	892888
6302.00	117980	476046	24246	1874087	152898
6303.00	878205	733640	232106	207093	705888E-01
6401.00	391158	575150	322203	754705	205775E-01
6402.00	935737	524501	694033	101851	326720E-01
6403.00	953775	140529E-01	113493	1246137	1927679E-01
6404.00	1971435	920729	10456	70705	133335
6501.00	315009	181954	10433	101851	1315914
6502.00	162203	128430	147265	3310152	108814
6504.00	910516	1911915	80977	104049	719143
6601.00	688710	178474	26354	10660	58743
6602.00	772395	889346	3739	10660	58743
6604.00	434545	615100	04367	19161	508169
6701.00	201602	139817	91367	19161	508169
6702.00	136854	113815	286042	1482989	191824
6703.00	458638E-01	985141E-01	64439E-01	426790	236863
6801.00	88678	877344	867902	878118	1914863

TABLE 6A(continued)

Quarter	REC F1	REC F2	REC F3	REC F4	REC F8
6802:00	- .980866	- .900789	- .904860	- .917245	- .913657
6803:00	- .339659	- .264984	- .335599	- .327033	- .146033
6804:00	- .624798	- .382518	- .106836	- .415773	- .134892
6901:00	- .153113	- .140602	- .140982	- .138404	- .155855
6902:00	- .215166	- .173178	- .290457	- .384603	- .144012
6903:00	- .407980	- .425308	- .365099	- .521813	- .295015
7001:00	- .130278	- .116806	- .113764	- .118716	- .120366
7002:00	- .303136	- .173291	- .112291	- .118627	- .108848
7003:00	- .135083	- .133713	- .151336	- .126927	- .133394
7004:00	- .355809	- .390214	- .394266	- .358922	- .344439
7101:00	- .179140	- .827114	- .732661	- .753879	- .764792
7102:00	- .432966	- .378695	- .340046	- .403969	- .371814
7103:00	- .133109	- .130613	- .151019	- .141188	- .149944
7201:00	- .102999	- .935642	- .636687	- .702085	- .681509
7202:00	- .142999	- .150641	- .140956	- .122080	- .127084
7203:00	- .167690	- .150698	- .140956	- .144080	- .164984

TABLE 6B

Gordon-Jorgenson Model ( $\hat{\rho} = .6223$ , Capital Deducted): Test Statistics

Number of residuals = 65

Global Location Tests			p-values			
	t-Test		-.4535	.6502		
No of Positive Residuals		35	.6201			
Wilcoxon Test		1112	.7962			
Runs Tests	No of Runs		34	.6460		
	Length of the Longest Run		7	.3892		
Serial Correlation Tests	Modified Von Neumann Ratio		1.974	$\geq .10$		
	Rank Tests					
	k	Signed-Rank Tests		Sign-Tests		
		$S_k$	$S'_k$	p-value	$S_k$	$S'_k$
1	1101	.4079	.6833	31	-.2500	.9007
2	1117	.7462	.4556	35	.8819	.4500
3	1106	.9079	.3639	30	-.2540	.8991
4	818	-.9158	.3598	25	-1.408	.2000
5	866	-.3607	.7183	25	-1.291	.2451
6	1100	1.623	.1046	34	1.172	.2976
7	1092	1.831	.0671	36	1.838	.0869
8	738	-.7032	.4820	25	-.9272	.4270
9	973	1.427	.1534	33	1.336	.2288
10	748	-.1843	.8538	26	-.4045	.7877
11	639	-.8912	.3728	25	-.5443	.6835
12	806	.8012	.4230	28	.4121	.7838

1,42

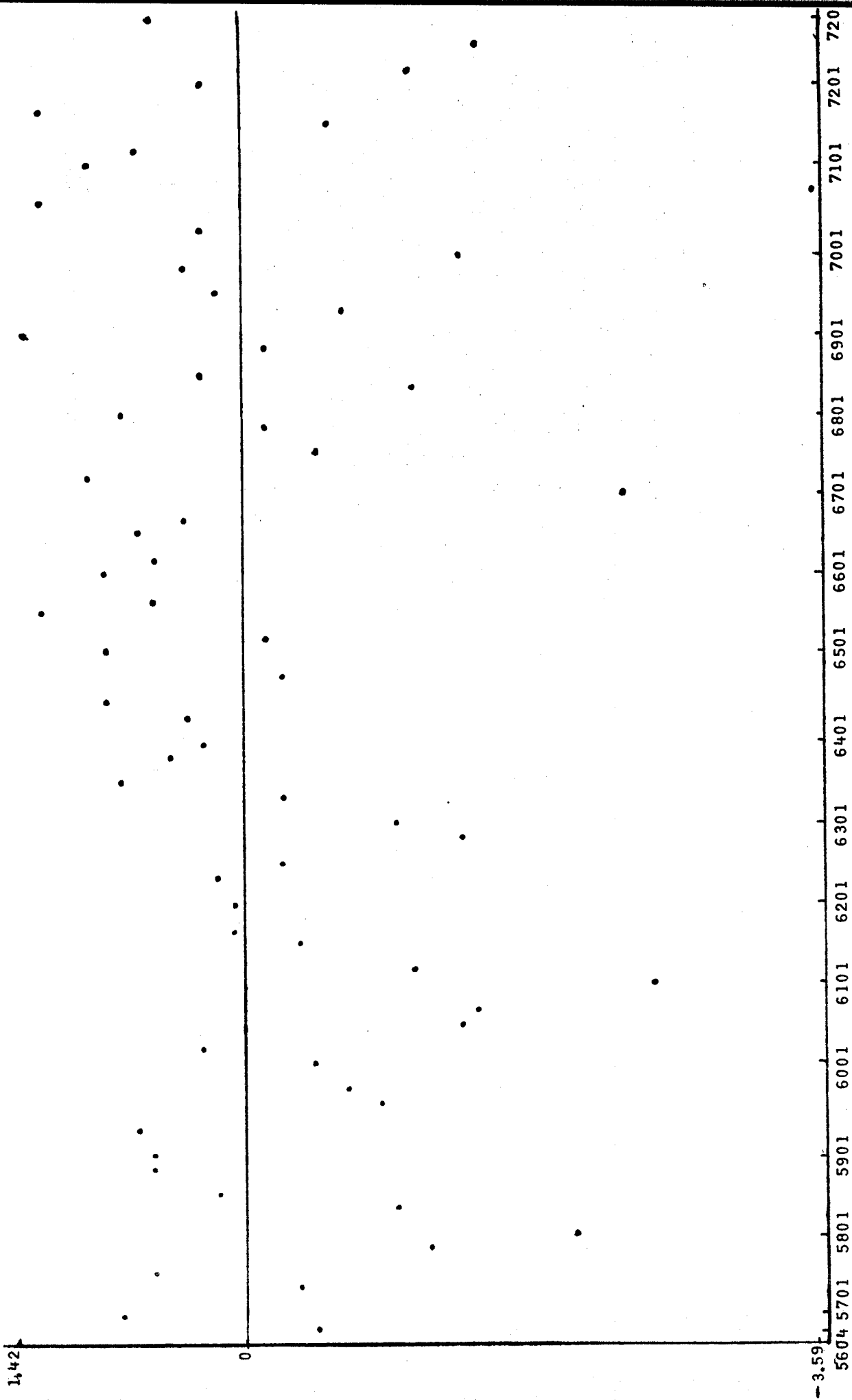


FIGURE 6A - Gordon-Jorgenson Model: One-Step Ahead Recursive Residuals ( $\hat{\rho} = .6223$ , Capital Subtracted)



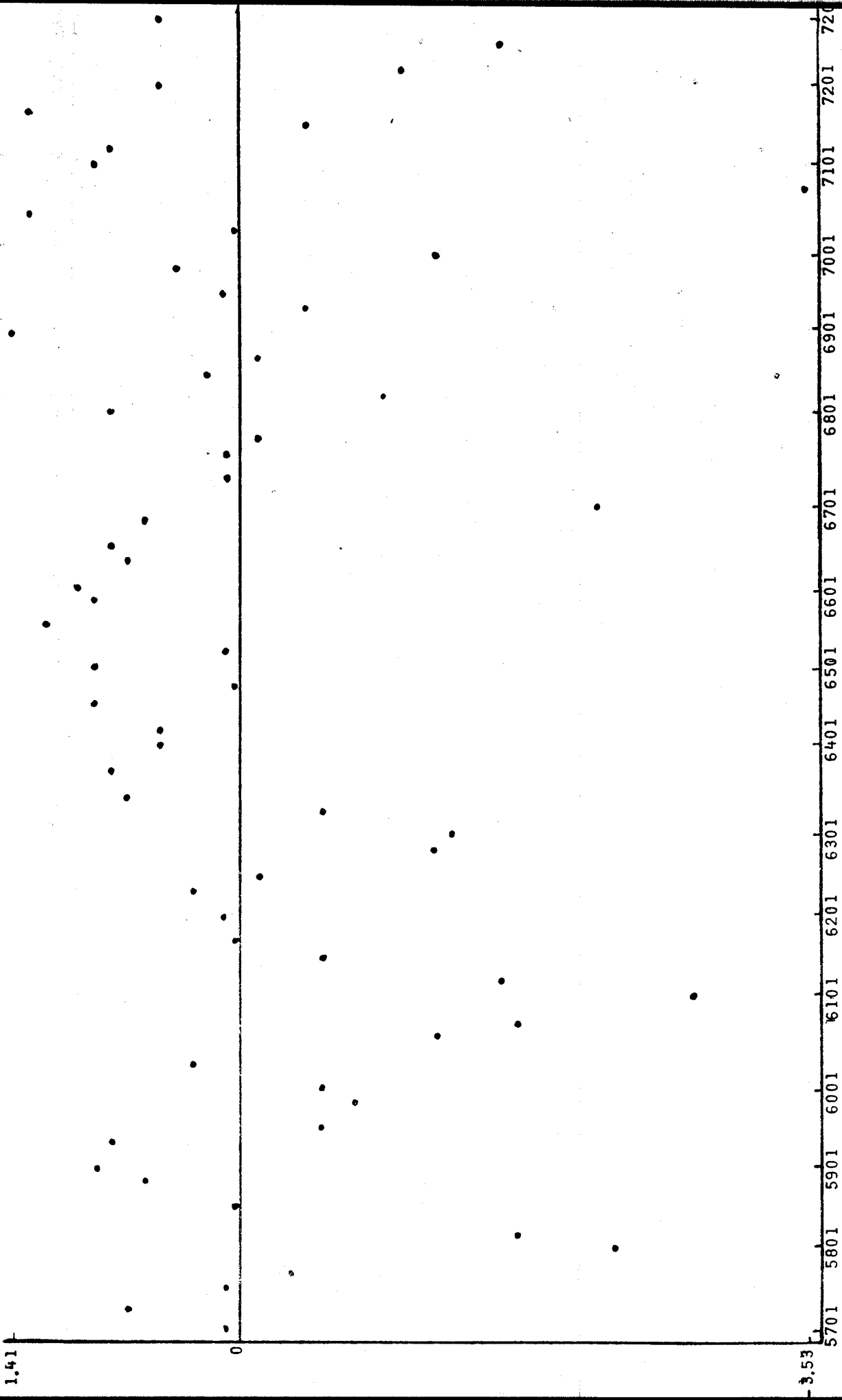
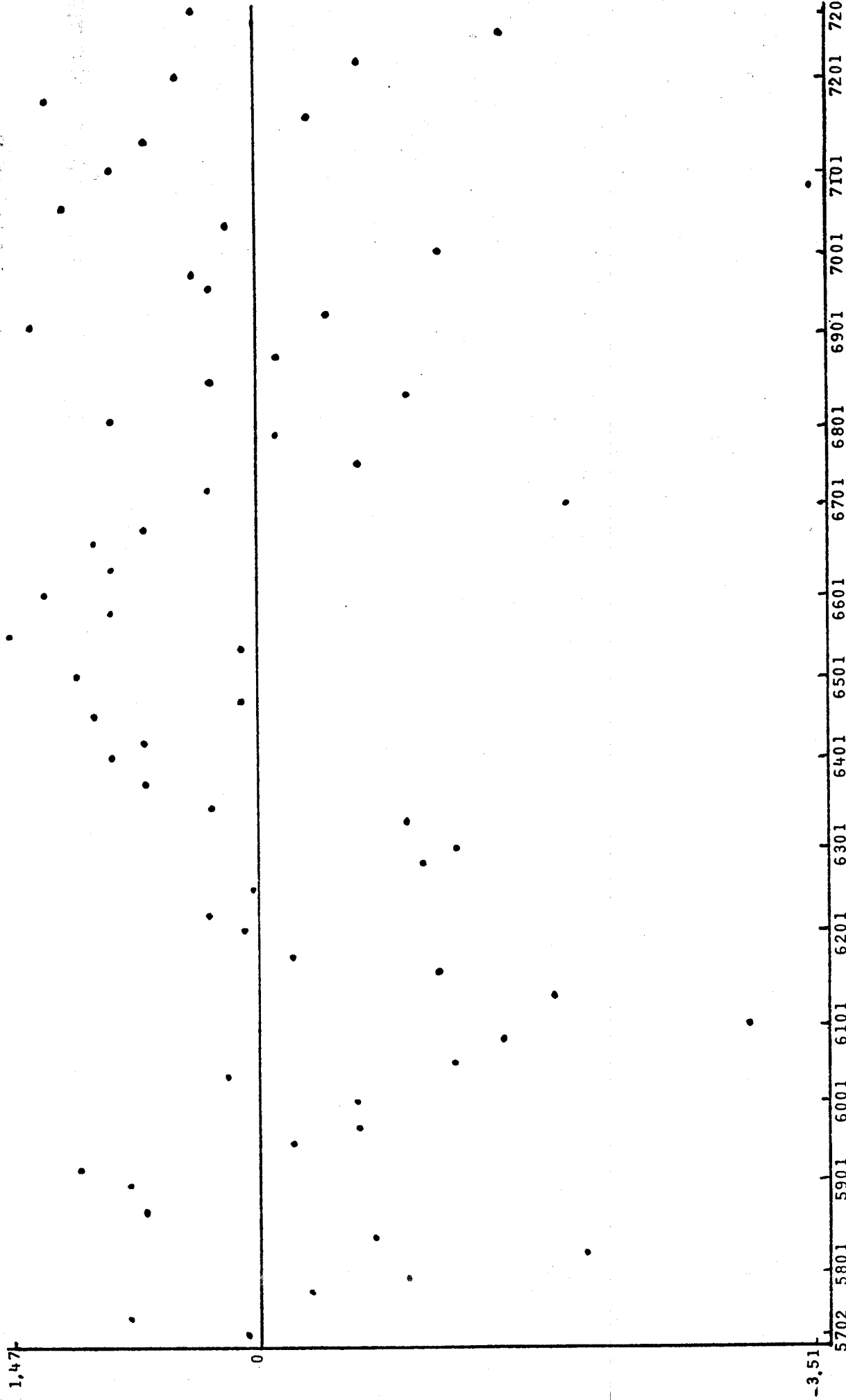


FIGURE 6B - Gordon-Jorgenson Model: Two-Steps Ahead Recursive Residuals ( $\hat{\rho} = .6223$ , Capital Subtracted)

1.47



-3.51

5702 5801 5901 6001 6101 6201 6301 6401 6501 6601 6701 6801 6901 7001 7101 7202

FIGURE 6C - Gordon-Jorgenson Model: Three-Steps Ahead Recursive Residuals ( $\hat{\rho} = .6223$ , Capital Subtracted)

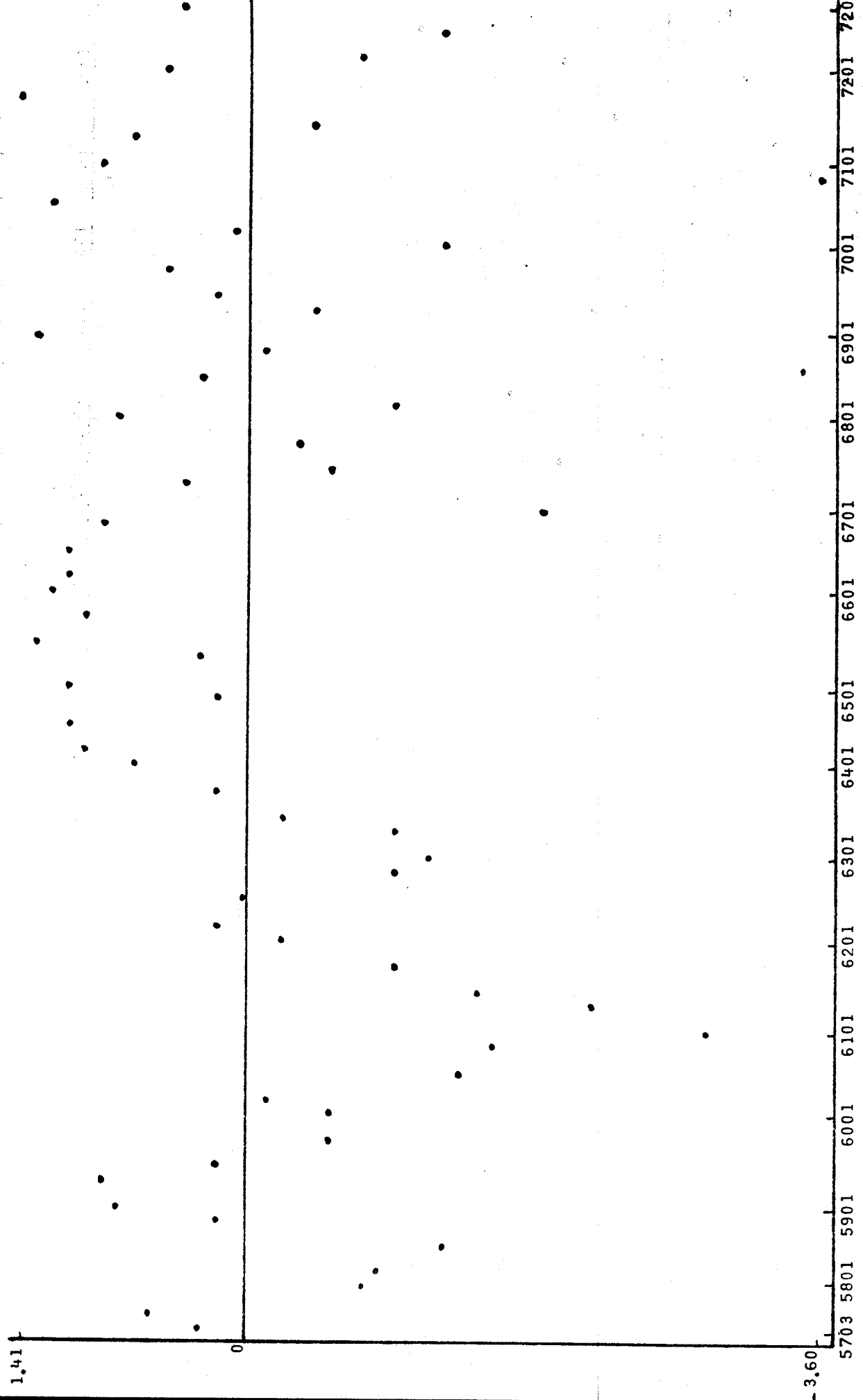


FIGURE 6D - Gordon-Jorgenson Model: Four-Steps Ahead Recursive Residuals ( $\hat{\rho} = .6223$ , Capital Subtracted)

1.72

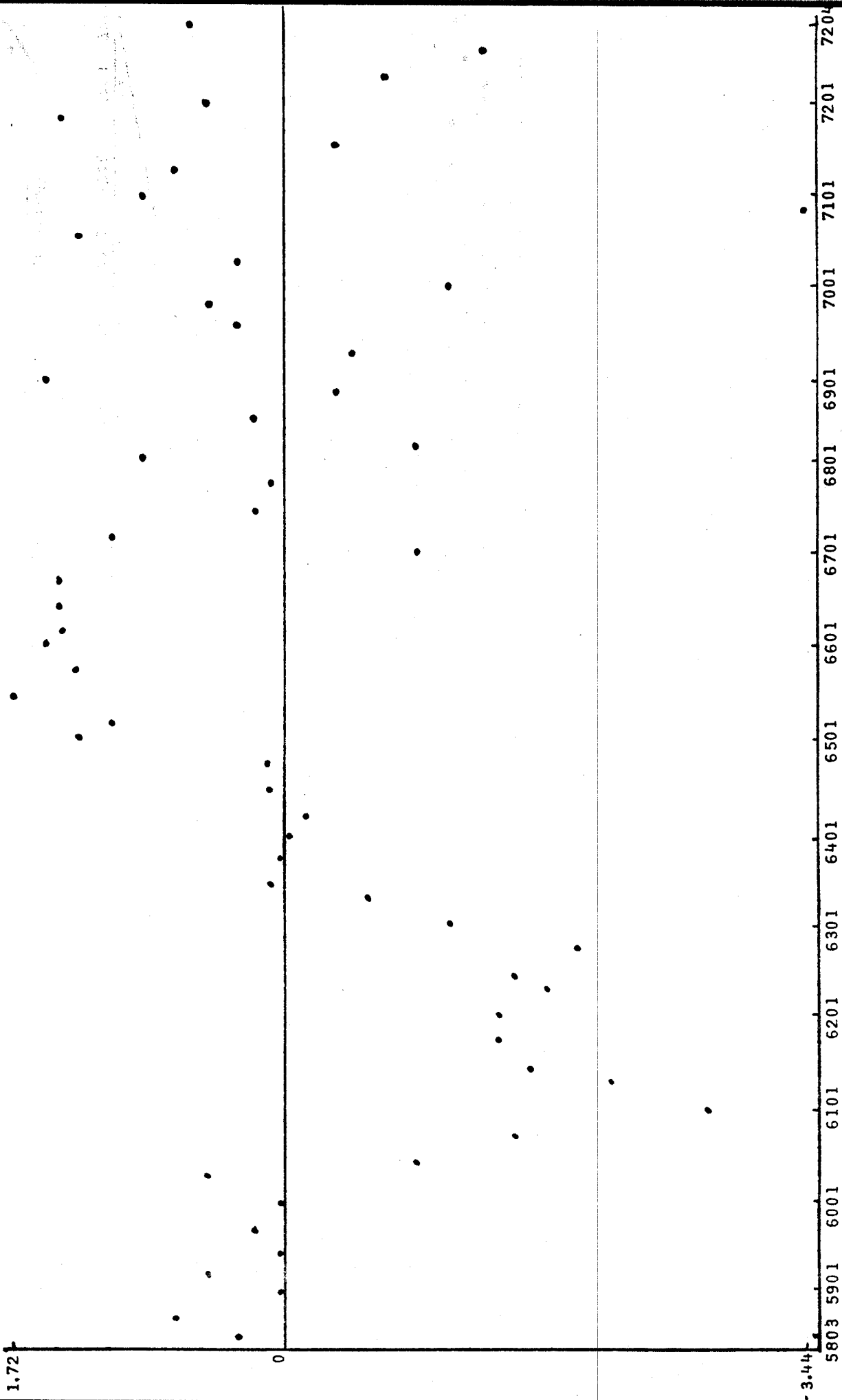
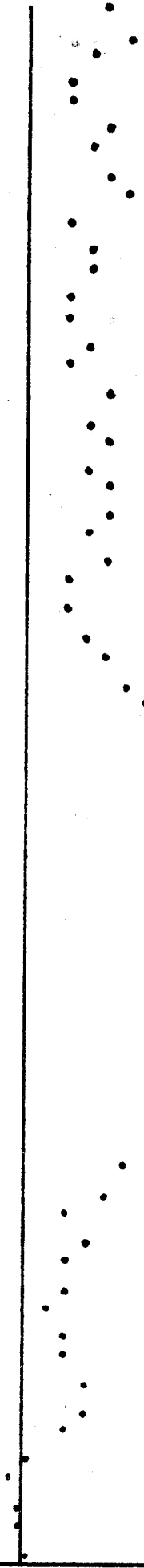


FIGURE 6E - Gordon-Jorgenson Model: Eight-Steps Ahead Recursive Residuals ( $\hat{\rho} = .6223$ , Capital Subtracted)

22.9

5 % SIGNIFICANCE LINE

0



5 % SIGNIFICANCE LINE

22.9

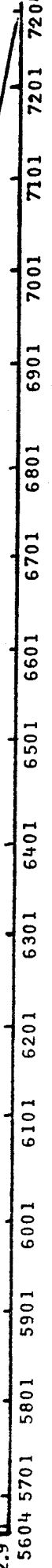


FIGURE 6F - Gordon-Jorgenson Model: CUSUM of Recursive Residuals ( $\hat{\rho} = .6229$ , Capital Subtracted)

06

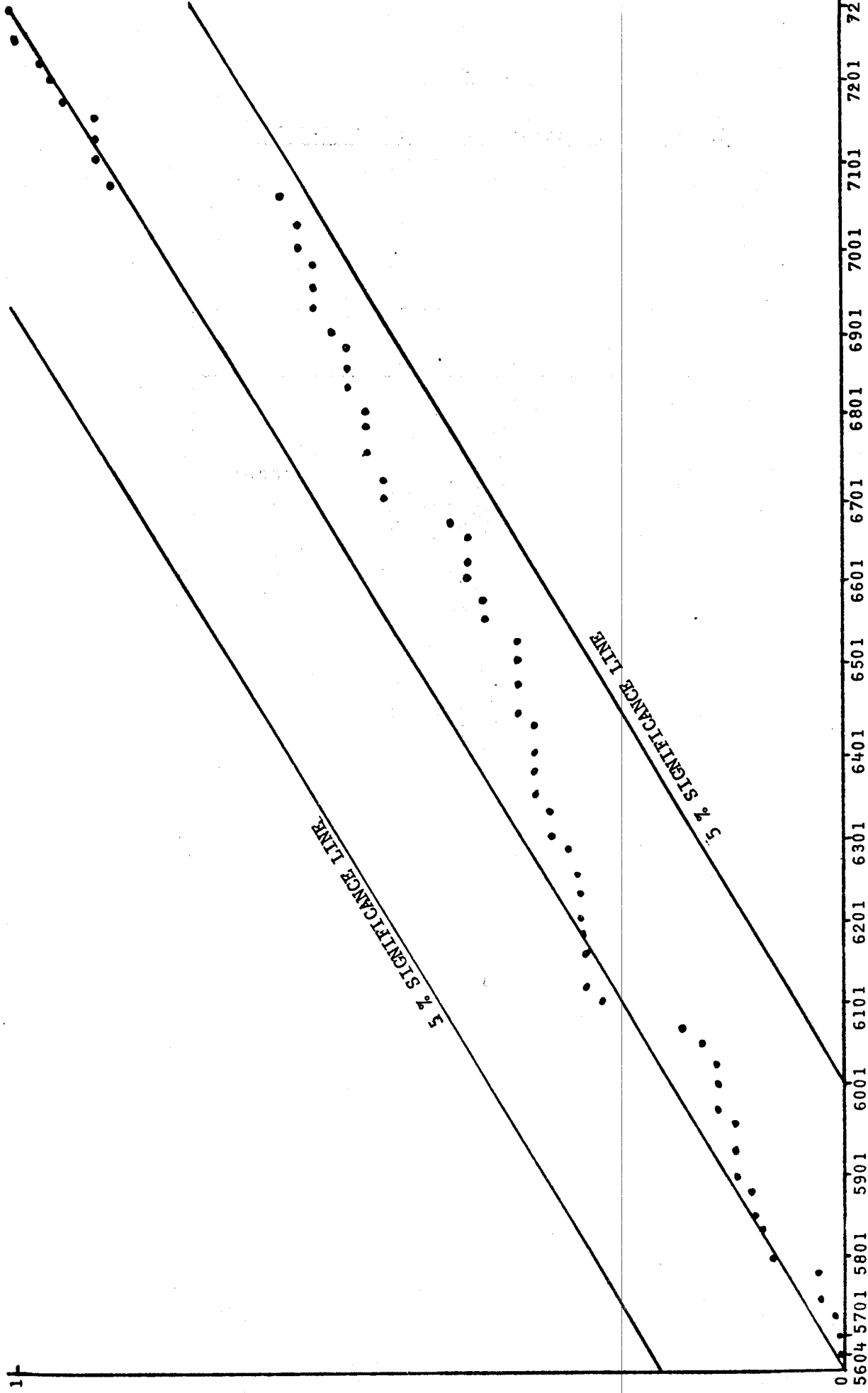


FIGURE 6G - Gordon-Jorgenson Model: CUSUM of Squares of Recursive Residuals ( $\hat{\rho} = .6223$ , Capital Subtracted)

TABLE 7

Effective Investment Tax Credit (1961-1972)<sup>1</sup>

Quarter	TC	Y
6101	.000000	.000000
6102	.000000	.000000
6103	.000000	.000000
6104	.000000	.000000
6201	.000000	.000000
6202	.33100000	.000000
6203	.35000000	.000000
6204	.39000000	.000000
6301	.43000000	.000000
6302	.47000000	.000000
6303	.51000000	.000000
6304	.55000000	.000000
6401	.56000000	.000000
6402	.56000000	.000000
6403	.56000000	.000000
6404	.56000000	.000000
6501	.56000000	.000000
6502	.56000000	.000000
6503	.56000000	.000000
6504	.56000000	.000000
6601	.56000000	.000000
6602	.56000000	.000000
6603	.56000000	.000000
6604	.56000000	.000000
6701	.56000000	.000000
6702	.56000000	.000000
6703	.56000000	.000000
6704	.56000000	.000000
6801	.56000000	.000000
6802	.56000000	.000000
6803	.56000000	.000000
6804	.56000000	.000000
6901	.56000000	.000000
6902	.56000000	.000000
6903	.56000000	.000000
6904	.56000000	.000000
7001	.56000000	.000000
7002	.56000000	.000000
7003	.56000000	.000000
7004	.56000000	.000000
7101	.56000000	.000000
7102	.56000000	.000000
7103	.56000000	.000000
7104	.56000000	.000000
7201	.56000000	.000000
7202	.56000000	.000000
7203	.56000000	.000000
7204	.56000000	.000000

<sup>1</sup>TC is the rate of the investment tax credit and Y is the dummy variable for the Long Amendment. Before 1961, TC and Y are both equal to zero.

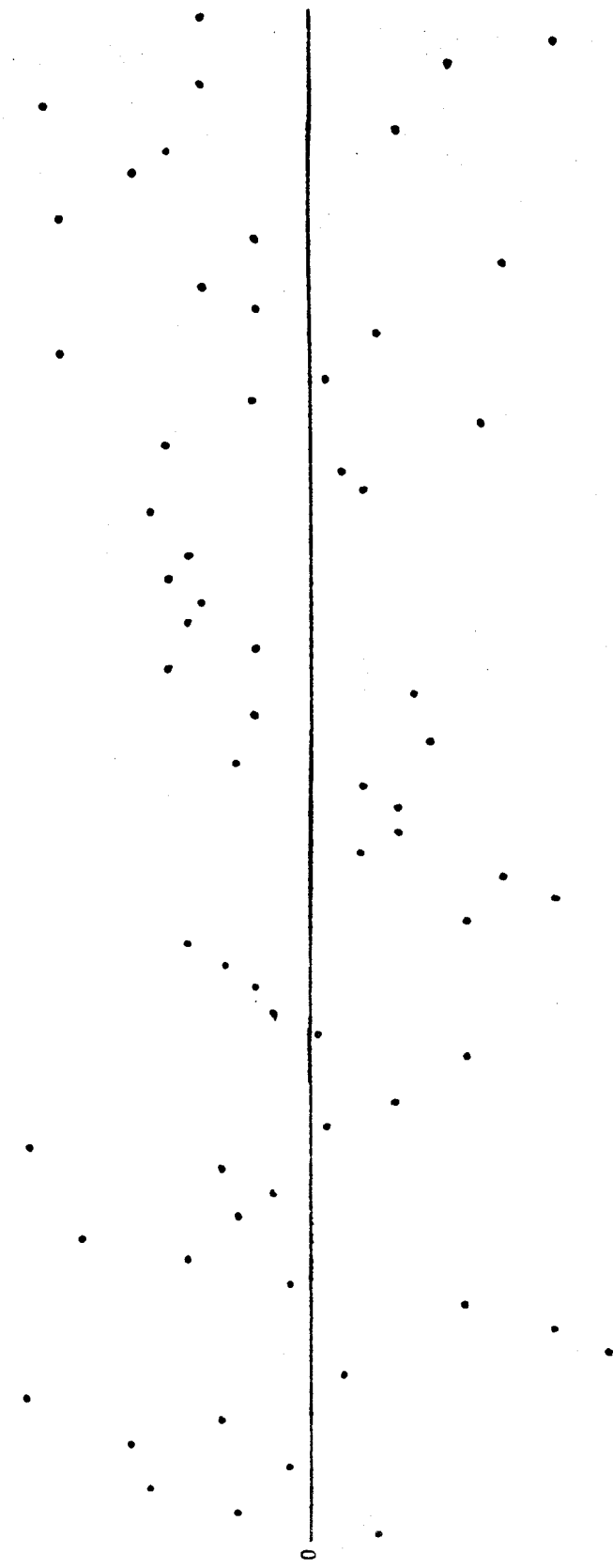


FIGURE 7 - Gordon-Jorgenson Model: Generalized Least Squares Residuals (1956/I-1972/IV)

After autoregressive transformation with  $\hat{\rho} = .6223$ .



## REFERENCES

- Brown, R.L., Durbin, J., and Evans, J.M. (1975), "Techniques for Testing the Constancy of Regression Relationships over Time" (With Discussion), Journal of the Royal Statistical Society, Series B, 37, 149-192.
- Christensen, L.R., and Jorgenson, D.W. (1969), "The Measurement of U.S. Real Capital Input, 1929-1967", Review of Income and Wealth, 15, 293-320.
- Dufour, J.-M. (1979), Methods for Specification Errors Analysis with Macroeconomic Applications, Ph.D. Dissertation, University of Chicago.
- Dufour, J.-M. (1981a), "Rank Tests for Serial Dependence", Journal of Time Series Analysis, 2, 117-128.
- Dufour, J.-M. (1981b), "The Demand for Money During the German Hyperinflation (1921-23): A Recursive Stability Analysis", Cahier 8130, Département de science économique et Centre de recherche en développement économique, Université de Montréal.
- Dufour, J.-M. (1982), "Recursive Stability Analysis of Linear Regression Relationships: An Exploratory Methodology", Journal of Econometrics, Vol. 19, 31-76.
- Gordon, R.H., and Jorgenson, D.W. (1976), "The Investment Tax Credit and Counter-Cyclical Policy", in O. Eckstein, ed., Parameters and Policies in the U.S. Economy, Amsterdam: North-Holland, 275-314.
- Hall, R.E., and Jorgenson, D.W. (1967), "Tax Policy and Investment Behaviour", American Economic Review, 59, 391-414.
- Lucas, R. (1976), "Econometric Policy Evaluation: A Critique", in Vol. 1 of the Carnegie-Rochester Conferences on Public Policy, Supplement to the Journal of Monetary Economics, 19-46.